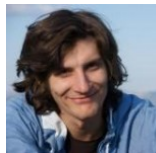
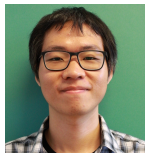


# Relative Goodness-of-Fit Tests for Models with Latent Variables

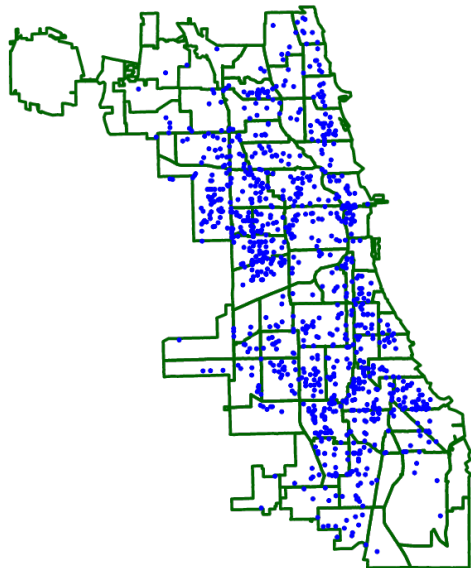
Arthur Gretton



Gatsby Computational Neuroscience Unit,  
Deepmind

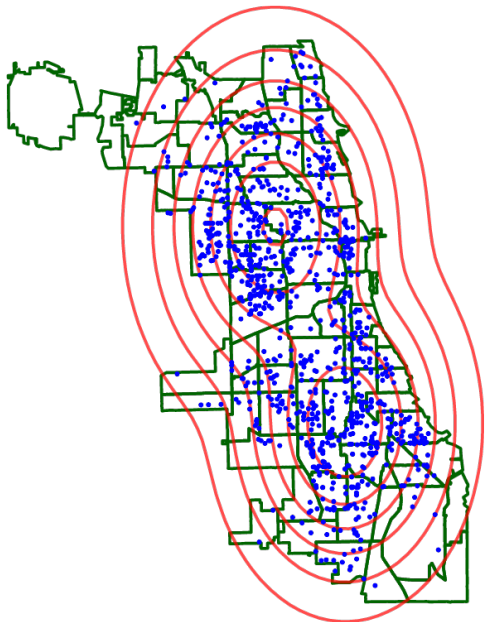
Department of Statistics, Columbia, 2023

## Model Criticism



Data = robbery events in  
Chicago in 2016.

# Model Criticism



Is this a good **model**?

# Model Criticism

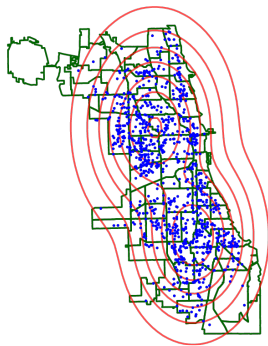
"All models are wrong."

G. Box (1976)

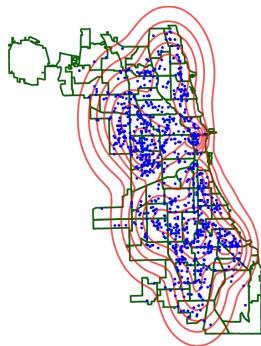


## Model comparison

- Have: two candidate models  $P$  and  $Q$ , and samples  $\{x_i\}_{i=1}^n$  from reference distribution  $R$
- Goal: which of  $P$  and  $Q$  is better?



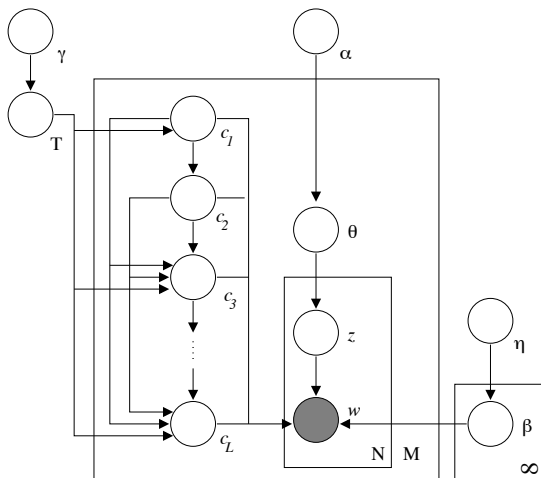
$P$  : two components



$Q$  : ten components

## Most interesting models have latent structure

Graphical model representation of hierarchical LDA with a nested CRP prior, Blei et al. (2003)



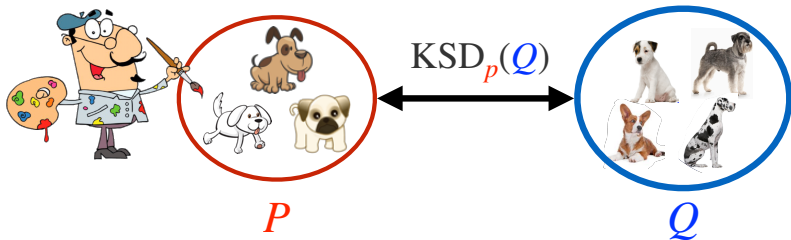
# Outline

## Relative goodness-of-fit tests for Models with Latent Variables

- The Maximum Mean Discrepancy: an integral probability metric
  - maximize difference in expectations using an RKHS witness class
- The kernel Stein discrepancy
  - Comparing a sample and a model: **Stein** modification of the witness class
- Constructing a **relative hypothesis test** using the KSD
- **Relative hypothesis tests with latent variables**

# Kernel Stein Discrepancy

- Model  $P$ , data  $\{x_i\}_{i=1}^n \sim Q$ .
- “All models are wrong” ( $P \neq Q$ ).

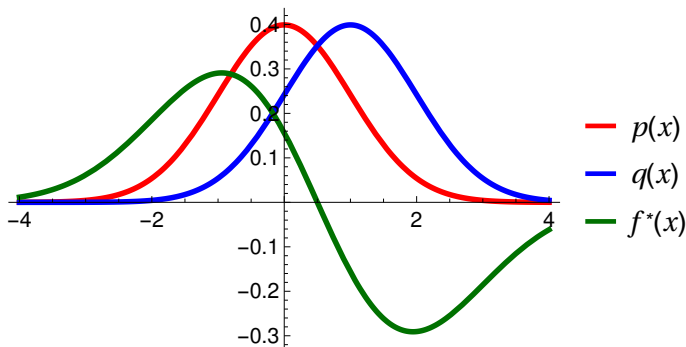


## Comparing a sample and model

Can we compute MMD with samples from  $Q$  and a model  $P$ ?

**Problem:** usually can't compute  $\mathbb{E}_{p f}$  in closed form.

$$\text{MMD}(P, Q) = \sup_{\|f\|_{\mathcal{F}} \leq 1} [\mathbb{E}_q f - \mathbb{E}_p f]$$



## Stein idea

To get rid of  $E_p f$  in

$$\sup_{\|f\|_{\mathcal{F}} \leq 1} [E_q f - E_p f]$$

we use the (1-D) **Langevin Stein operator**

$$[\mathcal{A}_p f](x) = \frac{1}{p(x)} \frac{d}{dx} (f(x)p(x))$$

Then

$$E_p \mathcal{A}_p f = 0$$

subject to appropriate boundary conditions.

$$E_p [\mathcal{A}_p f] = \int \left[ \frac{1}{p(x)} \frac{d}{dx} (f(x)p(x)) \right] p(x) dx = [f(x)p(x)]_{-\infty}^{\infty}$$

Gorham and Mackey (NeurIPS 15), Oates, Girolami, Chopin (JRSS B 2016)

## Stein idea

To get rid of  $E_{\textcolor{red}{p}}f$  in

$$\sup_{\|f\|_{\mathcal{F}} \leq 1} [E_{\textcolor{blue}{q}}f - E_{\textcolor{red}{p}}f]$$

we use the (1-D) **Langevin Stein operator**

$$[\mathcal{A}_{\textcolor{red}{p}}f](x) = \frac{1}{\textcolor{red}{p}(x)} \frac{d}{dx} (f(x)\textcolor{red}{p}(x))$$

Then

$$E_{\textcolor{red}{p}}\mathcal{A}_{\textcolor{red}{p}}f = 0$$

subject to appropriate boundary conditions.

Do not need to normalize  $\textcolor{red}{p}$ , or sample from it.

# Kernel Stein Discrepancy

Stein operator

$$\mathcal{A}_{\textcolor{red}{p}} f = \frac{1}{\textcolor{red}{p}(x)} \frac{d}{dx} (f(x) \textcolor{red}{p}(x))$$

Kernel Stein Discrepancy (KSD)

$$\text{KSD}_{\textcolor{red}{p}}(\textcolor{blue}{Q}) = \sup_{\|g\|_{\mathcal{F}} \leq 1} \mathbb{E}_{\textcolor{blue}{q}} \mathcal{A}_{\textcolor{red}{p}} g - \mathbb{E}_{\textcolor{red}{p}} \mathcal{A}_{\textcolor{red}{p}} g$$



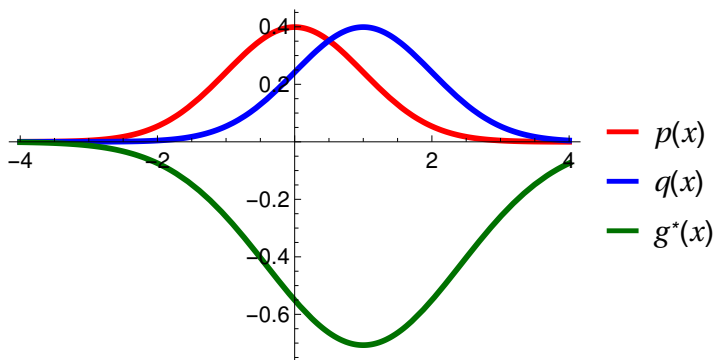
# Kernel Stein Discrepancy

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$$\mathcal{A}_p f = \frac{1}{p(x)} \frac{d}{dx} (f(x)p(x))$$

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$$\text{KSD}_p(Q) = \sup_{\|g\|_{\mathcal{F}} \leq 1} \mathbb{E}_q \mathcal{A}_p g - \cancel{\mathbb{E}_p \mathcal{A}_p g} = \sup_{\|g\|_{\mathcal{F}} \leq 1} \mathbb{E}_q \mathcal{A}_p g$$



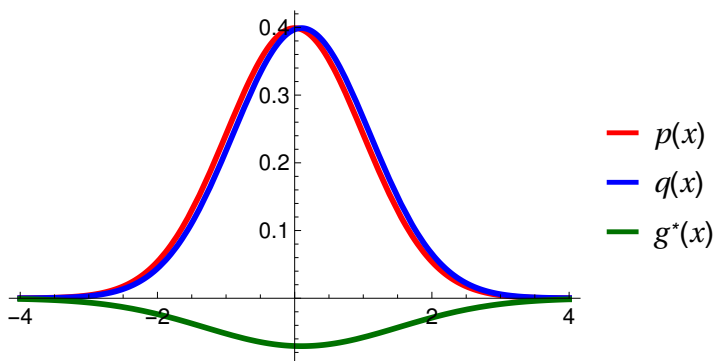
# Kernel Stein Discrepancy

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## Computing the kernel Stein discrepancy

How do we get the KSD in closed form (with kernels)?

Can we define “Stein features”?

$$[\mathcal{A}_{\textcolor{red}{p}}f](\textcolor{blue}{x}) = \frac{1}{\textcolor{red}{p}(x)} \frac{d}{dx} (f(x)\textcolor{red}{p}(x))$$

# Computing the kernel Stein discrepancy

How do we get the KSD in closed form (with kernels)?

Can we define “Stein features”?

$$\begin{aligned}[\mathcal{A}_{\textcolor{red}{p}}f](\textcolor{blue}{x}) &= \frac{1}{\textcolor{red}{p}(x)} \frac{d}{dx} (f(x)\textcolor{red}{p}(x)) \\&= \frac{d}{dx} f(x) + f(x) \frac{1}{\textcolor{red}{p}(x)} \frac{d}{dx} \textcolor{red}{p}(x) \\&= f(x) \frac{d}{dx} \log \textcolor{red}{p}(x) + \frac{d}{dx} f(x)\end{aligned}$$

# Computing the kernel Stein discrepancy

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Can we define “Stein features”?

$$\begin{aligned}[\mathcal{A}_p f](x) &= \frac{1}{p(x)} \frac{d}{dx} (f(x)p(x)) \\&= \frac{d}{dx} f(x) + f(x) \frac{1}{p(x)} \frac{d}{dx} p(x) \\&= f(x) \frac{d}{dx} \log p(x) + \frac{d}{dx} f(x) \\&\stackrel{?}{=} \langle f, \underbrace{\xi(x)}_{\text{stein features}} \rangle_{\mathcal{F}}\end{aligned}$$

where  $\mathbb{E}_{x \sim p} \xi(x) = 0$ .

# Computing the kernel Stein discrepancy

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Can we define “Stein features”?

$$\begin{aligned}[\mathcal{A}_{\textcolor{red}{p}} f](\textcolor{blue}{x}) &= \frac{1}{\textcolor{red}{p}(x)} \frac{d}{dx} (f(x) \textcolor{red}{p}(x)) \\&= \frac{d}{dx} f(x) + f(x) \frac{1}{\textcolor{red}{p}(x)} \frac{d}{dx} \textcolor{red}{p}(x) \\&= f(x) \frac{d}{dx} \log \textcolor{red}{p}(x) + \frac{d}{dx} f(x) \\&\stackrel{?}{=} \langle f, \underbrace{\textcolor{red}{\xi}(\textcolor{blue}{x})}_{\text{stein features}} \rangle_{\mathcal{F}}\end{aligned}$$

where  $\mathbb{E}_{x \sim \textcolor{red}{p}} \textcolor{red}{\xi}(x) = 0$ .

Intended destination:

$$\text{KSD}(\textcolor{red}{p}, \textcolor{blue}{q}, \mathcal{F}) = \sup_{\|g\|_{\mathcal{F}} \leq 1} \langle g, \mathbb{E}_{z \sim \textcolor{blue}{q}} \textcolor{red}{\xi}_z \rangle_{\mathcal{F}} = \|\mathbb{E}_{z \sim \textcolor{blue}{q}} \textcolor{red}{\xi}_z\|_{\mathcal{F}}$$

## Stein RKHS features

Reproducing property for the derivative: for differentiable  $k(x, x')$ ,

$$\frac{d}{dx}f(x) = \left\langle f, \frac{d}{dx}\varphi(x) \right\rangle_{\mathcal{F}} \quad \left\langle \frac{d}{dx}\varphi(x), \varphi(x') \right\rangle_{\mathcal{F}} = \frac{d}{dx}k(x, x')$$

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Using kernel derivative trick in (a),

$$\begin{aligned} [\mathcal{A}_p f](x) &= \left( \frac{d}{dx} \log p(x) \right) f(x) + \frac{d}{dx}f(x) \\ &= \left\langle f, \left( \frac{d}{dx} \log p(x) \right) \varphi(x) + \underbrace{\frac{d}{dx}\varphi(x)}_{(a)} \right\rangle_{\mathcal{F}} \\ &=: \langle f, \xi(x) \rangle_{\mathcal{F}}. \end{aligned}$$



## Proof: kernel derivative trick (on $[-\pi, \pi]$ )

**Proof:** differentiable translation invariant  $k(x, x')$ ,  $\mathcal{X} := [-\pi, \pi]$ ,  
periodic boundary

$$\frac{d}{dx}f(x) = \left\langle f, \frac{d}{dx}\varphi(x) \right\rangle_{\mathcal{F}} \quad \left\langle \frac{d}{dx}\varphi(x), \varphi(x') \right\rangle_{\mathcal{F}} = \frac{d}{dx}k(x, x')$$

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Fourier series representation:

$$f(x) = \sum_{\ell=-\infty}^{\infty} \hat{f}_{\ell} \exp(i\ell x), \quad \hat{f}_{\ell} = \frac{1}{2\pi} \int_{-\pi}^{\pi} f(x) \exp(-i\ell x) dx.$$

Fourier series representation of derivative:

$$\frac{d}{dx}f(x) \xrightarrow{F.S.} \left\{ (i\ell)\hat{f}_{\ell} \right\}_{\ell=-\infty}^{\infty}$$

## Proof: kernel derivative trick (on $[-\pi, \pi]$ )

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Fourier series representation of **derivative**:

$$\frac{d}{dx}f(x) \xrightarrow{F.S.} \left\{ (i\ell)\hat{f}_{\ell} \right\}_{\ell=-\infty}^{\infty} \quad \frac{d}{dx}k(x, \cdot) = \sum_{\ell=-\infty}^{\infty} (i\ell)\hat{k}_{\ell} \exp(i\ell(x - \cdot))$$

## Proof: kernel derivative trick (on $[-\pi, \pi]$ )

From previous slide,

$$\frac{d}{dx}f(x) \xrightarrow{F.S.} \left\{ (i\ell)\hat{f}_\ell \right\}_{\ell=-\infty}^{\infty} \qquad \frac{d}{dx}k(x, \cdot) = \sum_{\ell=-\infty}^{\infty} (i\ell)\hat{k}_\ell \exp(i\ell(x - \cdot))$$

We can write

$$\begin{aligned} \left\langle f, \frac{d}{dx}k(x, \cdot) \right\rangle_{\mathcal{F}} &= \sum_{\ell=-\infty}^{\infty} \frac{(\hat{f}_\ell) \left( \overline{-i\ell\hat{k}_\ell \exp(-i\ell x)} \right)}{\hat{k}_\ell} \\ &= \sum_{\ell=-\infty}^{\infty} (i\ell) (\hat{f}_\ell) (\exp(i\ell x)) = \frac{d}{dx}f(x). \end{aligned}$$

## Kernel Stein discrepancy: derivation

Closed-form expression for KSD: given independent  $x, x' \sim Q$ , then

$$\begin{aligned}\text{KSD}_p(Q) &= \sup_{\|g\|_{\mathcal{F}} \leq 1} \mathbb{E}_{x \sim q}([\mathcal{A}_p g](x)) \\ &= \sup_{\|g\|_{\mathcal{F}} \leq 1} \mathbb{E}_{x \sim q} \langle g, \xi_x \rangle_{\mathcal{F}} \\ &\stackrel{(a)}{=} \sup_{\|g\|_{\mathcal{F}} \leq 1} \langle g, \mathbb{E}_{x \sim q} \xi_x \rangle_{\mathcal{F}} = \|\mathbb{E}_{x \sim q} \xi_x\|_{\mathcal{F}}\end{aligned}$$

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Caution: (a) requires boundedness (Riesz),

$$|\mathbb{E}_{z \sim q} \langle f, \xi_z \rangle_{\mathcal{F}}| \leq \|f\|_{\mathcal{F}} \underbrace{\mathbb{E}_{z \sim q} \|\xi_z\|_{\mathcal{F}}}_{\text{bounded?}}$$

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Closed-form expression for KSD: given independent  $x, x' \sim Q$ , then

$$\begin{aligned}\text{KSD}_p(Q) &= \sup_{\|g\|_{\mathcal{F}} \leq 1} \mathbb{E}_{x \sim q} ([\mathcal{A}_p g](x)) \\ &= \sup_{\|g\|_{\mathcal{F}} \leq 1} \mathbb{E}_{x \sim q} \langle g, \xi_x \rangle_{\mathcal{F}} \\ &\stackrel{(a)}{=} \sup_{\|g\|_{\mathcal{F}} \leq 1} \langle g, \mathbb{E}_{x \sim q} \xi_x \rangle_{\mathcal{F}} = \|\mathbb{E}_{x \sim q} \xi_x\|_{\mathcal{F}}\end{aligned}$$

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Leading term

$$\|\xi_z\|_{\mathcal{F}}^2 = \left\langle \left( \frac{d}{dz} \log p(z) \right) k(z, \cdot), \left( \frac{d}{dz} \log p(z) \right) k(z, \cdot) \right\rangle_{\mathcal{F}} + \dots$$

implies  $\mathbb{E}_{x \sim q} \left( \frac{d}{dx} \log p(x) \right)^2 < \infty$ .



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Kernel expression in  $\mathbb{R}$ :

$$\begin{aligned}& \|\mathbb{E}_{x \sim q} \xi_x\|_{\mathcal{F}}^2 \\ &= \left\| \mathbb{E}_{x \sim q} \left( \varphi(x) \frac{d}{dx} \log p(x) + \frac{d}{dx} \varphi(x) \right) \right\|_{\mathcal{F}}^2 \\ &= \mathbb{E}_{x, x' \sim Q} \left( k(x, x') \frac{\partial p(x)}{p(x)} \frac{\partial p(x')}{p(x')} + \partial_1 k(x, x') \frac{\partial p(x')}{p(x')} \right. \\ & \quad \left. + \partial_2 k(x, x') \frac{\partial p(x)}{p(x)} + \partial_{12} k(x, x') \right)\end{aligned}$$

## Does the Riesz condition matter?

Consider the **standard normal**,

$$p(x) = \frac{1}{\sqrt{2\pi}} \exp(-x^2/2).$$

Then

$$\frac{d}{dx} \log p(x) = -x.$$

If  $q$  is a **Cauchy distribution**, then the integral

$$\mathbb{E}_{x \sim q} \left( \frac{d}{dx} \log p(x) \right)^2 = \int_{-\infty}^{\infty} x^2 q(x) dx$$

is undefined.

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# Kernel Stein discrepancy: population expression

Population kernel Stein discrepancy (in  $\mathbb{R}^D$ ):

$$\text{KSD}_p^2(\mathcal{Q}) = \mathbb{E}_{x, x' \sim \mathcal{Q}} h_p(x, x')$$

where

$$\begin{aligned} h_p(x, x') = & \mathbf{s}_p(x)^\top \mathbf{s}_p(x') k(x, x') + \mathbf{s}_p(x)^\top k_2(x, x') \\ & + \mathbf{s}_p(x')^\top k_1(x, x') + \text{tr}[k_{12}(x, x')] \end{aligned}$$

- $\mathbf{s}_p(x) \in \mathbb{R}^D = \frac{\nabla p(x)}{p(x)}$
- $k_1(a, b) := \nabla_x k(x, x')|_{x=a, x'=b} \in \mathbb{R}^D$ ,  
 $k_2(a, b) := \nabla_{x'} k(x, x')|_{x=a, x'=b} \in \mathbb{R}^D$ ,
- $k_{12}(a, b) := \nabla_x \nabla_{x'} k(x, x')|_{x=a, x'=b} \in \mathbb{R}^{D \times D}$

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- $\mathbf{s}_{\mathbf{p}}(\mathbf{x}) \in \mathbb{R}^D = \frac{\nabla \mathbf{p}(\mathbf{x})}{\mathbf{p}(\mathbf{x})}$
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If kernel is  $C_0$ -universal and  $\mathcal{Q}$  satisfies  $\mathbb{E}_{\mathbf{x} \sim \mathcal{Q}} \left\| \nabla \left( \log \frac{\mathbf{p}(\mathbf{x})}{\mathbf{q}(\mathbf{x})} \right) \right\|^2 < \infty$ ,  
then  $\text{KSD}_{\mathbf{p}}^2(\mathcal{Q}) = 0$  iff  $\mathbf{P} = \mathcal{Q}$ .

## KSD for discrete-valued variables

Discrete domains:  $\mathcal{X} = \{1, \dots, L\}^D$  with  $L \in \mathbb{N}$ .

The population KSD (discrete):

$$\text{KSD}_{\mathbf{p}}^2(\mathcal{Q}) = \mathbb{E}_{\mathbf{x}, \mathbf{x}' \sim \mathcal{Q}} h_{\mathbf{p}}(\mathbf{x}, \mathbf{x}')$$

where

$$h_{\mathbf{p}}(\mathbf{x}, \mathbf{x}') = \mathbf{s}_{\mathbf{p}}(\mathbf{x})^\top \mathbf{s}_{\mathbf{p}}(\mathbf{x}') k(\mathbf{x}, \mathbf{x}') - \mathbf{s}_{\mathbf{p}}(\mathbf{x})^\top k_2(\mathbf{x}, \mathbf{x}') \\ - \mathbf{s}_{\mathbf{p}}(\mathbf{x}')^\top k_1(\mathbf{x}, \mathbf{x}') + \text{tr}[k_{12}(\mathbf{x}, \mathbf{x}')]$$

$$k_1(\mathbf{x}, \mathbf{x}') = \Delta_{\mathbf{x}}^{-1} k(\mathbf{x}, \mathbf{x}'), \Delta_{\mathbf{x}}^{-1} \text{ is difference on } \mathbf{x}, \mathbf{s}_{\mathbf{p}}(\mathbf{x}) = \frac{\Delta \mathbf{p}(\mathbf{x})}{\mathbf{p}(\mathbf{x})}$$

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where

$$h_{\mathbf{p}}(x, x') = \mathbf{s}_{\mathbf{p}}(x)^\top \mathbf{s}_{\mathbf{p}}(x') k(x, x') - \mathbf{s}_{\mathbf{p}}(x)^\top k_2(x, x') \\ - \mathbf{s}_{\mathbf{p}}(x')^\top k_1(x, x') + \text{tr}[k_{12}(x, x')]$$

$$k_1(x, x') = \Delta_x^{-1} k(x, x'), \Delta_x^{-1} \text{ is difference on } x, \mathbf{s}_{\mathbf{p}}(x) = \frac{\Delta_{\mathbf{p}}(x)}{\mathbf{p}(x)}$$

**A discrete kernel:**  $k(x, x') = \exp(-d_H(x, x'))$ , where  $d_H(x, x') = D^{-1} \sum_{d=1}^D \mathbb{I}(x_d \neq x'_d)$ .



# KSD for discrete-valued variables

Discrete domains:  $\mathcal{X} = \{1, \dots, L\}^D$  with  $L \in \mathbb{N}$ .

The population KSD (discrete):

$$\text{KSD}_{\mathbf{p}}^2(\mathcal{Q}) = \mathbb{E}_{\mathbf{x}, \mathbf{x}' \sim \mathcal{Q}} h_{\mathbf{p}}(\mathbf{x}, \mathbf{x}')$$

where

$$h_{\mathbf{p}}(\mathbf{x}, \mathbf{x}') = \mathbf{s}_{\mathbf{p}}(\mathbf{x})^\top \mathbf{s}_{\mathbf{p}}(\mathbf{x}') k(\mathbf{x}, \mathbf{x}') - \mathbf{s}_{\mathbf{p}}(\mathbf{x})^\top k_2(\mathbf{x}, \mathbf{x}') \\ - \mathbf{s}_{\mathbf{p}}(\mathbf{x}')^\top k_1(\mathbf{x}, \mathbf{x}') + \text{tr}[k_{12}(\mathbf{x}, \mathbf{x}')]$$

$$k_1(\mathbf{x}, \mathbf{x}') = \Delta_{\mathbf{x}}^{-1} k(\mathbf{x}, \mathbf{x}'), \Delta_{\mathbf{x}}^{-1} \text{ is difference on } \mathbf{x}, \mathbf{s}_{\mathbf{p}}(\mathbf{x}) = \frac{\Delta \mathbf{p}(\mathbf{x})}{\mathbf{p}(\mathbf{x})}$$

**A discrete kernel:**  $k(\mathbf{x}, \mathbf{x}') = \exp(-d_H(\mathbf{x}, \mathbf{x}'))$ , where  
 $d_H(\mathbf{x}, \mathbf{x}') = D^{-1} \sum_{d=1}^D \mathbb{I}(x_d \neq x'_d)$ .

$\text{KSD}_{\mathbf{p}}^2(\mathcal{Q}) = 0$  iff  $\mathbf{P} = \mathcal{Q}$  if

- Gram matrix over all the configurations in  $\mathcal{X}$  is strictly positive definite,
- $\mathbf{P} > 0$  and  $\mathcal{Q} > 0$ .

## Constructing threshold for a statistical test

Given samples  $\{z_i\}_{i=1}^n \sim q$ , empirical KSD (test statistic) is:

$$\widehat{\text{KSD}}(\mathbf{p}, q, \mathcal{F}) := \frac{1}{n(n-1)} \sum_{i=1}^n \sum_{j \neq i}^n h_{\mathbf{p}}(z_i, z_j).$$

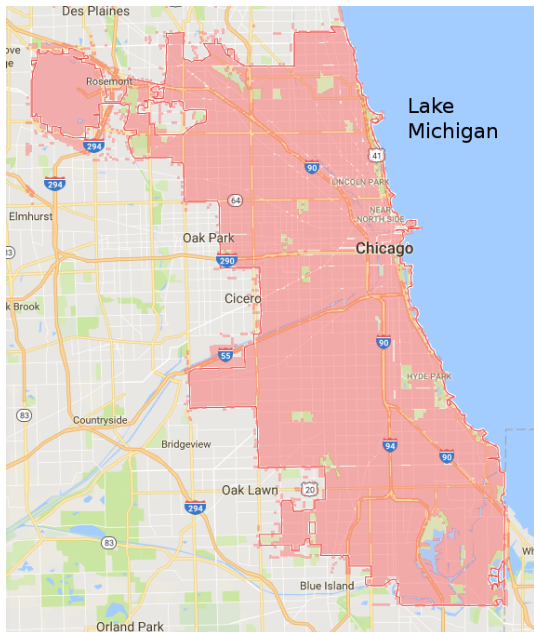
When  $q = \mathbf{p}$ , U-statistic is degenerate. Estimate of null distribution with wild bootstrap:

$$\widetilde{\text{KSD}}(\mathbf{p}, q, \mathcal{F}) := \frac{1}{n(n-1)} \sum_{i=1}^n \sum_{j \neq i}^n \sigma_i \sigma_j h_{\mathbf{p}}(z_i, z_j).$$

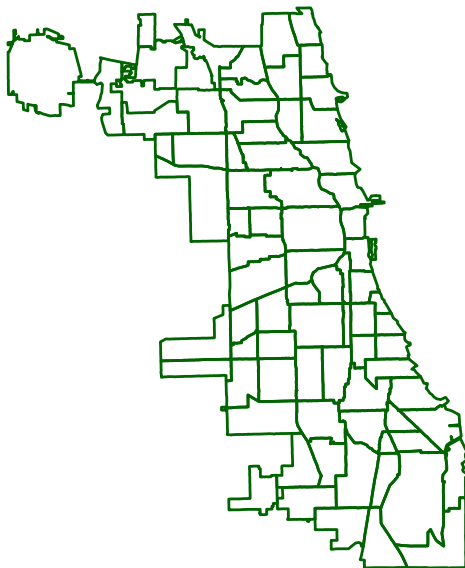
where  $\{\sigma_i\}_{i=1}^n$  i.i.d,  $E(\sigma_i) = 0$ , and  $E(\sigma_i^2) = 1$

- Consistent estimate of the null distribution when  $q = \mathbf{p}$
- Consistent test (Type II error goes to zero) under a rich class of alternatives Chwialkowski, Strathmann, G., ICML 2016

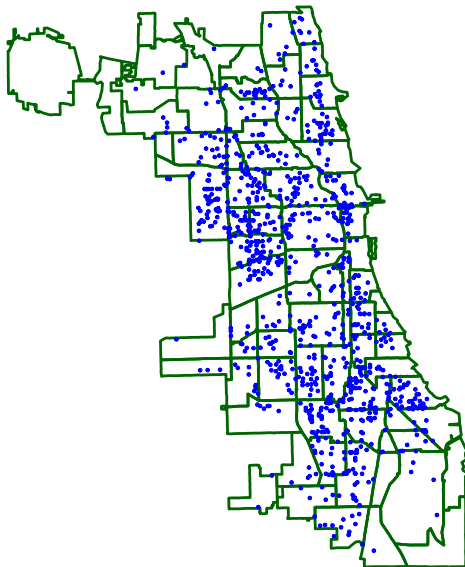
# Model Criticism



# Model Criticism

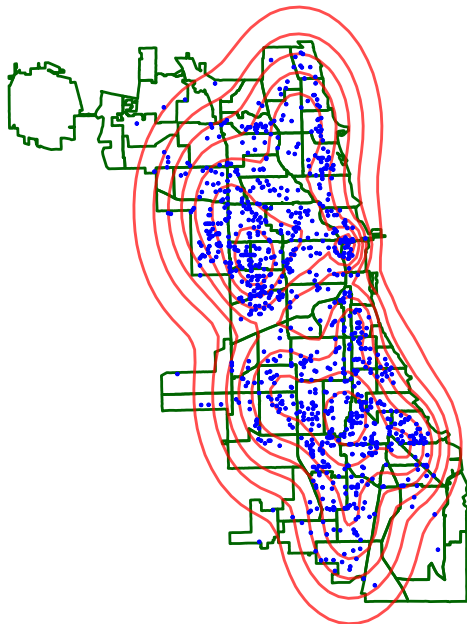


# Model Criticism



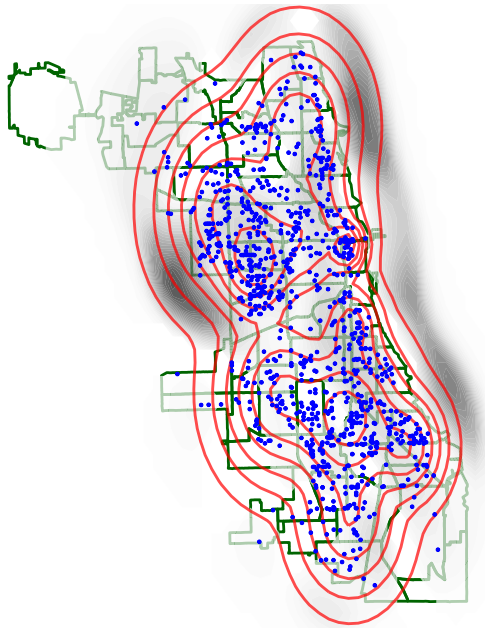
Data = robbery events in  
Chicago in 2016.

## The witness function: Chicago Crime



Model  $p$  = 10-component  
Gaussian mixture.

## The witness function: Chicago Crime



Witness function  $g$  shows mismatch

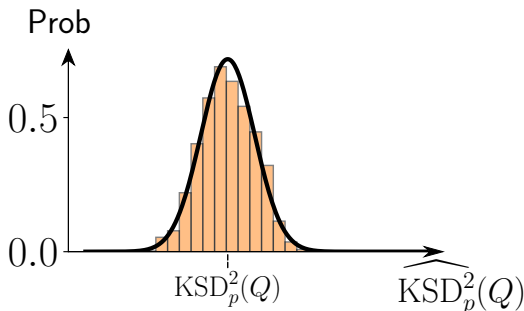
## Empirical statistic, asymptotic normality for $P \neq Q$

The empirical statistic:

$$\widehat{\text{KSD}}^2_p(Q) := \frac{1}{n(n-1)} \sum_{i \neq j} h_p(x_i, x_j).$$

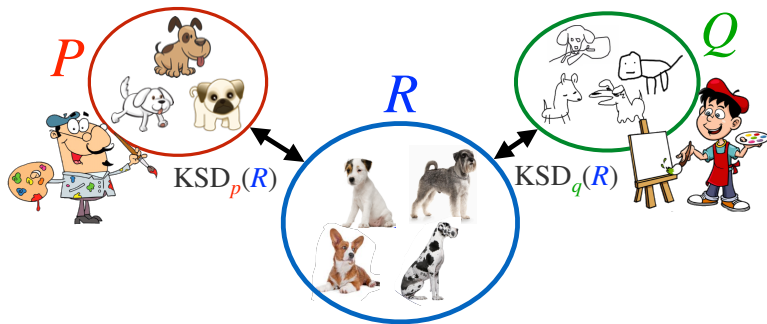
Asymptotic distribution when  $q \neq p$ :

$$\sqrt{n} \left( \widehat{\text{KSD}}^2_p(Q) - \text{KSD}_p(Q) \right) \xrightarrow{d} \mathcal{N}(0, \sigma_{h_p}^2) \quad \sigma_{h_p}^2 = 4 \text{Var}[\mathbb{E}_{x'}[h_p(x, x')]].$$





# Relative goodness-of-fit testing



- Two latent variable models  $P$  and  $Q$ , data  $\{x_i\}_{i=1}^n \stackrel{\text{i.i.d.}}{\sim} R$ .
- Distinct models  $p \neq q$

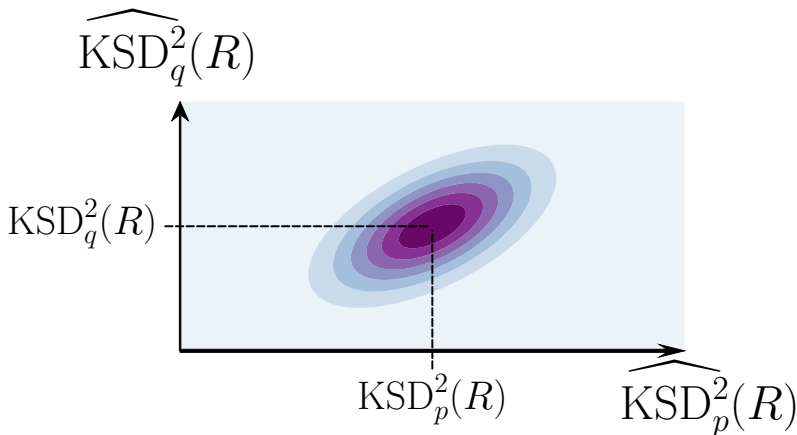
Hypotheses:

$$H_0 : KSD_p(R) \leq KSD_q(R) \text{ vs. } H_1 : KSD_p(R) > KSD_q(R) \\ (H_0 : 'P \text{ is as good as } Q, \text{ or better}' \text{ vs. } H_1 : 'Q \text{ is better}') )$$

## Relative GOF testing: joint asymptotic normality

Joint asymptotic normality when  $P \neq R$  and  $Q \neq R$

$$\sqrt{n} \begin{bmatrix} \widehat{\text{KSD}}_P^2(R) - \text{KSD}_P(R) \\ \widehat{\text{KSD}}_Q^2(R) - \text{KSD}_Q(R) \end{bmatrix} \xrightarrow{d} \mathcal{N} \left( \begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} \sigma_{h_P}^2 & \sigma_{h_P h_Q} \\ \sigma_{h_P h_Q} & \sigma_{h_Q}^2 \end{bmatrix} \right)$$



## Relative GOF testing: joint asymptotic normality

Joint asymptotic normality when  $P \neq R$  and  $Q \neq R$

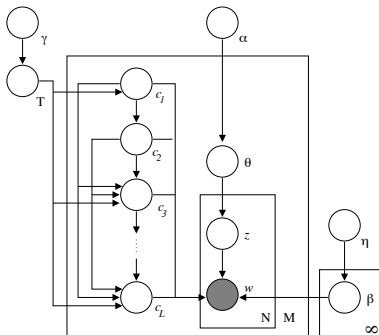
$$\sqrt{n} \begin{bmatrix} \widehat{\text{KSD}}_{\textcolor{red}{p}}^2(\textcolor{blue}{R}) - \text{KSD}_{\textcolor{red}{p}}(\textcolor{blue}{R}) \\ \widehat{\text{KSD}}_{\textcolor{teal}{q}}^2(\textcolor{blue}{R}) - \text{KSD}_{\textcolor{teal}{q}}(\textcolor{blue}{R}) \end{bmatrix} \xrightarrow{d} \mathcal{N} \left( \begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} \sigma_{h_{\textcolor{red}{p}}}^2 & \sigma_{h_{\textcolor{red}{p}}h_{\textcolor{teal}{q}}} \\ \sigma_{h_{\textcolor{red}{p}}h_{\textcolor{teal}{q}}} & \sigma_{h_{\textcolor{teal}{q}}}^2 \end{bmatrix} \right)$$

**Difference** in statistics is asymptotically normal:

$$\begin{aligned} \sqrt{n} \left[ \widehat{\text{KSD}}_{\textcolor{red}{p}}^2(\textcolor{blue}{R}) - \widehat{\text{KSD}}_{\textcolor{teal}{q}}^2(\textcolor{blue}{R}) - (\text{KSD}_{\textcolor{red}{p}}(\textcolor{blue}{R}) - \text{KSD}_{\textcolor{teal}{q}}(\textcolor{blue}{R})) \right] \\ \xrightarrow{d} \mathcal{N} \left( 0, \sigma_{h_{\textcolor{red}{p}}}^2 + \sigma_{h_{\textcolor{teal}{q}}}^2 - 2\sigma_{h_{\textcolor{red}{p}}h_{\textcolor{teal}{q}}} \right) \end{aligned}$$

$\implies$  a statistical test with **null hypothesis**  $\text{KSD}_{\textcolor{red}{p}}(\textcolor{blue}{R}) - \text{KSD}_{\textcolor{teal}{q}}(\textcolor{blue}{R}) \leq 0$  is straightforward.

# Latent variable models

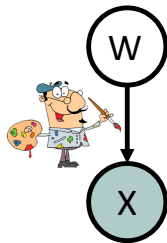
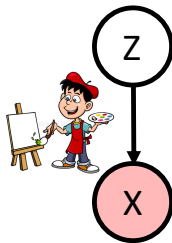


# Latent variable models

Can we compare latent variable models with KSD?

$$p(x) = \int p(x|z)p(z)dz$$

$$q(x) = \int q(x|w)p(w)dw$$



Multi-dimensional Stein operator:

$$[T_{\textcolor{red}{p}}f](x) = \left\langle f(x), \underbrace{\frac{\nabla \textcolor{red}{p}(x)}{\textcolor{red}{p}(x)}}_{(\textcolor{brown}{a})} \right\rangle + \langle \nabla, f(x) \rangle.$$

Expression  $(\textcolor{brown}{a})$  requires **marginal  $p(x)$ , often intractable...**

## What not to do

Approximate the integral using  $\{z_j\}_{j=1}^m \sim p(z)$ :

$$\begin{aligned} p(x) &= \int p(x|z)p(z)dz \\ &\approx p_m(x) = \frac{1}{m} \sum_{j=1}^m p(x|z_j) \end{aligned}$$

Estimate KSD with approximate density:

$$\widehat{\text{KSD}}_{\textcolor{blue}{p}}^2(\textcolor{blue}{R}) \approx \widehat{\text{KSD}}_{\textcolor{blue}{p}_m}^2(\textcolor{blue}{R})$$

## What not to do

Approximate the integral using  $\{z_j\}_{j=1}^m \sim p(z)$ :

$$\begin{aligned} p(x) &= \int p(x|z)p(z)dz \\ &\approx p_m(x) = \frac{1}{m} \sum_{j=1}^m p(x|z_j) \end{aligned}$$

Estimate KSD with approximate density:

$$\widehat{\text{KSD}}_p^2(R) \approx \widehat{\text{KSD}}_{p_m}^2(R)$$

**Problem:**  $\widehat{\text{KSD}}_{p_m}^2(R)$  asymptotically normal but slow bias decay.

# MCMC approximation of score function

Result we use:

$$\mathbf{s}_{\mathbf{p}}(\mathbf{x}) = \mathbb{E}_{z|x}[\mathbf{s}_{\mathbf{p}}(\mathbf{x}|z)]$$

Proof:

$$\begin{aligned}\mathbf{s}_{\mathbf{p}}(\mathbf{x}) &= \frac{\nabla \mathbf{p}(\mathbf{x})}{\mathbf{p}(\mathbf{x})} = \frac{1}{\mathbf{p}(\mathbf{x})} \int \nabla \mathbf{p}(\mathbf{x}|z) d\mathbf{p}(z) \\ &= \int \frac{\nabla \mathbf{p}(\mathbf{x}|z)}{\mathbf{p}(\mathbf{x}|z)} \cdot \frac{\mathbf{p}(\mathbf{x}|z) d\mathbf{p}(z)}{\mathbf{p}(\mathbf{x})} = \mathbb{E}_{z|x}[\mathbf{s}_{\mathbf{p}}(\mathbf{x}|z)],\end{aligned}$$

Friel, N., Mira, A. and Oates, C. J. (2016) Exploiting multi-core architectures for reduced-variance estimation with intractable likelihoods. *Bayesian Analysis*, 11, 215–245.



# MCMC approximation of score function

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Proof:

$$\begin{aligned}\mathbf{s}_{\textcolor{red}{p}}(x) &= \frac{\nabla \textcolor{red}{p}(x)}{\textcolor{red}{p}(x)} = \frac{1}{\textcolor{red}{p}(x)} \int \nabla \textcolor{red}{p}(x|z) dp(z) \\ &= \int \frac{\nabla \textcolor{red}{p}(x|z)}{\textcolor{red}{p}(x|z)} \cdot \frac{\textcolor{red}{p}(x|z) dp(z)}{\textcolor{red}{p}(x)} = \mathbb{E}_{z|x}[\mathbf{s}_{\textcolor{red}{p}}(x|z)],\end{aligned}$$

Approximate intractable posterior  $\mathbb{E}_{z|x_i}[\mathbf{s}_{\textcolor{red}{p}}(x_i|z)]$

$$\bar{\mathbf{s}}_{\textcolor{red}{p}}(x_i; z_i^{(t)}) := \frac{1}{m} \sum_{j=1}^m \mathbf{s}_{\textcolor{red}{p}}(x_i|z_{i,j}^{(t)}) \approx \mathbf{s}_{\textcolor{red}{p}}(x_i)$$

with  $z_i^{(t)} = (z_{i,1}^{(t)}, \dots, z_{i,m}^{(t)})$  via **MCMC** (after  $t$  burn-in steps)

Friel, N., Mira, A. and Oates, C. J. (2016) Exploiting multi-core architectures for reduced-variance estimation with intractable likelihoods. Bayesian Analysis, 11, 215–245.

## KSD for latent variable models

Recall earlier KSD estimate:

$$U_n(\textcolor{red}{P}) = \frac{1}{n(n-1)} \sum_{i \neq j} h_{\textcolor{red}{p}}(x_i, x_j) \ (\approx \text{KSD}_{\textcolor{red}{p}}^2(\textcolor{blue}{R}))$$

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KSD estimate for latent variable models:

$$U_n^{(t)}(\textcolor{red}{P}) := \frac{1}{n(n-1)} \sum_{i \neq j} \bar{H}_{\textcolor{red}{p}}[(x_i, z_i^{(t)}), (x_j, z_j^{(t)})] \ (\approx \text{KSD}_{\textcolor{red}{p}}^2(\textcolor{blue}{R}))$$

where  $\bar{H}_{\textcolor{red}{p}}$  is the Stein kernel  $h_{\textcolor{red}{p}}$  with  $s_{\textcolor{red}{p}}(x_i)$  replaced with  $\bar{s}_{\textcolor{red}{p}}(x_i; z_i^{(t)})$ .

## Return to relative GOF test, latent variable models

Hypotheses:

$$H_0 : \text{KSD}_{\textcolor{red}{p}}(\textcolor{blue}{R}) \leq \text{KSD}_{\textcolor{teal}{q}}(\textcolor{blue}{R}) \text{ vs. } H_1 : \text{KSD}_{\textcolor{red}{p}}(\textcolor{blue}{R}) > \text{KSD}_{\textcolor{teal}{q}}(\textcolor{blue}{R})$$

(  $H_0$  : ' $\textcolor{red}{P}$  is as good as  $\textcolor{teal}{Q}$ , or better' vs.  $H_1$  : ' $\textcolor{teal}{Q}$  is better' )

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( $H_0$  : ' $\textcolor{red}{P}$  is as good as  $\textcolor{teal}{Q}$ , or better' vs.  $H_1$  : ' $\textcolor{teal}{Q}$  is better' )

Strategy:

- Estimate the difference  $\text{KSD}_{\textcolor{red}{p}}^2(\textcolor{blue}{R}) - \text{KSD}_{\textcolor{teal}{q}}^2(\textcolor{blue}{R})$  by

$$D_n^{(t)}(\textcolor{red}{P}, \textcolor{teal}{Q}) = U_n^{(t)}(\textcolor{red}{P}) - U_n^{(t)}(\textcolor{teal}{Q}).$$

- If  $D_n^{(t)}(\textcolor{red}{P}, \textcolor{teal}{Q})$  is sufficiently large, reject  $H_0$ .
  - “Sufficient”: control type-I error (falsely rejecting  $H_0$ )
  - Requires the (asymptotic) behaviour of  $D_n^{(t)}(\textcolor{red}{P}, \textcolor{teal}{Q})$

## Asymptotic distribution for relative KSD test

Asymptotic distribution of approximate KSD estimate  $n, t \rightarrow \infty$ :

$$\sqrt{n} \left[ D_n^{(t)}(\textcolor{red}{P}, \textcolor{teal}{Q}) - \mu_{\textcolor{red}{P}\textcolor{teal}{Q}} \right] \xrightarrow{d} \mathcal{N}(0, \sigma_{\textcolor{red}{P}\textcolor{teal}{Q}}^2)$$

where

$$\begin{aligned} \mu_{\textcolor{red}{P}\textcolor{teal}{Q}} &= \text{KSD}_{\textcolor{red}{p}}^2(\textcolor{blue}{R}) - \text{KSD}_{\textcolor{teal}{q}}^2(\textcolor{blue}{R}), \\ \sigma_{\textcolor{red}{P}\textcolor{teal}{Q}}^2 &= \lim_{n, t \rightarrow \infty} n \cdot \text{Var} \left[ D_n^{(t)}(\textcolor{red}{P}, \textcolor{teal}{Q}) \right]. \end{aligned}$$

Fine print:

- The double limit requires fast bias decay

$$\sqrt{n} [\mathbb{E}\{D_n^{(t)}(\textcolor{red}{P}, \textcolor{teal}{Q})\} - \mu_{\textcolor{red}{P}\textcolor{teal}{Q}}] \rightarrow 0$$

- The fourth moment of  $\bar{H}_{\textcolor{red}{p}}^{(t)} - \bar{H}_{\textcolor{teal}{q}}^{(t)}$  has finite limit sup. ( $t \rightarrow \infty$ ).

# Asymptotic distribution for relative KSD test

Asymptotic distribution of approximate KSD estimate  $n, t \rightarrow \infty$ :

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where

$$\mu_{\textcolor{red}{P}\textcolor{teal}{Q}} = \text{KSD}_{\textcolor{red}{p}}^2(\textcolor{blue}{R}) - \text{KSD}_{\textcolor{teal}{q}}^2(\textcolor{blue}{R}),$$

$$\sigma_{\textcolor{red}{P}\textcolor{teal}{Q}}^2 = \lim_{n, t \rightarrow \infty} n \cdot \text{Var} \left[ D_n^{(t)}(\textcolor{red}{P}, \textcolor{teal}{Q}) \right].$$

Level- $\alpha$  test:

$$\text{Reject } H_0 \text{ if } D_n^{(t)}(\textcolor{red}{P}, \textcolor{teal}{Q}) \geq \frac{\hat{\sigma}_{\textcolor{red}{P}\textcolor{teal}{Q}}}{\sqrt{n}} c_{1-\alpha}$$

- $c_{1-\alpha}$  is  $(1 - \alpha)$ -quantile of  $\mathcal{N}(0, 1)$ .
- $\hat{\sigma}_{\textcolor{red}{P}\textcolor{teal}{Q}}$  estimated via jackknife

# Experiments



## Experiment 1: sensitivity to model difference

- Data  $R$  : Probabilistic Principal Component Analysis PPCA( $A$ ):

$$x_i \in \mathbb{R}^{100} \sim \mathcal{N}(Az_i, I), \quad z_i \in \mathbb{R}^{10} \sim \mathcal{N}(0, I_z)$$

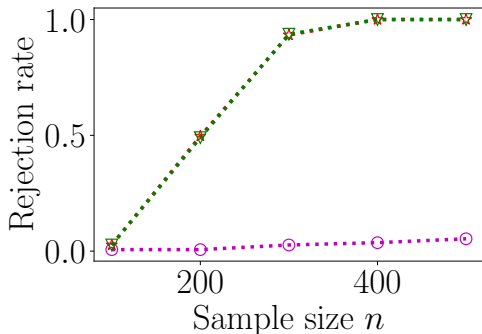
- Generate  $P$ ,  $Q$  : perturb (1, 1)-entry :  $A_\delta = A + \delta E_{1,1}$

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- Generate  $P$ ,  $Q$  : perturb (1, 1)-entry :  $A_\delta = A + \delta E_{1,1}$



- Alt.  $H_1$  ( $Q$  is better):

- $P$ 's perturbation  $\delta_P = 2$
- $Q$ 's perturbation  $\delta_Q = 1$

- IMQ kernel:  $k(x, x') = (1 + \|x - x'\|_2^2 / \sigma_{\text{med}}^2)^{-1/2}$

- NUTS-HMC with sample size  $m = 500$  (after  $t = 200$  steps).

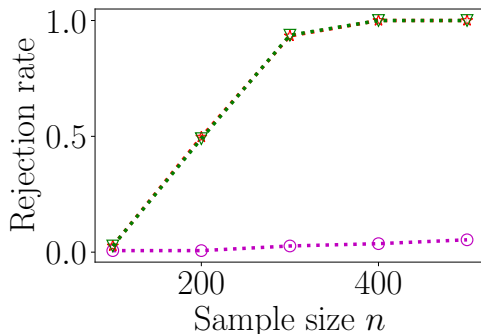
.....○..... MMD      .....☆..... KSD      .....▽..... LKSD

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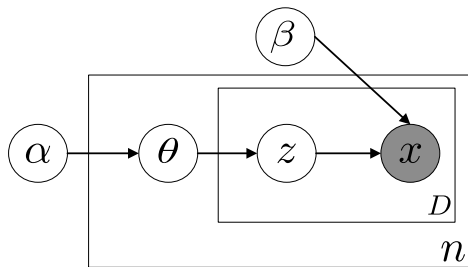
(L)KSD = higher power

- Sample-wise difference in models = subtle (MMD fails)
- Model information is helpful

.....○..... MMD      .....☆..... KSD      .....▽..... LKSD

## Experiment 2: topic models for arXiv articles

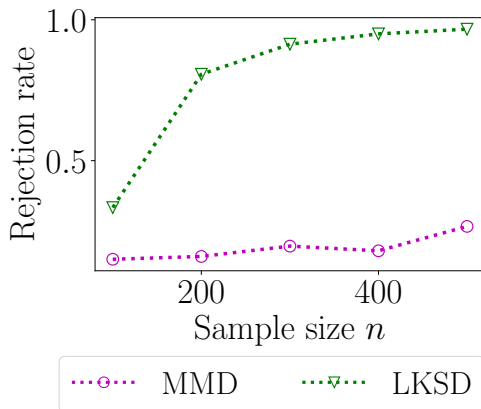
- Data  $R$  : arXiv articles from category stat.TH (stat theory) :
- Models  $P$ ,  $Q$  : LDAs trained on articles from different categories
  - $P$  : math.PR (math probability theory)
  - $Q$  : stat.ME (stat methodology).  $H_1$ :  $Q$  is better



Graphical model of LDA

## Experiment 2: topic models for arXiv articles

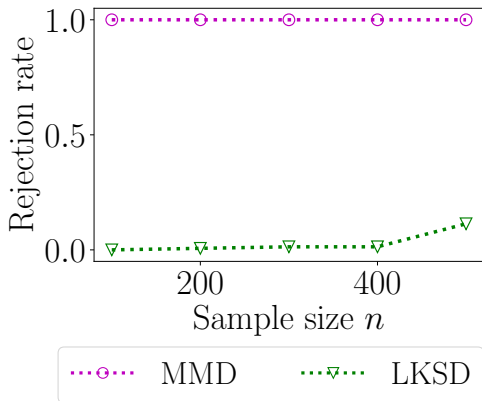
- Data  $R$ : arXiv articles from category stat.TH (stat theory):
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  - $P$ : math.PR (math probability theory)
  - $Q$ : stat.ME (stat methodology).  $H_1$ :  $Q$  is better



- $\mathcal{X} = \{1, \dots, L\}^D$ ,  $D = 100$ ,  $L = 126, 190$ .
- IMQ kernel in BoW rep.:
$$k(x, x') = (1 + \|B(x) - B(x')\|_2^2)^{-1/2}$$
- MCMC size  $m = 5000$  (after  $t = 500$  steps).

## A failure mode

- Data  $R$  : arXiv articles from category stat.TH (stat theory) :
- Models  $P$ ,  $Q$  : LDAs trained on articles from different categories (100 topics)
  - $P$  : cs.LG (CS machine learning)
  - $Q$  : stat.ME (stat methodology).  $H_1$ :  $Q$  is better



- $\mathcal{X} = \{1, \dots, L\}^D$ ,  $D = 100$ ,  $L = 208,671$ .
- IMQ kernel in BoW rep.:
$$k(x, x') = (1 + \|B(x) - B(x')\|_2^2)^{-1/2}$$
- MCMC size  $m = 5000$  (after  $t = 500$  steps).

## What went wrong?

Recall (one-dimension, informally)

$$s_p(x) = \frac{p(x+1)}{p(x)} - 1$$

Numerical instability arises when

- Observed word  $x$  has low probability
- Word next to  $x$  in vocabulary has non-negligible probability

# Zanella-Barker Stein operator

Zanella-Barker Stein operator (1-D):

$$\mathcal{A}_p^{\text{ZB}} f(x) = \sum_{\tilde{x} \in \{x+1, x-1\}} \frac{p(\tilde{x})}{p(\tilde{x}) + p(x)} \cdot \{f(\tilde{x}) - f(x)\}$$

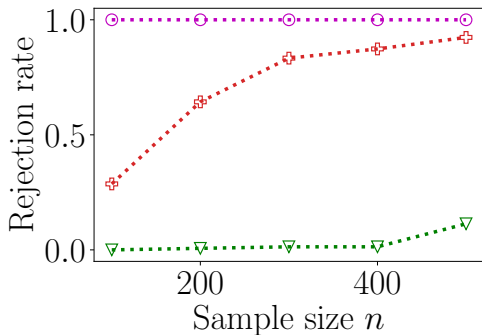
- More stable: the ratio  $p(\tilde{x})/\{p(\tilde{x}) + p(x)\}$  is always between 0 and 1.
- Similarly applies to latent variable models.

Hodgkinson, Salomone, and Roosta (2020); Shi, Zhou, Hwang, Titsias, and Mackey. (2022)



# A resolution to the failure mode

- Data  $R$  : arXiv articles from category stat.TH (stat theory) :
- Models  $P$ ,  $Q$  : LDAs trained on articles from different categories (100 topics)
  - $P$  : cs.LG (CS machine learning)
  - $Q$  : stat.ME (stat methodology).  $H_1$ :  $Q$  is better



■ Improved performance by an alternative Stein operator

.....○..... MMD      .....▽..... LKSD      .....+..... LKSD (Alt.)

## Can sampler influence test power?

How important is the quality of  $\frac{1}{m} \sum_{j=1}^m s_p(x|z_j^{(t)})$ ?

Experiment with PPCA:

- $P$  : MALA with a bad step size (poor sampler)
- $Q$  : NUTS-HMC (good sampler)

Expectation:

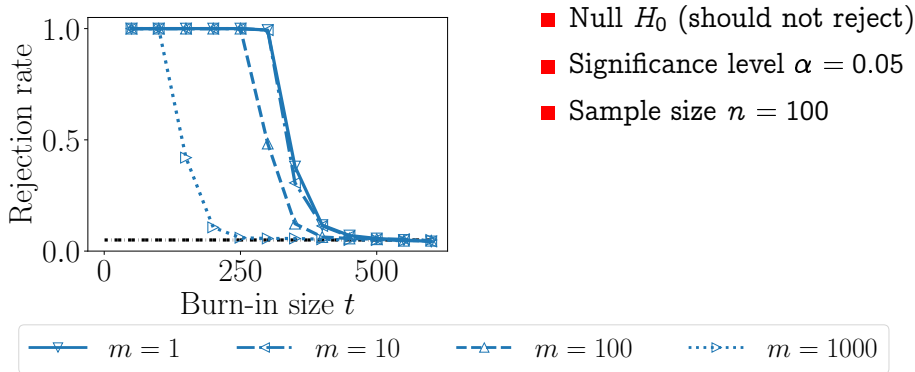
If poor, the test would reject even if  $P$  and  $Q$  are equally good

# Can sampler influence test power?

How important is the quality of  $\frac{1}{m} \sum_{j=1}^m s_p(x|z_j^{(t)})$ ?

Experiment with PPCA:

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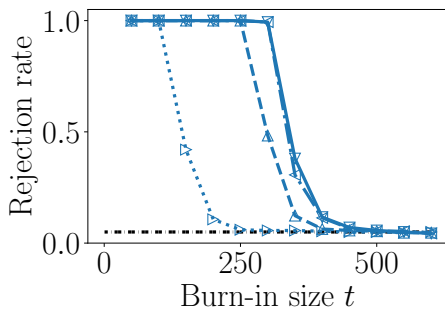


# Can sampler influence test power?

How important is the quality of  $\frac{1}{m} \sum_{j=1}^m \mathbf{s}_{\textcolor{red}{P}}(x|z_j^{(t)})$ ?

Experiment with PPCA:

- $\textcolor{red}{P}$  : MALA with a bad step size (poor sampler)
- $\textcolor{teal}{Q}$  : NUTS-HMC (good sampler)



- Null  $H_0$  (should not reject)
- Significance level  $\alpha = 0.05$
- Sample size  $n = 100$

Sufficient burn-in  
→ correct type-I error

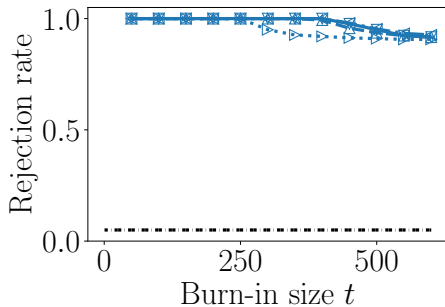
$\textcolor{teal}{\nabla}$  —  $m = 1$        $\textcolor{teal}{\triangleleft}$  - -  $m = 10$       - -  $\triangle$   $m = 100$        $\cdots \triangleright \cdots$   $m = 1000$

## Can sampler influence test power?

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- Significance level  $\alpha = 0.05$
- Sample size  $n = 300$

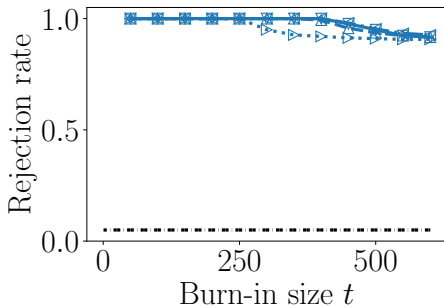
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## Can sampler influence test power?

How important is the quality of  $\frac{1}{m} \sum_{j=1}^m \mathbf{s}_P(x|z_j^{(t)})$ ?

Experiment with PPCA:

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Larger  $n \implies$  more  
sensitive to mismatch

—▽—  $m = 1$       -◀-  $m = 10$       -△-  $m = 100$       ...▶...  $m = 1000$

# Conclusion

## Relative goodness-of-fit tests for Models with Latent Variables

- The kernel Stein discrepancy
  - Comparing two models via samples: MMD and the witness function.
  - Comparing a sample and a model: Stein modification of the witness class
- Constructing a relative hypothesis test using the KSD
- Relative hypothesis tests with latent variables

## References

A Kernel Test of Goodness of Fit

Kacper Chwialkowski, Heiko Strathmann, Arthur Gretton

<https://arxiv.org/abs/1602.02964>

A Kernel Stein Test for Comparing Latent Variable Models

Heishiro Kanagawa, Wittawat Jitkrittum, Lester Mackey,

Kenji Fukumizu, Arthur Gretton

<https://arxiv.org/abs/1907.00586>



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The Gatsby Charitable Foundation



Deepmind



# Questions?



## KSD Riesz condition proof (detailed)

The KSD is written:

$$\begin{aligned}[T_p f](z) &= \left( \frac{d}{dz} \log p(z) \right) f(z) + \frac{d}{dz} f(z) \\ &= \left\langle f, \left( \frac{d}{dz} \log p(z) \right) k(z, \cdot) + \frac{d}{dz} k(z, \cdot) \right\rangle_{\mathcal{F}} \\ &=: \langle f, \xi_z \rangle_{\mathcal{F}} .\end{aligned}$$

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Step 2: show that

$$E_{z \sim q} [T_p f] = E_{z \sim q} \langle f, \xi_z \rangle_{\mathcal{F}} = \langle f, E_{z \sim q} \xi_z \rangle_{\mathcal{F}}.$$

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Riesz theorem!

## Next step: taking expectations

Riesz theorem: need boundedness,

$$|E_{z \sim q} \langle f, \xi_z \rangle_{\mathcal{F}}| \leq \|f\|_{\mathcal{F}} \lambda$$

for some  $\lambda \in \mathbb{R}$ .

By Jensen and Cauchy-Schwarz,

$$\begin{aligned} |E_{z \sim q} \langle f, \xi_z \rangle_{\mathcal{F}}| &\leq E_{z \sim q} |\langle f, \xi_z \rangle_{\mathcal{F}}| \\ &\leq \|f\|_{\mathcal{F}} \underbrace{E_{z \sim q} \|\xi_z\|_{\mathcal{F}}}_{\text{bounded?}}. \end{aligned}$$

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Compute the squared norm:

$$\begin{aligned}\|\xi_z\|_{\mathcal{F}}^2 &= \langle \xi_z, \xi_z \rangle_{\mathcal{F}} \\&= \left\langle \left( \frac{d}{dz} \log p(z) \right) k(z, \cdot) + \|\xi_z\|_{\mathcal{F}}^2 \|\xi_z\|_{\mathcal{F}}^2 \|\xi_z\|_{\mathcal{F}}^2 \frac{d}{dz} k(z, \cdot), \dots \right\rangle_{\mathcal{F}} \\&= \underbrace{\left\langle \left( \frac{d}{dz} \log p(z) \right) k(z, \cdot), \left( \frac{d}{dz} \log p(z) \right) k(z, \cdot) \right\rangle_{\mathcal{F}}}_{(A)} \\&\quad + \underbrace{\left\langle \frac{d}{dx} k(x, \cdot), \frac{d}{dx'} k(x', \cdot) \right\rangle_{\mathcal{F}} \Big|_{x=x'=z}}_{(B) = \frac{d}{dx} \frac{d}{dx'} k(x-x') \Big|_{x=x'=z}} \\&\quad + 2 \underbrace{\left\langle \left( \frac{d}{dx} \log p(x) \right) k(x, \cdot), \frac{d}{dx'} k(x', \cdot) \right\rangle_{\mathcal{F}} \Big|_{x=x'=z}}_{(C)}\end{aligned}$$



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## First two (easy) terms

First term (A):

$$\begin{aligned}(A) &= \left\langle \left( \frac{d}{dz} \log p(z) \right) k(z, \cdot), \left( \frac{d}{dz} \log p(z) \right) k(z, \cdot) \right\rangle_{\mathcal{F}} \\ &= \left[ \left( \frac{d}{dz} \log p(z) \right)^2 \underbrace{k(z, z)}_{=c} \right]\end{aligned}$$

## First two (easy) terms

Second term (B):

$$\begin{aligned}(B) &= \left\langle \frac{d}{dx} k(x, \cdot), \frac{d}{dx'} k(x', \cdot) \right\rangle_{\mathcal{F}} \Big|_{x=x'=z} \\&= \sum_{\ell=-\infty}^{\infty} \frac{[-i\ell \hat{k}_{\ell} \exp(-i\ell x)] \overline{[-i\ell \hat{k}_{\ell} \exp(-i\ell x')]} }{\hat{k}_{\ell}} \Big|_{x=x'=z} \\&= \sum_{\ell=-\infty}^{\infty} -(i\ell)^2 \hat{k}_{\ell} \underbrace{\exp(i\ell(x' - x))}_{=1 \text{ when } x=x'=z} \\&= \sum_{\ell=-\infty}^{\infty} \ell^2 \hat{k}_{\ell} =: C > 0\end{aligned}$$

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Third term (C):

$$\begin{aligned}
 (C) &= \left\langle \left( \frac{d}{dx} \log p(x) \right) k(x, \cdot), \frac{d}{dx'} k(x', \cdot) \right\rangle_{\mathcal{F}} \Big|_{x=x'=z} \\
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## Putting it all together

We found:

$$\|\xi_z\|_{\mathcal{F}}^2 = C + \left( \frac{d}{dz} \log p(z) \right)^2 c,$$

Thus for boundedness, we have the condition:

$$\begin{aligned} E_{z \sim q} \|\xi_z\|_{\mathcal{F}} &= E_{z \sim q} \sqrt{C + \left( \frac{d}{dx} \log p(x) \right)^2 c} \\ &\leq \sqrt{E_{z \sim q} \left[ C + \left( \frac{d}{dz} \log p(z) \right)^2 c \right]}, \end{aligned}$$

So Riesz holds when  $E_{z \sim q} \left( \frac{d}{dz} \log p(z) \right)^2 < \infty$

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