### **Probabilistic & Unsupervised Learning**

# Factored Variational Approximations and Variational Bayes

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### **Expectations in Statistical Modelling**

Parameter estimation

$$\begin{split} \hat{\theta} &= \operatorname*{argmax}_{\theta} \int \frac{d\mathcal{Z}}{P(\mathcal{Z}|\theta)} P(\mathcal{X}|\mathcal{Z},\theta) \\ \text{(or, using EM)} \\ \theta^{\text{new}} &= \operatorname*{argmax}_{\theta} \int \frac{d\mathcal{Z}}{P(\mathcal{Z}|\mathcal{X},\theta^{\text{old}})} \log P(\mathcal{X},\mathcal{Z}|\theta) \end{split}$$

Prediction

$$p(x|\mathcal{D}, m) = \int d\theta \ p(\theta|\mathcal{D}, m)p(x|\theta, \mathcal{D}, m)$$

Model selection or weighting (by marginal likelihood)

$$p(\mathcal{D}|m) = \int d\theta \ p(\theta|m)p(\mathcal{D}|\theta, m)$$

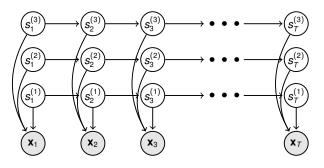
#### These integrals are often intractable:

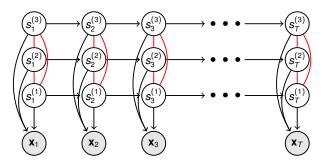
- ► Analytic intractability: integrals may not have closed form in non-linear, non-Gaussian models ⇒ numerical integration.
- Computational intractability: Numerical integral (or sum if Z or θ are discrete) may be exponential in data or model size.

### Intractabilities and approximations

- Inference computational intractability
  - Factored variational approx
  - Loopy BP/EP/Power EP
  - LP relaxations/ convexified BP
  - Gibbs sampling, other MCMC
- Inference analytic intractability
  - Laplace approximation (global)
    - Parametric variational approx
    - Message approximations (linearised, sigma-point, Laplace)
    - Assumed-density methods and Expectation-Propagation
    - (Sequential) Monte-Carlo methods
- Learning intractable partition function
  - Sampling parameters
  - Constrastive divergence
  - Score-matching
- Model selection
  - Laplace approximation / BIC
  - Variational Bayes
  - (Annealed) importance sampling
  - Reversible jump MCMC

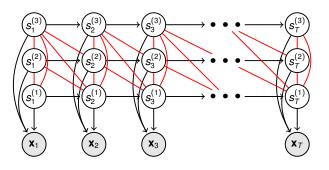
Not a complete list!



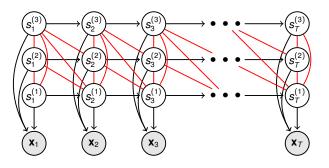


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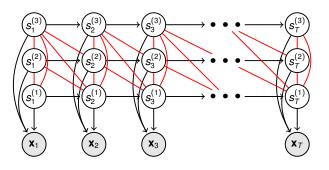
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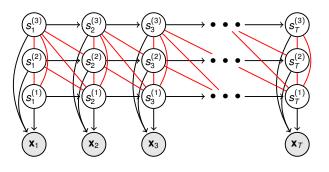
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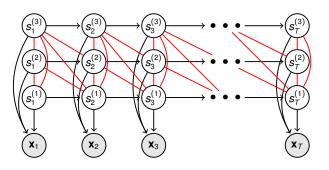
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To see how they work, we need to review the free-energy interpretation of EM.

### The Free Energy for a Latent Variable Model

Observed data  $\mathcal{X} = \{\mathbf{x}_i\}$ ; Latent variables  $\mathcal{Z} = \{\mathbf{z}_i\}$ ; Parameters  $\theta$ .

**Goal:** Maximize the log likelihood wrt  $\theta$  (i.e. ML learning):

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$$\begin{split} \int q(\mathcal{Z}) \log \frac{P(\mathcal{Z}, \mathcal{X}|\theta)}{q(\mathcal{Z})} \ d\mathcal{Z} &= \int q(\mathcal{Z}) \log P(\mathcal{Z}, \mathcal{X}|\theta) \ d\mathcal{Z} - \int q(\mathcal{Z}) \log q(\mathcal{Z}) \ d\mathcal{Z} \\ &= \int q(\mathcal{Z}) \log P(\mathcal{Z}, \mathcal{X}|\theta) \ d\mathcal{Z} + \mathbf{H}[q], \end{split}$$

where  $\mathbf{H}[q]$  is the entropy of  $q(\mathcal{Z})$ .

So: 
$$\mathcal{F}(q,\theta) = \langle \log P(\mathcal{Z}, \mathcal{X}|\theta) \rangle_{q(\mathcal{Z})} + \mathbf{H}[q]$$

### The E and M steps of EM

The log likelihood is bounded below by:

$$\mathcal{F}(q,\theta) = \langle \log P(\mathcal{Z},\mathcal{X}|\theta) \rangle_{q(\mathcal{Z})} + \mathbf{H}[q] = \ell(\theta) - \mathbf{KL}[q(\mathcal{Z}) \| P(\mathcal{Z}|\mathcal{X},\theta)]$$

EM alternates between:

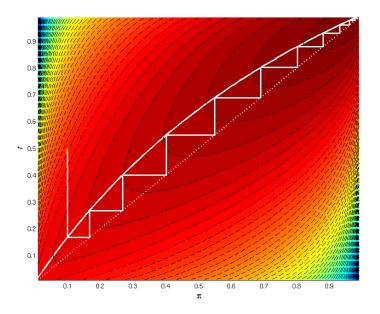
**E step:** optimise  $\mathcal{F}(q,\theta)$  wrt distribution over hidden variables holding parameters fixed:

$$q^{(k)}(\mathcal{Z}) := \underset{q(\mathcal{Z})}{\operatorname{argmax}} \ \mathcal{F}\big(q(\mathcal{Z}), \theta^{(k-1)}\big) = P\big(\mathcal{Z}|\mathcal{X}, \theta^{(k-1)}\big)$$

**M step:** maximise  $\mathcal{F}(q,\theta)$  wrt parameters holding hidden distribution fixed:

$$\theta^{(k)} := \underset{\theta}{\operatorname{argmax}} \ \mathcal{F}\big(\mathbf{q^{(k)}(\mathcal{Z})}, \theta\big) = \underset{\theta}{\operatorname{argmax}} \ \langle \log P(\mathcal{Z}, \mathcal{X} | \theta) \rangle_{q^{(k)}(\mathcal{Z})}$$

### EM as Coordinate Ascent in ${\mathcal F}$



#### **EM Never Decreases the Likelihood**

The E and M steps together never decrease the log likelihood:

$$\ell\big(\boldsymbol{\theta}^{(k-1)}\big) \ \underset{\text{E step}}{=} \ \mathcal{F}\big(\boldsymbol{q}^{(k)},\boldsymbol{\theta}^{(k-1)}\big) \underset{\text{M step}}{\leq} \ \mathcal{F}\big(\boldsymbol{q}^{(k)},\boldsymbol{\theta}^{(k)}\big) \underset{\text{Jensen}}{\leq} \ \ell\big(\boldsymbol{\theta}^{(k)}\big),$$

- The E step brings the free energy to the likelihood.
- ▶ The M-step maximises the free energy wrt  $\theta$ .
- $\mathcal{F} \leq \ell$  by Jensen or, equivalently, from the non-negativity of KL

If the M-step is executed so that  $\theta^{(k)} \neq \theta^{(k-1)}$  iff  $\mathcal{F}$  increases, then the overall EM iteration will step to a new value of  $\theta$  iff the likelihood increases.

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#### For the E-step we could:

- ▶ Parameterise  $q = q_{\rho}(\mathcal{Z})$  and take a gradient step in  $\rho$ .
- Assume some simplified form for q, usually factored:  $q = \prod_i q_i(\mathcal{Z}_i)$  where  $\mathcal{Z}_i$  partition  $\mathcal{Z}$ , and maximise within this form.

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In either case, we choose q from within a limited set Q:

**VE step**: maximise  $\mathcal{F}(q,\theta)$  wrt constrained latent distribution given parameters:

$$q^{(k)}(\mathcal{Z}) := \underset{q(\mathcal{Z}) \in \mathcal{Q} \leftarrow \text{Constraint}}{\operatorname{argmax}} \mathcal{F}(q(\mathcal{Z}), \theta^{(k-1)}).$$

M step: unchanged

$$\theta^{(k)} := \underset{\boldsymbol{\theta}}{\operatorname{argmax}} \ \mathcal{F}\big(\mathbf{q}^{(k)}(\boldsymbol{\mathcal{Z}}), \boldsymbol{\theta}\big) = \underset{\boldsymbol{\theta}}{\operatorname{argmax}} \ \int \mathbf{q}^{(k)}(\boldsymbol{\mathcal{Z}}) \log p(\boldsymbol{\mathcal{Z}}, \boldsymbol{\mathcal{X}}|\boldsymbol{\theta}) d\boldsymbol{\mathcal{Z}},$$

Unlike in GEM, the fixed point may not be at an unconstrained optimum of  $\mathcal{F}$ .

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$$\mathcal{F}(q,\theta) \leq \ell(\theta^{\mathsf{ML}})$$

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#### **KL** divergence

Recall that

$$\begin{split} \mathcal{F}(q,\theta) &= \langle \log P(\mathcal{X},\mathcal{Z}|\theta) \rangle_{q(\mathcal{Z})} + \mathbf{H}[q] \\ &= \langle \log P(\mathcal{X}|\theta) + \log P(\mathcal{Z}|\mathcal{X},\theta) \rangle_{q(\mathcal{Z})} - \langle \log q(\mathcal{Z}) \rangle_{q(\mathcal{Z})} \\ &= \langle \log P(\mathcal{X}|\theta) \rangle_{q(\mathcal{Z})} - \mathbf{KL}[q \| P(\mathcal{Z}|\mathcal{X},\theta)]. \end{split}$$

Thus,

E step maximise  $\mathcal{F}(q,\theta)$  wrt the distribution over latents, given parameters:

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is equivalent to:

E step minimise  $KL[q||p(\mathcal{Z}|\mathcal{X},\theta)]$  wrt distribution over latents, given parameters:

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So, in each E step, the algorithm is trying to find the best approximation to  $P(\mathcal{Z}|\mathcal{X})$  in  $\mathcal{Q}$  in a KL sense. This is related to ideas in *information geometry*. It also suggests generalisations to other distance measures.

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In this case the E-step is itself iterative:

(Factored VE step)<sub>i</sub>: maximise  $\mathcal{F}(q,\theta)$  wrt  $q_i(\mathcal{Z}_i)$  given other  $q_j$  and parameters:

$$q_i^{(k)}(\mathcal{Z}_i) := \underset{q_i(\mathcal{Z}_i)}{\operatorname{argmax}} \ \mathcal{F}\big(q_i(\mathcal{Z}_i) \prod_{j \neq i} q_j(\mathcal{Z}_j), \theta^{(k-1)}\big).$$

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- q<sub>i</sub> updates iterated to convergence to "complete" VE-step.
- In fact, every (VE)₁-step separately increases F, so any schedule of (VE)₁- and M-steps will converge. Choice can be dictated by practical issues (rarely efficient to fully converge E-step before updating parameters).

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The free energy is:

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$$(=0) \quad \Rightarrow \quad q_i(\mathcal{Z}_i) \propto \exp\left\langle \log P(\mathcal{X}, \mathcal{Z} | \theta^{(k-1)}) \right\rangle_{\prod_{j \neq i} q_j(\mathcal{Z}_j)}$$

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Now, taking the variational derivative of the Lagrangian (enforcing normalisation of  $q_i$ ):

$$\begin{split} &\frac{\delta}{\delta q_i} \left( \mathcal{F} + \lambda \left( \int q_i - 1 \right) \right) = \left\langle \log P(\mathcal{X}, \mathcal{Z} | \theta^{(k-1)}) \right\rangle_{\prod_{j \neq i} q_j(\mathcal{Z}_j)} - \log q_i(\mathcal{Z}_i) - \frac{q_i(\mathcal{Z}_i)}{q_i(\mathcal{Z}_i)} + \lambda \end{split}$$

$$(= 0) \quad \Rightarrow \quad q_i(\mathcal{Z}_i) \propto \exp \left\langle \log P(\mathcal{X}, \mathcal{Z} | \theta^{(k-1)}) \right\rangle_{\prod_{i \neq i} q_i(\mathcal{Z}_i)}$$

In general, this depends only on the expected sufficient statistics under  $q_i$ . Thus, again, we don't actually need the *entire* distributions, just the relevant expectations (now for approximate inference as well as learning).

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$$P(\mathcal{X}, \mathcal{Z}) = \frac{1}{Z} \exp \left( \sum_{ij} W_{ij} s_i s_j + \sum_{i} b_i s_i \right)$$

with some  $s_i \in \mathcal{Z}$  and others observed.

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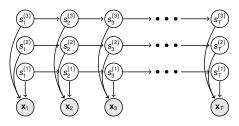
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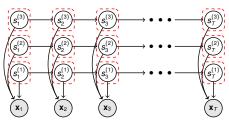
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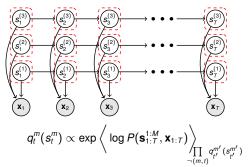
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- Thus, we can update each q<sub>i</sub> in turn given the means (or, in general, mean sufficient statistics) of the others.
- Each variable sees the mean field imposed by its neighbours, and we update these fields until they all agree.

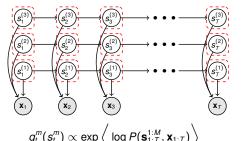




$$q(s_{1:T}^{1:M}) = \prod_{m,t} q_t^m(s_t^m)$$



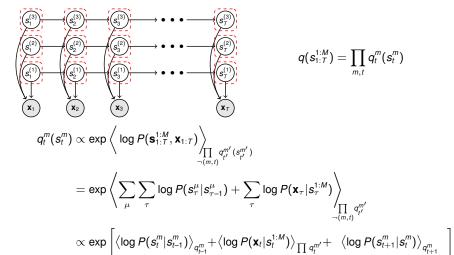
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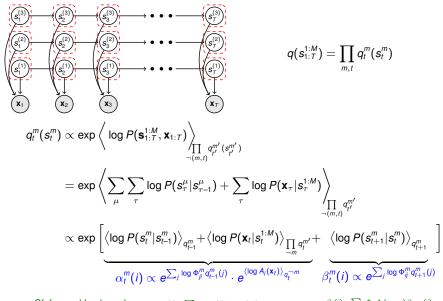


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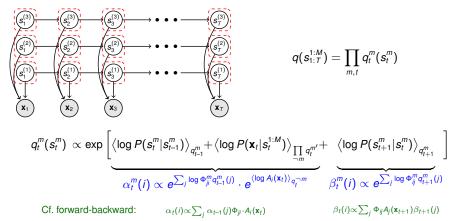




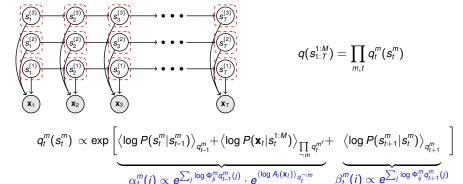
Cf. forward-backward:

$$\alpha_t(i) \propto \sum_j \alpha_{t-1}(j) \Phi_{ji} \cdot A_i(\mathbf{x}_t)$$

$$\beta_t(i) \propto \sum_j \Phi_{ij} A_j(\mathbf{x}_{t+1}) \beta_{t+1}(j)$$



Yields a message-passing algorithm like forward-backward

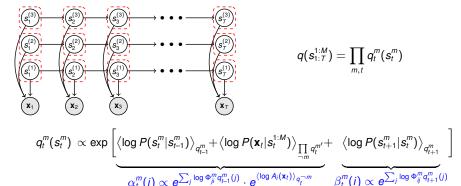


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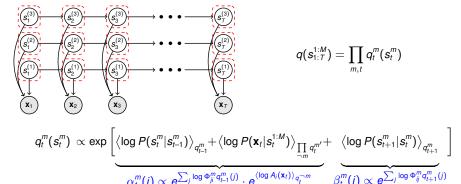


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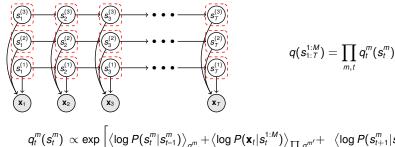
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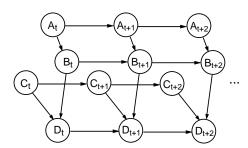


$$q_t^m(\boldsymbol{s}_t^m) \propto \exp\left[\frac{\left\langle \log P(\boldsymbol{s}_t^m|\boldsymbol{s}_{t-1}^m) \right\rangle_{q_{t-1}^m} + \left\langle \log P(\boldsymbol{\mathbf{x}}_t|\boldsymbol{s}_t^{1:M}) \right\rangle_{\prod\limits_{m} q_t^{m'}} + \frac{\left\langle \log P(\boldsymbol{s}_{t+1}^m|\boldsymbol{s}_t^m) \right\rangle_{q_{t-1}^m}}{\beta_t^m(i) \propto e^{\sum_{j} \log \Phi_{jj}^m q_{t-1}^m(j)} \cdot e^{\left\langle \log A_j(\boldsymbol{\mathbf{x}}_t) \right\rangle_{q_t^{m-m}}} \beta_t^m(i) \propto e^{\sum_{j} \log \Phi_{jj}^m q_{t+1}^m(j)}$$
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- Evidence does not appear explicitly in backward message (cf Kalman smoothing)

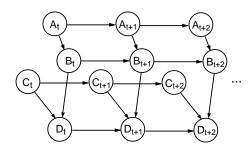
# Structured variational approximation

•  $q(\mathcal{Z})$  need not be completely factorized.



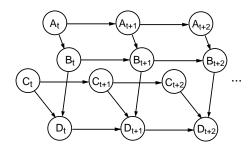
## Structured variational approximation

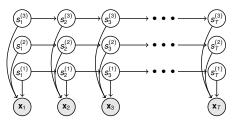
- q(Z) need not be completely factorized.
- ▶ For example, suppose  $\mathcal{Z}$  can be partitioned into sets  $\mathcal{Z}_1$  and  $\mathcal{Z}_2$  such that computing the expected sufficient statistics under  $P(\mathcal{Z}_1|\mathcal{Z}_2,\mathcal{X})$  and  $P(\mathcal{Z}_2|\mathcal{Z}_1,\mathcal{X})$  would be tractable.
- $\Rightarrow$  Then the factored approximation  $q(\mathcal{Z}) = q(\mathcal{Z}_1)q(\mathcal{Z}_2)$  is tractable.

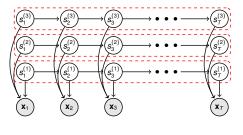


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- $\Rightarrow$  Then the factored approximation  $q(\mathcal{Z}) = q(\mathcal{Z}_1)q(\mathcal{Z}_2)$  is tractable.
- In particular, any factorisation of  $q(\mathcal{Z})$  into a product of distributions on trees, yields a tractable approximation.

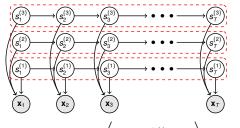






For the FHMM we can factor the chains:

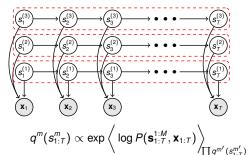
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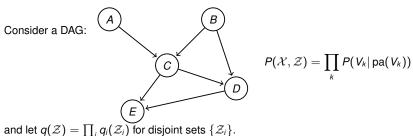
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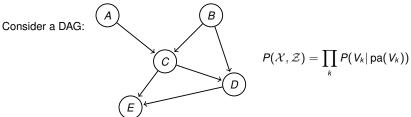
$$= \exp\left\langle \sum_{\mu} \sum_{t} \log P(s_{t}^{\mu} | s_{t-1}^{\mu}) + \sum_{t} \log P(\mathbf{x}_{t} | s_{t}^{1:M}) \right\rangle_{\prod_{m} q^{m'}}$$

$$\propto \exp\left[ \sum_{t} \log P(s_{t}^{m} | s_{t-1}^{m}) + \sum_{t} \left\langle \log P(\mathbf{x}_{t} | s_{t}^{1:M}) \right\rangle_{\prod_{m} q^{m'}(s_{t}^{m'})} \right]$$

$$= \prod_{t} P(s_{t}^{m} | s_{t-1}^{m}) \prod_{t} e^{\left\langle \log P(\mathbf{x}_{t} | s_{t}^{1:M}) \right\rangle_{\prod_{m} q^{m'}(s_{t}^{m'})}}$$

This looks like a standard HMM joint, with a modified likelihood term  $\Rightarrow$  cycle through multiple forward-backward passes, updating likelihood terms each time.

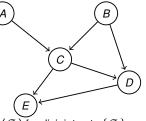




and let  $q(\mathcal{Z}) = \prod_i q_i(\mathcal{Z}_i)$  for disjoint sets  $\{\mathcal{Z}_i\}$ .

We have that the VE update for  $q_i$  is given by  $q_i^*(\mathcal{Z}_i) \propto \exp{\langle \log p(\mathcal{Z}, \mathcal{X}) \rangle_{q_{\neg i}(\mathcal{Z})}}$  where  $\langle \cdot \rangle_{q_{\neg i}(\mathcal{Z})}$  denotes averaging with respect to  $q_i(\mathcal{Z}_j)$  for all  $j \neq i$ 

Consider a DAG:



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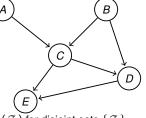
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This defines messages that are passed between nodes in the graph. Each node receives messages from its Markov boundary: parents, children and parents of children (all neighbours in the corresponding factor graph).

#### Non-factored variational methods

The term variational approximation is used whenever a bound on the likelihood (or on another estimation cost function) is optimised, but does not necessarily become tight.

Many further variational approximations have been developed, including:

- parametric forms (e.g. Gaussian) for non-linear models
  - closed form updates in special cases
  - numerical or sampling-based computation of expectations
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Variational methods can also be used to find an approximate posterior on the parameters.

So far, we have applied Jensen's bound and factorisations to help with integrals over latent variables.

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Some call this the "Evidence Lower Bound" (ELBO). I'm not fond of that term.

# **Variational Bayesian EM...**

Coordinate maximization of the VB free-energy lower bound

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leads to EM-like updates:

$$\begin{array}{ll} Q_{\mathcal{Z}}^*(\mathcal{Z}) \propto \exp \left\langle \log P(\mathcal{Z}, \mathcal{X} | \boldsymbol{\theta}) \right\rangle_{Q_{\boldsymbol{\theta}}(\boldsymbol{\theta})} & \textit{E-like step} \\ Q_{\boldsymbol{\theta}}^*(\boldsymbol{\theta}) \propto P(\boldsymbol{\theta}) \exp \left\langle \log P(\mathcal{Z}, \mathcal{X} | \boldsymbol{\theta}) \right\rangle_{Q_{\mathcal{Z}}(\mathcal{Z})} & \textit{M-like step} \end{array}$$

# Variational Bayesian EM ...

Coordinate maximization of the VB free-energy lower bound

$$\mathcal{F}(Q_{\mathcal{Z}},Q_{ heta}) = \iint\!\!d\mathcal{Z}\,d heta\,\,\,Q_{\mathcal{Z}}(\mathcal{Z})Q_{ heta}( heta)\lograc{p(\mathcal{X},\mathcal{Z}, heta|\mathcal{M})}{Q_{\mathcal{Z}}(\mathcal{Z})Q_{ heta}( heta)}$$

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Maximizing  $\mathcal{F}$  is equivalent to minimizing KL-divergence between the *approximate posterior*,  $Q(\theta)Q(\mathcal{Z})$  and the *true posterior*,  $P(\theta,\mathcal{Z}|\mathcal{X})$ .

$$\begin{split} \log P(\mathcal{X}) - \mathcal{F}(Q_{\mathcal{Z}}, Q_{\theta}) &= \log P(\mathcal{X}) - \iint \! d\mathcal{Z} \, d\theta \ \, Q_{\mathcal{Z}}(\mathcal{Z}) Q_{\theta}(\theta) \log \frac{P(\mathcal{X}, \mathcal{Z}, \theta)}{Q_{\mathcal{Z}}(\mathcal{Z}) Q_{\theta}(\theta)} \\ &= \iint \! \! d\mathcal{Z} \, d\theta \ \, Q_{\mathcal{Z}}(\mathcal{Z}) Q_{\theta}(\theta) \log \frac{Q_{\mathcal{Z}}(\mathcal{Z}) Q_{\theta}(\theta)}{P(\mathcal{Z}, \theta | \mathcal{X})} = \mathit{KL}(Q || P) \end{split}$$

# **Conjugate-Exponential models**

Let's focus on *conjugate-exponential* (**CE**) latent-variable models:

► Condition (1). The joint probability over *variables* is in the exponential family:

$$P(\mathcal{Z}, \mathcal{X} | \boldsymbol{\theta}) = f(\mathcal{Z}, \mathcal{X}) \ g(\boldsymbol{\theta}) \exp \left\{ \phi(\boldsymbol{\theta})^\mathsf{T} \mathsf{T}(\mathcal{Z}, \mathcal{X}) \right\}$$

where  $\phi(\theta)$  is the vector of *natural parameters*, T are *sufficient statistics* 

Condition (2). The prior over parameters is conjugate to this joint probability:

$$P(\boldsymbol{\theta}|\nu, au) = h(\nu, au) \; g(\boldsymbol{\theta})^{
u} \exp\left\{\phi(\boldsymbol{\theta})^{\mathsf{T}} au
ight\}$$

where  $\nu$  and  $\tau$  are hyperparameters of the prior.

Conjugate priors are computationally convenient and have an intuitive interpretation:

- ν: number of pseudo-observations
- $\triangleright$   $\tau$ : values of pseudo-observations

# **Conjugate-Exponential examples**

#### In the **CE** family:

- Gaussian mixtures
- factor analysis, probabilistic PCA
- hidden Markov models and factorial HMMs
- linear dynamical systems and switching models
- discrete-variable belief networks

Other as yet undreamt-of models combinations of Gaussian, Gamma, Poisson, Dirichlet, Wishart, Multinomial and others.

#### Not in the **CE** family:

- Boltzmann machines, MRFs (no simple conjugacy)
- logistic regression (no simple conjugacy)
- sigmoid belief networks (not exponential)
- independent components analysis (not exponential)

Note: one can often approximate such models with a suitable choice from the **CE** family.

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 with  $\tilde{\nu} = \nu + n$  and  $\tilde{\tau} = \tau + \sum_{i} \left\langle \mathsf{T}(\mathbf{z}_{i}, \mathbf{x}_{i}) \right\rangle_{Q_{\mathcal{Z}}}$ 

ullet  $Q_{\mathcal{Z}}(\mathcal{Z}) = \prod_{i=1}^n Q_{\mathbf{z}_i}(\mathbf{z}_i)$  takes the same form as in the E-step of regular EM

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with natural parameters  $\overline{\phi}( heta) = \langle \phi( heta) 
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with natural parameters  $\overline{\phi}(\theta)=\langle\phi(\theta)\rangle_{\mathcal{Q}_{\theta}}\Rightarrow$  inference unchanged from regular EM.

#### **EM for MAP estimation**

Goal: maximize  $P(\theta|\mathcal{X}, m)$  wrt  $\theta$ 

E Step: compute

$$Q_{\mathcal{Z}}(\mathcal{Z}) \leftarrow p(\mathcal{Z}|\mathcal{X}, \boldsymbol{\theta})$$

M Step:

$$\theta \leftarrow \underset{\theta}{\operatorname{argmax}} \int d\mathcal{Z} \ Q_{\mathcal{Z}}(\mathcal{Z}) \log P(\mathcal{Z}, \mathcal{X}, \theta)$$

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## Properties:

▶ Reduces to the EM algorithm if  $Q_{\theta}(\theta) = \delta(\theta - \theta^*)$ .

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- VB-E step has same complexity as corresponding E step.
- We can use the junction tree, belief propagation, Kalman filter, etc, algorithms in the VB-E step of VB-EM, but using expected natural parameters,  $\bar{\phi}$ .

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- These lower bounds can be compared amongst models to learn the right (structure, connectivity . . . of the) model
- If a continuous domain of models is specified by a hyperparameter η, then the VB free energy depends on that parameter:

$$\mathcal{F}(Q_{\mathcal{Z}}, Q_{\theta}, \eta) = \iint \!\! d\mathcal{Z} \, d\theta \ Q_{\mathcal{Z}}(\mathcal{Z}) Q_{\theta}(\theta) \log \frac{P(\mathcal{X}, \mathcal{Z}, \theta | \eta)}{Q_{\mathcal{Z}}(\mathcal{Z}) Q_{\theta}(\theta)} \leq P(\mathcal{X} | \eta)$$

A hyper-M step maximises the current bound wrt  $\eta$ :

$$\eta \leftarrow \operatorname*{argmax} \iint \!\! d\mathcal{Z} \, d\theta \, \, \, Q_{\mathcal{Z}}(\mathcal{Z}) Q_{\theta}(\theta) \log P(\mathcal{X}, \mathcal{Z}, \theta | \eta)$$

Recall that ARD (automatic relevance determination) was a hyperparameter method to select relevant or useful inputs in regression.

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- A similar idea used with variational Bayesian methods can learn a latent dimensionality.
- Consider factor analysis:

$$\mathbf{x} \sim \mathcal{N}\left(\Lambda \mathbf{z}, \Psi\right)$$
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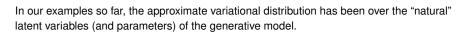
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- ► In this case, these parameters select "relevant" latent dimensions, effectively learning the dimensionality of **z**.

# Augmented Variational Methods



Sometimes it may be useful to introduce additional latent variables, solely to achieve computational tractability.

Two examples are GP regression and the GPLVM.

# **Sparse GP approximations**

GP predictions:

$$y'|X,Y,\boldsymbol{x}' \sim \mathcal{N}\left(K_{\boldsymbol{x}'X}(K_{XX} + \sigma^2I)^{-1}Y,K_{\boldsymbol{x}'\boldsymbol{x}'} - K_{\boldsymbol{x}'X}(K_{XX+\sigma^2I)}^{-1}K_{X\boldsymbol{x}'} + \sigma^2\right)$$

Evidence (for learning kernel hyperparameters):

$$\log P(Y|X) = -\frac{1}{2}\log|2\pi(K_{XX} + \sigma^2 I)| - \frac{1}{2}Y(K_{XX} + \sigma^2 I)^{-1}Y^{\mathsf{T}}$$

Computing either form requires inverting the  $N \times N$  matrix  $K_{XX}$ , in  $\mathcal{O}(N^3)$  time.

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What values should *U* and *Z* take?

Write F for the (smooth) GP function values that underlie Y (so  $Y \sim \mathcal{N}\left(F, \sigma^2 I\right)$ ).

Introduce latent measurements U at inputs Z (and integrate over U).

The likelihood can be written

$$P(Y|X) = \iint dF \, dU \, P(Y,F,U|X,Z) = \iint dF \, dU \, P(Y|F) P(F|U,X,Z) P(U|Z)$$

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$$P(Y|X) = \iint dF \, dU \, P(Y,F,U|X,Z) = \iint dF \, dU \, P(Y|F) P(F|U,X,Z) P(U|Z)$$

Now, both U and F are latent, so we introduce a variational distribution q(F, U) to form a free-energy.

$$\mathcal{F}(q(F,U),\theta) = \left\langle \log \frac{P(Y|F)P(F|U,X,Z)P(U|Z)}{q(F,U)} \right\rangle_{q(F,U)}$$

Write F for the (smooth) GP function values that underlie Y (so  $Y \sim \mathcal{N}(F, \sigma^2 I)$ ).

Introduce latent measurements U at inputs Z (and integrate over U).

The likelihood can be written

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Now, choose the variational form q(F, U) = P(F|U, X, Z)q(U). That is, fix F|U without reference to Y – so information about Y will need to be "compressed" into q(U).

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$$\mathcal{F}(q(F,U),\theta, \mathbf{Z}) = \left\langle \log \frac{P(Y|F) \ P(F|U,X,Z) \ P(U|Z)}{P(F|U,X,Z) \ q(U)} \right\rangle_{P(F|U)q(U)}$$

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$$\mathcal{F}(q(F,U),\theta, \textcolor{red}{Z}) = \left\langle \log \frac{P(Y|F) \, P(F|U, \textcolor{blue}{X, \textcolor{blue}{Z}}) \, P(U|\textcolor{blue}{Z})}{P(F|U, \textcolor{blue}{X, \textcolor{blue}{Z}}) \, q(U)} \right\rangle_{P(F|U)q(U)}$$

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$$\mathcal{F}(q(F,U),\theta,\mathbf{Z}) = \left\langle \log \frac{P(Y|F)}{P(F|U,X,Z)} \frac{P(U|Z)}{P(F|U,X,Z)} \right\rangle_{P(F|U)q(U)}$$
$$= \left\langle \left\langle \log P(Y|F) \right\rangle_{P(F|U)} + \log P(U|Z) - \log q(U) \right\rangle_{q(U)}$$

$$\mathcal{F}(q(U), \theta, Z) = \left\langle \left\langle \log P(Y|F) \right\rangle_{P(F|U)} + \log P(U|Z) - \log q(U) \right\rangle_{q(U)}$$

Now P(F|U) is fixed by the generative model (rather than being subject to free optimisation). So we can evaluate that expectation:

$$\begin{split} &\langle \log P(Y|F) \rangle_{P(F|U)} \\ &= \left\langle -\frac{1}{2} \log \left| 2\pi \sigma^2 I \right| - \frac{1}{2\sigma^2} \mathrm{Tr} \left[ (Y-F)(Y-F)^{\mathsf{T}} \right] \right\rangle_{P(F|U)} \\ &= -\frac{1}{2} \log \left| 2\pi \sigma^2 I \right| - \frac{1}{2\sigma^2} \mathrm{Tr} \left[ (Y-\langle F \rangle_{P(F|U)}) (Y-\langle F \rangle_{P(F|U)})^{\mathsf{T}} \right] - \frac{1}{2\sigma^2} \mathrm{Tr} \left[ \Sigma_{F|U} \right] \\ &= \log \mathcal{N} \left( Y | K_{XZ} K_{ZZ}^{-1} U, \sigma^2 I \right) - \frac{1}{2\sigma^2} \mathrm{Tr} \left[ K_{XX} - K_{XZ} K_{ZZ}^{-1} K_{ZX} \right] \end{split}$$

$$\mathcal{F}(q(U), \theta, Z) = \left\langle \left\langle \log P(Y|F) \right\rangle_{P(F|U)} + \log P(U|Z) - \log q(U) \right\rangle_{q(U)}$$

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So,

$$\begin{split} \mathcal{F}(q(U),\theta,Z) &= \left\langle \log \mathcal{N}\left(Y|K_{XZ}K_{ZZ}^{-1}U,\sigma^2I\right) + \log P(U|Z) - \log q(U)\right\rangle_{q(U)} \\ &- \frac{1}{2\sigma^2} \text{Tr}\left[K_{XX} - K_{XZ}K_{ZZ}^{-1}K_{ZX}\right] \;. \end{split}$$

$$\mathcal{N}(Y|K_{XZ}K_{ZZ}^{-1}U,\sigma^2I)P(U|Z)$$

$$\mathcal{F}(q(U), \theta, Z) = \left\langle \log \frac{\mathcal{N}\left(Y | K_{XZ} K_{ZZ}^{-1} U, \sigma^2 I\right) P(U | Z)}{q(U)} \right\rangle_{q(U)} - \frac{1}{2\sigma^2} \operatorname{Tr}\left[K_{XX} - K_{XZ} K_{ZZ}^{-1} K_{ZX}\right].$$

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The expectation is the free energy of a PPCA-like model with normal prior  $U \sim \mathcal{N}\left(0, K_{ZZ}\right)$  and loading matrix  $K_{XZ}K_{ZZ}^{-1}$ . The maximum of this free energy is the log-likelihood (achieved with q equal to the posterior under the PPCA-like model).

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This gives

$$\mathcal{F}(q^*(U), \theta, Z) = \log \mathcal{N}\left(Y|0, K_{XZ} K_{ZZ}^{-1} K_{ZZ} K_{ZZ}^{-1} K_{ZX} + \sigma^2 I\right) - \frac{1}{2\sigma^2} \text{Tr}\left[K_{XX} - K_{XZ} K_{ZZ}^{-1} K_{ZX}\right].$$

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Note that we have eliminated all terms in  $K_{xx}^{-1}$ .

We can optimise the free energy numerically with respect to Z and  $\theta$  to adjust the GP prior and quality of variational approximation.

$$\mathcal{F}(q(U), \theta, Z) = \left\langle \log \frac{\mathcal{N}\left(Y | K_{XZ} K_{ZZ}^{-1} U, \sigma^2 I\right) P(U | Z)}{q(U)} \right\rangle_{q(U)} - \frac{1}{2\sigma^2} \text{Tr}\left[K_{XX} - K_{XZ} K_{ZZ}^{-1} K_{ZX}\right].$$

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$$\mathcal{F}(q^*(U), \theta, Z) = \log \mathcal{N}\left(Y|0, K_{XZ} K_{ZZ}^{-1} K_{ZZ} K_{ZZ}^{-1} K_{ZX} + \sigma^2 I\right) - \frac{1}{2\sigma^2} \text{Tr}\left[K_{XX} - K_{XZ} K_{ZZ}^{-1} K_{ZX}\right].$$

Note that we have eliminated all terms in  $K_{XX}^{-1}$ .

We can optimise the free energy numerically with respect to Z and  $\theta$  to adjust the GP prior and quality of variational approximation.

A similar approach can be used to learn X if they are unobserved (i.e. in the GPLVM). Assume q(X, F, U) = q(X)P(F|X, U)q(U). Then  $\mathcal{F} = \langle \log P(Y, F, U|X) \log P(X) \rangle_{q(U)q(X)}$  which simplifies into tractable components in much the same way as above.

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