

Advanced Markov Chain Monte Carlo

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April 3, 2009

Set-up

- ▶ Models: m belong to a class of models \mathcal{M} .
- ▶ Data: $X = \{x_1, x_2, \dots, x_n\}$.
- ▶ Latent variables: $Y = \{y_1, y_2, \dots, y_n\}$.
- ▶ Parameters: θ_m for model $m \in \mathcal{M}$.
- ▶ Likelihood:

$$p(X|Y, \theta_m) = \prod_{i=1}^n p(x_i|y_i, \theta_m)$$

$$p(Y|\theta_m) = \prod_{i=1}^n p(y_i|\theta_m)$$

- ▶ Prior:

$$p(m)$$
$$p(\theta_m|m)$$

Bayesian Inference

- ▶ Posterior distribution:

$$\pi(\theta_m, Y|m) = p(\theta_m, Y|m, X) = \frac{p(\theta_m|m)p(Y|\theta_m)p(X|Y, \theta_m)}{p(X|m)}$$

$$\pi(m) = p(m|X)$$

- ▶ Evidence:

$$p(X|m)$$

- ▶ Bayes Factors:

$$\frac{p(X|m)}{p(X|m')}$$

Metropolis-Hastings

- ▶ Metropolis-Hastings is a standard and flexible Markov chain Monte Carlo transition kernel.
- ▶ Distribution we wish to sample from:

$$\pi(\theta)$$

- ▶ Proposal kernel:

$$q(\theta \rightarrow \tilde{\theta})$$

- ▶ Acceptance probability:

$$\alpha(\theta \rightarrow \tilde{\theta}) = \min \left(1, \frac{\pi(\tilde{\theta})q(\tilde{\theta} \rightarrow \theta)}{\pi(\theta)q(\theta \rightarrow \tilde{\theta})} \right)$$

Reversible-Jump Markov Chain Monte Carlo

- ▶ To do Bayesian model averaging, we are interested in computing

$$\pi(m, \theta_m)$$

where for different m 's θ_m 's come from different spaces with different dimensionalities.

- ▶ Most MCMC techniques can only deal with sampling from distributions over a single space, e.g. the posterior distribution over parameters of a *single* model.
- ▶ Reversible-Jump MCMC allows us to sample from the full posterior $\pi(m, \theta_m)$ over both models and parameters properly.

Simple Example

- ▶ Imagine mixtures of either 1 or 2 Gaussians in 1D:

$$m = 1 : \quad \mathcal{N}(\theta_0, \sigma^2)$$

$$m = 2 : \quad .5\mathcal{N}(\theta_1, \sigma^2) + .5\mathcal{N}(\theta_2, \sigma^2)$$

where σ^2 is known and parameters consist of the mean θ_0 for model 1 and the means (θ_1, θ_2) for model 2.

- ▶ Conditioned on m we can easily sample θ_0 or (θ_1, θ_2) depending on value of m .
- ▶ How do we sample m ? The θ spaces are different for different m 's, so we have to change the θ 's if we propose changing m .

Simple Example

- ▶ Say $m = 1$ and we propose changing $m = 1 \rightarrow m = 2$. We would also need to propose (θ_1, θ_2) as well. Since this proposal is a “split” of one Gaussian into two, a reasonable proposal would be:

$$\delta \sim \mathcal{N}(0, v)$$

$$\theta_1 \leftarrow \theta_0 + \delta$$

$$\theta_2 \leftarrow \theta_0 - \delta$$

This preserves the overall mean of the distribution over the data.

- ▶ To compute the MH acceptance probability we need the reverse proposal probabilities for $m = 2 \rightarrow m = 1$. Given a particular (θ_1, θ_2) , notice that *the only* θ_0 that could have been used to propose (θ_1, θ_2) is their average:

$$\theta_0 = (\theta_1 + \theta_2)/2$$

Simple Example

- ▶ Recap of the proposal and the reverse proposal:

$$m = 1 \rightarrow m = 2$$

$$m = 2 \rightarrow m = 1$$

$$\theta_0 \rightarrow (\theta_1, \theta_2) = (\theta_0 + \delta, \theta_0 - \delta) \quad (\theta_1, \theta_2) \rightarrow \theta_0 = (\theta_1 + \theta_2)/2$$

$$\delta \sim \mathcal{N}(\mathbf{0}, \mathbf{v})$$

- ▶ The acceptance probability is:

$$\min \left(1, \frac{\pi(m = 2, \theta_0 + \delta, \theta_0 - \delta) \mathbf{1}}{\pi(m = 1, \theta_0) \mathcal{N}(\delta; \mathbf{0}, \mathbf{v})} \right)$$

Notice that the dimensions match up.

Another Example

- ▶ If the $\theta_0 \rightarrow (\theta_1, \theta_2)$ proposal is simply:

$$\theta_1 = \theta_0$$

$$\theta_2 \sim \mathcal{N}(\mathbf{0}, \nu)$$

i.e. the first Gaussian stays the same, create a second Gaussian.

- ▶ Then the reverse proposal $(\theta_1, \theta_2) \rightarrow \theta_0$ would have needed to be:

$$\theta_0 = \theta_1$$

i.e. drop the second Gaussian and keep the first one.

General Case

- ▶ We want a proposal $q(m \rightarrow w)$ where $m, w \in \mathcal{M}$.
- ▶ We need to design proposals for the parameters $q(\theta_m \rightarrow \theta_w)$ and reverse proposals $q(\theta_w \rightarrow \theta_m)$.
- ▶ The acceptance probability of RJMCMC,

$$\min \left(1, \frac{\pi(w)\pi(\theta_w|w)q(w \rightarrow m)q(\theta_w \rightarrow \theta_m)}{\pi(m)\pi(\theta_m|m)q(m \rightarrow w)q(\theta_m \rightarrow \theta_w)} \right)$$

only makes sense if the θ densities are both wrt the same underlying measure. Simple case: measure lies on a 2D plane in $(\theta_0, \theta_1, \theta_2)$ space satisfying $\theta_0 = (\theta_1 + \theta_2)/2$.

