RKHS in ML: Comparing Two Samples

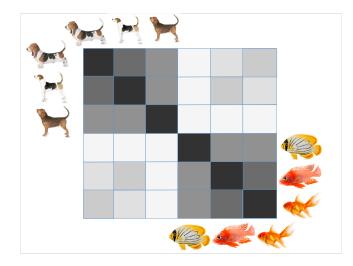
Arthur Gretton

Gatsby Computational Neuroscience Unit, University College London

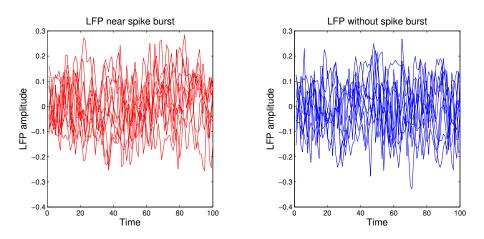
November 2, 2024

Comparing two samples

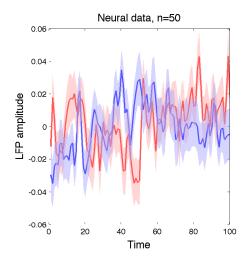
Given: Samples from unknown distributions P and Q.
Goal: do P and Q differ?



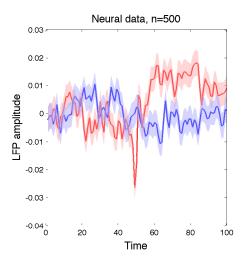
• The problem:Do local field potential (LFP) signals change when measured near a spike burst?



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CIFAR 10 samples

Cifar 10.1 samples

Significant difference?

Feng, Xu, Lu, Zhang, G., Sutherland, Learning Deep Kernels for Non-Parametric Two-Sample Tests, ICML 2020

Sutherland, Tung, Strathmann, De, Ramdas, Smola, G., ICLR 2017.

A real-life example: discrete domains

How do you compare distributions in a discrete domain?

 X_1 : Now disturbing reports out of Newfoundland show that the fragile snow crab industry is in serious decline. First the west coast salmon, the east coast salmon and the cod, and now the snow crabs off Newfoundland.

 X_2 : To my pleasant surprise he responded that he had personally visited those wharves and that he had already announced money to fix them. What wharves did the minister visit in my riding and how much additional funding is he going to provide for Delaps Cove, Hampton, Port Lorne, Y_1 : Honourable senators, I have a question for the Leader of the Government in the Senate with regard to the support funding to farmers that has been announced. Most farmers have not received any money yet.

 Y_2 :On the grain transportation system we have had the Estey report and the Kroeger report. We could go on and on. Recently programs have been announced over and over by the government such as money for the disaster in agriculture on the prairies and across Canada.

. . .

Are the gray extracts from the same distribution as the pink ones?

 $P_X = Q_Y$

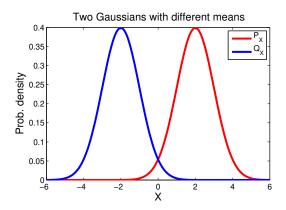
Outline

Two sample testing

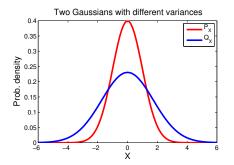
- Test statistic: Maximum Mean Discrepancy (MMD)...
 - ...as a difference in feature means
 - ...as an integral probability metric (not just a technicality!)
- Statistical testing with the MMD
- "How to choose the best kernel"
 - when are feature means unique?
 - what kernel gives the most powerful test?

Maximum Mean Discrepancy

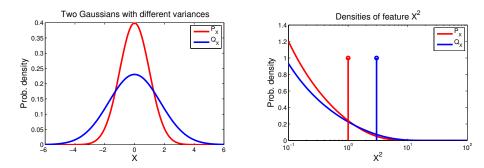
Simple example: 2 Gaussians with different meansAnswer: t-test



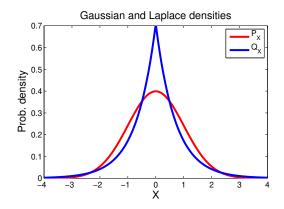
- Two Gaussians with same means, different variance
- Idea: look at difference in means of features of the RVs
- In Gaussian case: second order features of form $arphi(x)=x^2$



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- Idea: look at difference in means of features of the RVs
- In Gaussian case: second order features of form $arphi(x)=x^2$



- Gaussian and Laplace distributions
- Same mean *and* same variance
- Difference in means using higher order features...RKHS



Infinitely many features using kernels

Kernels: dot products of features

Feature map $\varphi(x) \in \mathcal{F}$,

$$arphi(x) = [\dots arphi_i(x) \dots] \in \ell_2$$

For positive definite k,

$$k(x,x')=\langle arphi(x),arphi(x')
angle_{\mathcal{F}}$$

Infinitely many features $\varphi(x)$, dot product in closed form!

Infinitely many features using kernels

Kernels: dot products of Exponentiated quadratic kernel features

$$k(x,x') = \exp\left(-\gamma \left\|x-x'
ight\|^2
ight)$$

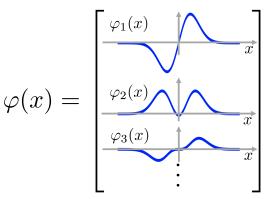
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Infinitely many features $\varphi(x)$, dot product in closed form!



Features: Gaussian Processes for Machine learning, Rasmussen and Williams, Ch. 4. 11/89

Infinitely many features of distributions

Given P a Borel probability measure on \mathcal{X} , define feature map of probability P,

$$\mu_P = [\dots \mathbb{E}_P [\varphi_i(X)] \dots]$$

For positive definite k(x, x'),

$$\langle \mu_P, \mu_Q
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for $x \sim P$ and $y \sim Q$.

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Expectations of RKHS functions

Function evaluation in an RKHS:

$$f(\boldsymbol{x}) = \langle f, \boldsymbol{\varphi}_{\boldsymbol{x}} \rangle_{\mathcal{F}}$$

Expectation evaulation in an RKHS:

$$\mathbb{E}_{\mathbb{P}}(f(X)) = \langle f, \boldsymbol{\mu}_{\mathbb{P}} \rangle_{\mathcal{F}}$$

 μ_P gives you expectations of all RKHS functions

Empirical mean embedding:

$$\widehat{\mu}_P = m^{-1} \sum_{i=1}^m arphi_{x_i} \qquad x_i \stackrel{ ext{i.i.d.}}{\sim} P$$

... does this reasoning work in infinite dimensions?

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... does this reasoning work in infinite dimensions?

Does there exist an element $\mu_P \in \mathcal{F}$ such that

$$\mathrm{E}_{P}f(x) = \langle f, \mu_{P}
angle_{\mathcal{F}} \qquad orall f \in \mathcal{F}$$

We recall the concept of a bounded operator: a linear operator $A : \mathcal{F} \to \mathbb{R}$ is bounded when

$$|Af| \leq \lambda_A ||f||_{\mathcal{F}} \quad \forall f \in \mathcal{F}.$$

Riesz representation theorem: In a Hilbert space \mathcal{F} , all bounded linear operators A can be written $\langle \cdot, g_A \rangle_{\mathcal{F}}$, for some $g_A \in \mathcal{F}$,

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Existence of mean embedding: If $\mathbb{E}_P \sqrt{k(x,x)} = \mathbb{E}_P \|\varphi(x)\|_{\mathcal{F}} < \infty$ then $\exists \mu_P \in \mathcal{F}$.

Proof:

The linear operator $T_P f := \mathrm{E}_P f(x)$ for all $f \in \mathcal{F}$ is bounded under the assumption, since

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μ_P as a function in the RKHS

Embedding of P to feature space

• Mean embedding $\mu_P \in \mathcal{F}$,

 $\langle \mu_P, f \rangle_{\mathcal{F}} = E_P f(x).$

• What does prob. feature map look like?

$$egin{aligned} \mu_P(t) &= \langle \mu_P, arphi(t)
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Expectation of kernel!

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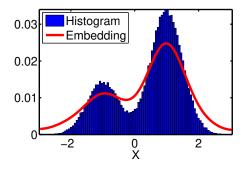
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The maximum mean discrepancy is the distance between feature means:

$$MMD^{2}(P, Q) = \|\mu_{P} - \mu_{Q}\|_{\mathcal{F}}^{2}$$
$$= \underbrace{\mathbb{E}_{P}k(x, x')}_{(a)} + \underbrace{\mathbb{E}_{Q}k(y, y')}_{(a)} - 2\underbrace{\mathbb{E}_{P,Q}k(x, y)}_{(b)}$$

(a) = within distrib. similarity, (b) = cross-distrib. similarity.

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Illustration of MMD

- **Dogs** (= P) and fish (= Q) example revisited
- Each entry is one of $k(\text{dog}_i, \text{dog}_j)$, $k(\text{dog}_i, \text{fish}_j)$, or $k(\text{fish}_i, \text{fish}_j)$

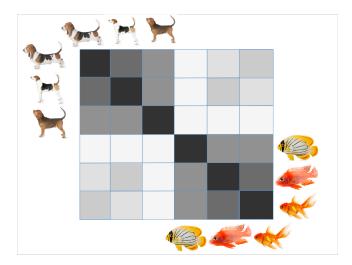
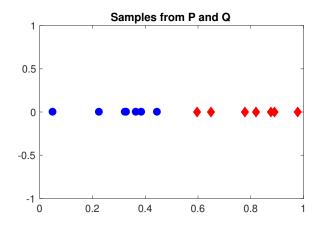


Illustration of MMD

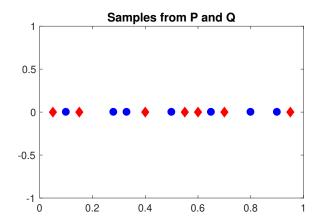
The maximum mean discrepancy:

$$\widehat{MMD}^{2} = \frac{1}{n(n-1)} \sum_{i \neq j} k(\operatorname{dog}_{i}, \operatorname{dog}_{j}) + \frac{1}{n(n-1)} \sum_{i \neq j} k(\operatorname{fish}_{i}, \operatorname{fish}_{j})$$
$$- \frac{2}{n^{2}} \sum k(\operatorname{dog}_{i}, \operatorname{fish}_{j})$$
$$k(\operatorname{dog}_{i}, \operatorname{dog}_{j}) k(\operatorname{dog}_{i}, \operatorname{fish}_{j})$$
$$k(\operatorname{fish}_{i}, \operatorname{dog}_{i}) k(\operatorname{fish}_{i}, \operatorname{fish}_{j})$$

Are P and Q different?

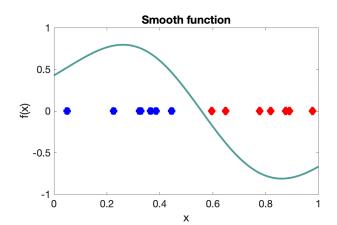


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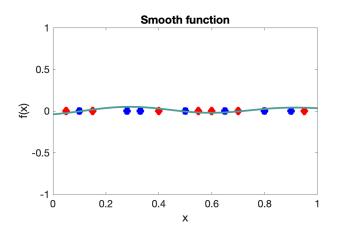
Integral probability metric:

Find a "well behaved function" f(x) to maximize

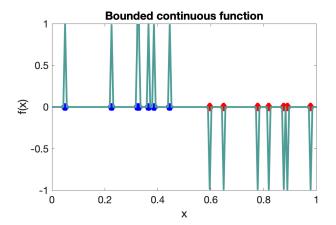


Integral probability metric:

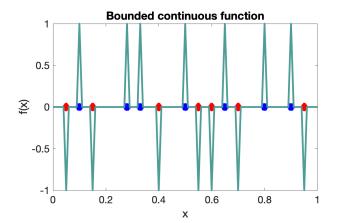
Find a "well behaved function" f(x) to maximize



What if the function is not smooth?

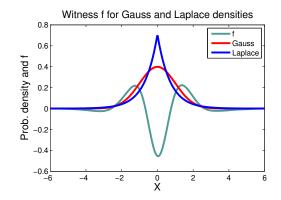


What if the function is not smooth?



Maximum mean discrepancy: smooth function for P vs Q

$$egin{aligned} MMD(P,oldsymbol{Q};F) := \sup_{\|f\|\leq 1} \left[\mathrm{E}_P f(X) - \mathrm{E}_{oldsymbol{Q}} f(Y)
ight] \ (F = \mathrm{unit\ ball\ in\ RKHS\ \mathcal{F}}) \end{aligned}$$



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For characteristic RKHS \mathcal{F} , MMD(P, Q; F) = 0 iff P = Q

Other choices for witness function class:

- Bounded continuous [Dudley, 2002]
- Bounded varation 1 (Kolmogorov metric) [Müller, 1997]
- Bounded Lipschitz (Wasserstein distances) [Dudley, 2002]

Maximum mean discrepancy: smooth function for P vs Q

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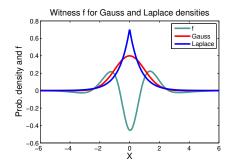
A reminder for the proof on the next slide:

$$\mathrm{E}_{P}(f(X)) = \langle f, \mathrm{E}_{P} arphi(X)
angle_{\mathcal{F}} = \langle f, \mu_{P}
angle_{\mathcal{F}}$$

(always true if kernel is bounded)

The MMD:

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The MMD:

MMD(P, Q; F)

- $= \sup_{\|f\|_{\mathcal{F}} \leq 1} \left[\operatorname{E}_{\mathcal{P}} f(X) \operatorname{E}_{\mathcal{Q}} f(Y)
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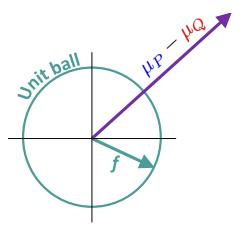
use

 $\mathrm{E}_{P}f(X) = \langle \mu_{P}, f \rangle_{\mathcal{F}}$

The MMD:

MMD(P, Q; F)

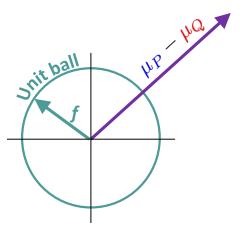
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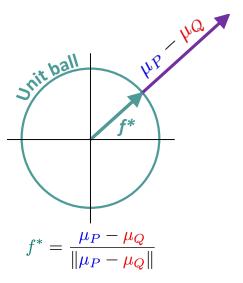
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The MMD:

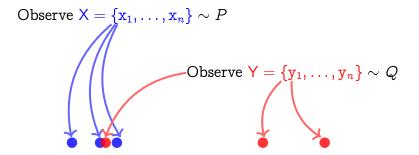
- MMD(P, Q; F)
- $= \sup_{\|f\|_{\mathcal{F}} \leq 1} \left[\operatorname{E}_{P} f(X) \operatorname{E}_{\mathcal{Q}} f(Y)
 ight]$
- $= \sup_{\|f\|_{\mathcal{F}} \leq 1} \langle f, \mu_P \mu_Q
 angle_{\mathcal{F}}$

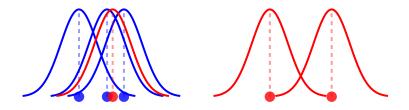


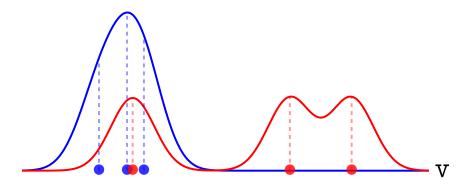
The MMD:

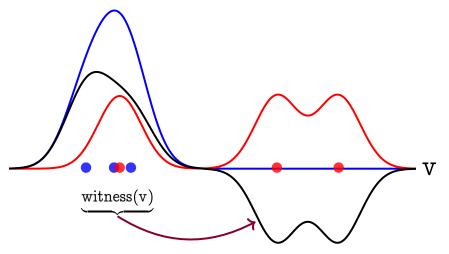
- MMD(P, Q; F)
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 ight]$
- $= \sup_{\|f\|_{\mathcal{F}} \leq 1} \langle f, \mu_P \mu_Q
 angle_{\mathcal{F}}$
- $= \|\mu_P \mu_Q\|$

Function view and feature view equivalent









Recall the witness function expression

 $f^* \propto \mu_P - \mu_Q$

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The empirical feature mean for P

$$\widehat{\mu}_P := rac{1}{n}\sum_{i=1}^n arphi(x_i)$$

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The empirical witness function at v

$$f^*(v) = \langle f^*, arphi(v)
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angle_\mathcal{F} \ &\propto \langle \widehat{\mu}_P - \widehat{\mu}_Q, arphi(v)
angle_\mathcal{F} \end{aligned}$$

Recall the witness function expression

 $f^* \propto \mu_P - \mu_Q$

The empirical feature mean for P

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The empirical witness function at v

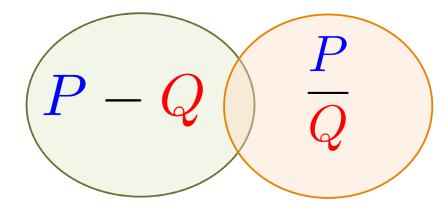
$$egin{aligned} f^*(v) &= \langle f^*, arphi(v)
angle_{\mathcal{F}} \ &\propto \langle \widehat{\mu}_P - \widehat{\mu}_{oldsymbol{Q}}, arphi(v)
angle_{\mathcal{F}} \ &= rac{1}{n} \sum_{i=1}^n k(x_i, v) - rac{1}{n} \sum_{i=1}^n k(\mathbf{y}_i, v) \end{aligned}$$

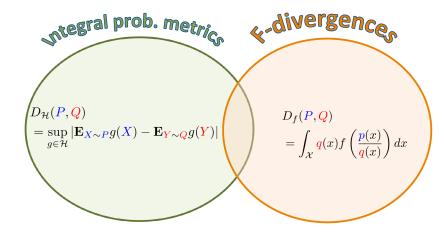
Don't need explicit feature coefficients $f^* := \begin{bmatrix} f_1^* & f_2^* & \dots \end{bmatrix}$

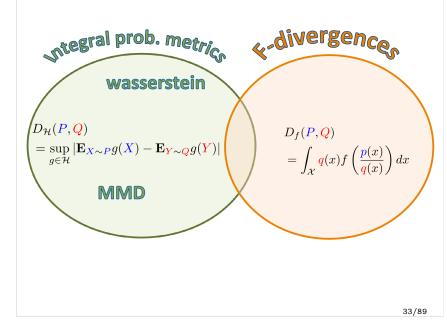
29/89

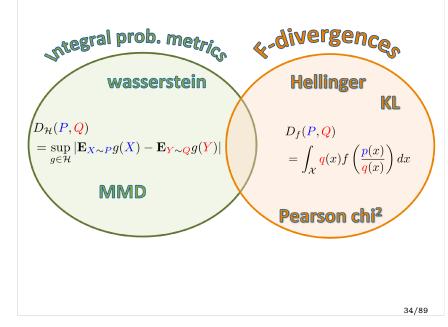
Interlude: divergence measures

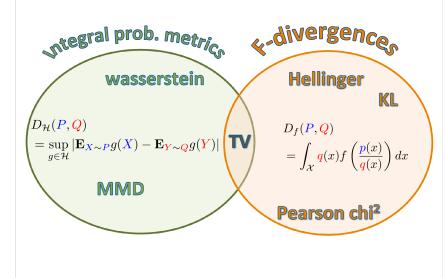












Sriperumbudur, Fukumizu, G, Schoelkopf, Lanckriet (EJS, 2012, Theorem A.1)

Two-Sample Testing with MMD

A statistical test using MMD

The empirical MMD:

$$egin{aligned} \widehat{MMD}^2 =& rac{1}{n(n-1)} \sum_{i
eq j} k(\pmb{x}_i, \pmb{x}_j) + rac{1}{n(n-1)} \sum_{i
eq j} k(\pmb{y}_i, \pmb{y}_j) \ &- rac{2}{n^2} \sum_{i,j} k(\pmb{x}_i, \pmb{y}_j) \end{aligned}$$

How does this help decide whether P = Q?

A statistical test using MMD

The empirical MMD:

$$egin{aligned} \widehat{MMD}^2 =& rac{1}{n(n-1)}\sum_{i
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Perspective from statistical hypothesis testing:

Null hypothesis H₀ when P = Q
should see MMD² "close to zero".
Alternative hypothesis H₁ when P ≠ Q
should see MMD² "far from zero"

A statistical test using MMD

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Perspective from statistical hypothesis testing:

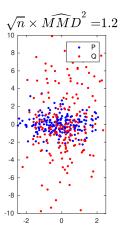
Null hypothesis H₀ when P = Q
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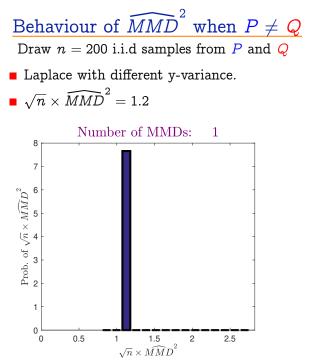
Want Threshold c_{α} for \widehat{MMD}^2 to get false positive rate α

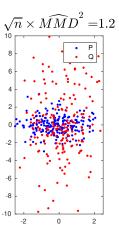
Draw n = 200 i.i.d samples from P and Q

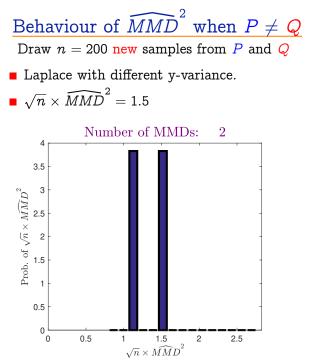
• Laplace with different y-variance.

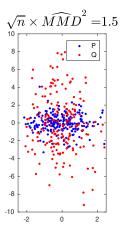
 $\sqrt{n} \times \widehat{MMD}^2 = 1.2$





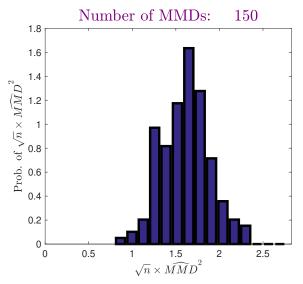




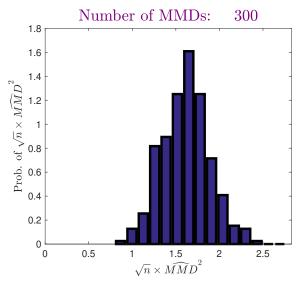


40/89

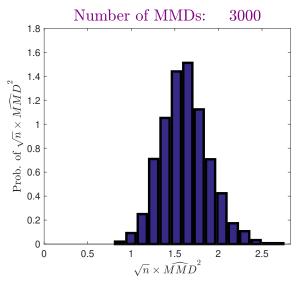
Repeat this 150 times ...



Repeat this 300 times ...



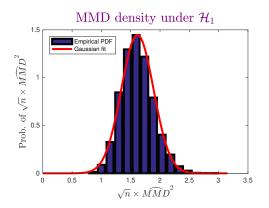
Repeat this 3000 times ...

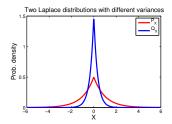


Asymptotics of \widehat{MMD}^2 when $P \neq Q$

When $P \neq Q$, statistic is asymptotically normal, $\frac{\widehat{\mathrm{MMD}}^2 - \mathrm{MMD}^2(P, Q)}{\sqrt{V_n(P, Q)}} \xrightarrow{D} \mathcal{N}(0, 1),$

where variance $V_n(P,Q) = O(n^{-1})$.

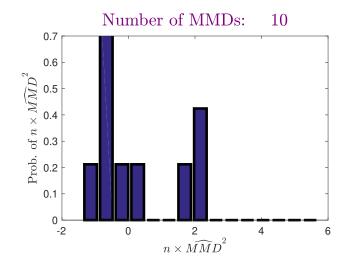


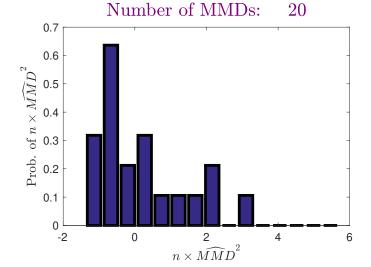


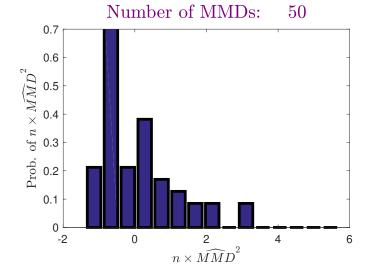


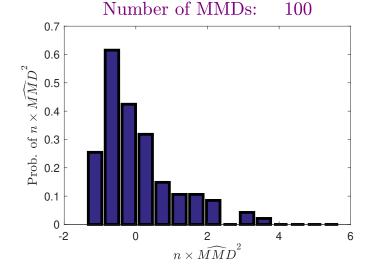
What happens when P and Q are the same?

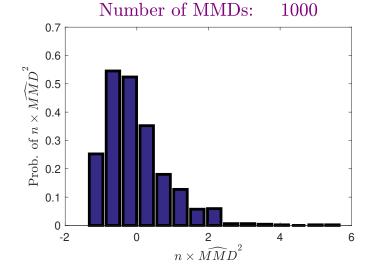








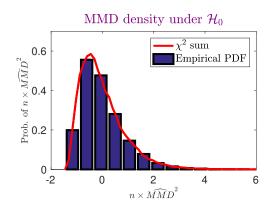




Asymptotics of \widehat{MMD}^2 when P = Q

Where P = Q, statistic has asymptotic distribution

$$\widehat{n ext{MMD}}^2\sim\sum_{l=1}^\infty\lambda_l\left[z_l^2-2
ight]$$

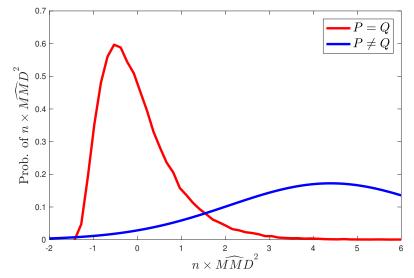


where

$$\lambda_i \psi_i(x') = \int_{\mathcal{X}} \underbrace{ ilde{k}(x,x')}_{ ext{centred}} \psi_i(x) dP(x)$$

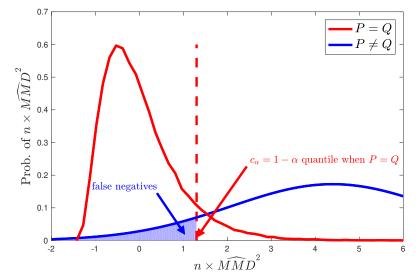
$$z_l \sim \mathcal{N}(0,2)$$
 i.i.d.

A summary of the asymptotics:



A statistical test

Test construction: (G., Borgwardt, Rasch, Schoelkopf, and Smola, JMLR 2012)



How do we get test threshold c_{α} ?

Permuted dog and fish samples (merdogs):





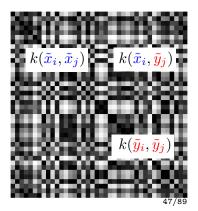
How do we get test threshold c_{α} ?

Permuted dog and fish samples (merdogs):

$$\widetilde{X} = \llbracket \bigotimes \bigotimes \bigotimes \bigotimes \ldots \rrbracket$$
$$\widetilde{Y} = \llbracket \bigotimes \bigotimes \bigotimes \bigotimes \bigotimes \bigotimes \ldots \rrbracket$$

$$egin{aligned} \widehat{MMD}^2 =& rac{1}{n(n-1)}\sum_{i
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Permutation simulates P = Q



How do we get test threshold c_{α} ?

Permuted dog and fish samples (merdogs):

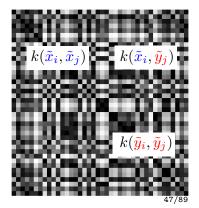


Exact level α (upper bound on false positive rate) at finite *n* and number of permutations (when unpermuted statistic included in pool)

Proposition 1, Schrab, Kim, Albert, Lau-

rent, Guedj, Gretton (2021), MMD Aggre-

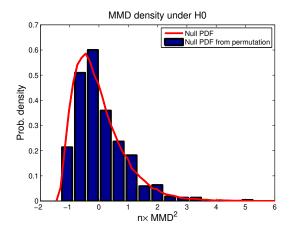
gated Two-Sample Test, arXiv:2110.15073



Approx. null distribution of \widehat{MMD}^2 via permutation

Null distribution estimated from 500 permutations

Example: $P = Q = \mathcal{N}(0, 1)$



Consistent test w/o bootstrap (not examinable)

Maximum mean discrepancy (MMD):

$$MMD^2(P, \boldsymbol{Q}; \mathcal{F}) = \left\| \boldsymbol{\mu}_P - \boldsymbol{\mu}_{\boldsymbol{Q}} \right\|_{\mathcal{F}}^2$$

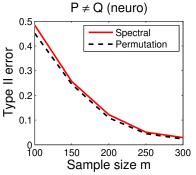
Is $\widehat{\text{MMD}}^2$ significantly > 0?

P = Q, null distrib. of $\widehat{\text{MMD}}$:

$$\widehat{\mathrm{MMD}} \stackrel{
ightarrow}{
ightarrow} \sum_{l=1}^\infty \lambda_l (z_l^2-2),$$

 λ_l is *l*th eigenvalue of centered kernel $ilde{k}(x_i, x_j)$

Use Gram matrix spectrum for $\hat{\lambda}_l$: consistent test without permutation



Gretton, Fukumizu, Harchaoui, Sriperumbudur, A fast, consistent kernel two-sample test, NeurIPS (2009)

How to choose the best kernel (1) characteristic kernels

Characteristic: MMD a metric MMD = 0 iff P = Q) [NeurIPS07b, JMLR10]

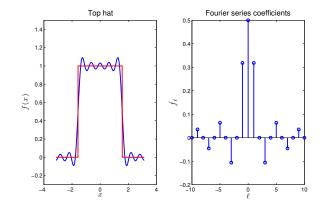
In the next slides:

- Characteristic property on $[-\pi, \pi]$ with periodic boundary
- Characteristic property on \mathbb{R}^d
- Characteristic property via Universality

Reminder: Fourier series

Function on $[-\pi, \pi]$ with periodic boundary.

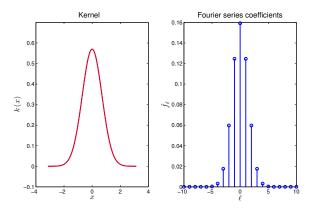
$$f(x) = \sum_{\ell=-\infty}^{\infty} \hat{f}_{\ell} \exp(\imath \ell x) = \sum_{l=-\infty}^{\infty} \hat{f}_{\ell} \left(\cos(\ell x) + \imath \sin(\ell x) \right).$$



Jacobi theta kernel (close to exponentiated quadratic):

$$k(x-y) = rac{1}{2\pi}artheta\left(rac{x-y}{2\pi},rac{\imath\sigma^2}{2\pi}
ight), \qquad \hat{k}_\ell = rac{1}{2\pi}\exp\left(rac{-\sigma^2\ell^2}{2}
ight)$$

 ϑ is the Jacobi theta function, close to Gaussian when σ^2 small



- Fourier series for P is characteristic function $\varphi_{P,\ell}$
- Fourier series for mean embedding is product of fourier series! (convolution theorem)

$$egin{aligned} \mu_P(x) &= \langle \mu_P, k(\cdot, x)
angle_{\mathcal{F}} \ &= E_{t \sim P} k(t-x) \ &= \int_{-\pi}^{\pi} k(t-x) dP(t) \qquad \hat{\mu}_{P,\ell} = \hat{k}_\ell imes ar{arphi}_{P,\ell} \end{aligned}$$

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Maximum mean embedding via Fourier series:

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MMD can be written in terms of Fourier series:

$$MMD(P, Q; \mathcal{F}) = \left\| \mu_P - \mu_Q \right\|_{\mathcal{F}} \ = \left\| \sum_{\ell = -\infty}^{\infty} \left[\left(\bar{\varphi}_{P,\ell} - \bar{\varphi}_{Q,\ell} \right) \hat{k}_\ell \right] \exp(\imath \ell x) \right\|_{\mathcal{F}}$$

A simpler Fourier representation for MMD

From previous slide,

$$MMD(P, \boldsymbol{Q}; \mathcal{F}) = \left\|\sum_{\ell=-\infty}^{\infty} \left[\left(ar{arphi}_{P,\ell} - ar{arphi}_{\boldsymbol{Q},\ell}
ight) \hat{k}_{\ell}
ight] \exp(\imath \ell x)
ight\|_{\mathcal{F}}$$

Reminder: the squared norm of a function f in \mathcal{F} is:

$$\|f\|_{\mathcal{F}}^2 = \sum_{l=-\infty}^\infty rac{|\widehat{f}_\ell|^2}{\widehat{k}_\ell}.$$

Simple, interpretable expression for squared MMD:

$$MMD^2(\pmb{P}, \pmb{Q}; \mathcal{F}) = \sum_{\ell=-\infty}^\infty rac{|arphi_{P,\ell} - arphi_{Q,\ell}|^2 \hat{k}_\ell^2}{\hat{k}_\ell} = \sum_{\ell=-\infty}^\infty |arphi_{P,\ell} - arphi_{Q,\ell}|^2 \hat{k}_\ell$$

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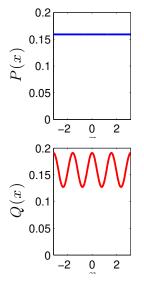
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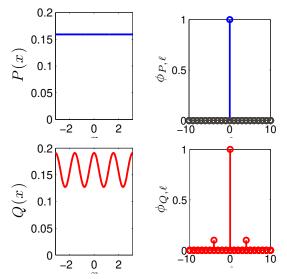
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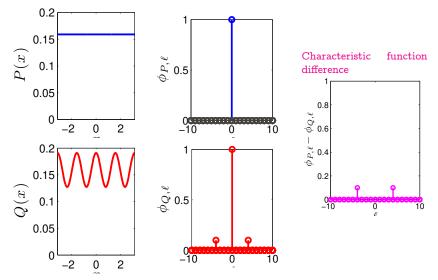
Example: P differs from Q at one frequency:



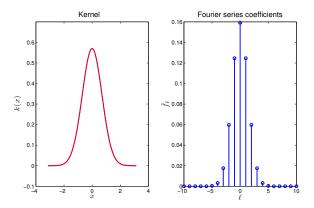
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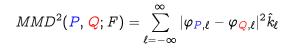


Example: P differs from Q at one frequency:



Is the Gaussian spectrum kernel characteristic?

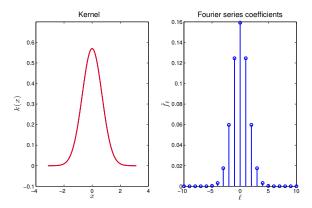


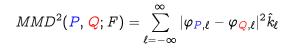


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Characteristic kernels on $[-\pi,\pi]$

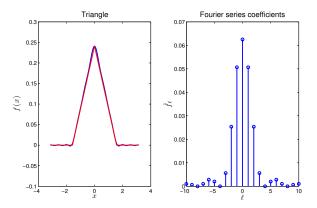
Is the Gaussian spectrum kernel characteristic? YES

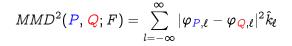




Characteristic kernels on $[-\pi,\pi]$

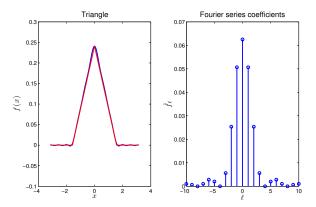
Is the triangle kernel characteristic?

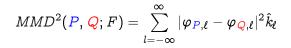




Characteristic kernels on $[-\pi,\pi]$

Is the triangle kernel characteristic? NO





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Can we prove characteristic on \mathbb{R}^d ?

Characteristic function of P via Fourier transform

$$arphi_P(\omega) = \int_{\mathbb{R}^d} e^{ix^ op \omega} dP(x)$$

For translation invariant kernels: k(x, y) = k(x - y), Bochner's theorem:

$$k(x-y) = \int_{\mathbb{R}^d} e^{-i(x-y)^ op \omega} d\Lambda(\omega) \; .$$

 $\Lambda(\omega)$ finite non-negative Borel measure.

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Fourier representation of MMD on \mathbb{R}^d :

$$MMD^2(P, oldsymbol{Q}; F) = \int \left| arphi_P(\omega) - arphi_{oldsymbol{Q}}(\omega)
ight|^2 d\Lambda(\omega)$$

Proof:

$$\begin{split} MMD^2(P,Q;F) &:= E_P k(x-x') + E_Q k(y-y') - 2E_{P,Q} k(x-y) \\ &= \int \int \left[k(s-t) d(P-Q)(s) \right] d(P-Q)(t) \\ &\stackrel{(a)}{=} \int \int_{\mathbb{R}^d} e^{-i(s-t)^T \omega} d\Lambda(\omega) d(P-Q)(s) d(P-Q)(t) \\ &\stackrel{(b)}{=} \int \int_{\mathbb{R}^d} e^{-is^T \omega} d(P-Q)(s) \int_{\mathbb{R}^d} e^{it^T \omega} d(P-Q)(t) d\Lambda(\omega) \\ &= \int_{\mathbb{R}^d} |\phi_P(\omega) - \phi_Q(\omega)|^2 d\Lambda(\omega) \end{split}$$

Fourier representation of MMD on \mathbb{R}^d :

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ight] d(P-Q)(t)\ &\stackrel{(a)}{=} \int \int \int_{\mathbb{R}^d} e^{-i(s-t)^T \omega} \ d\Lambda(\omega) \ d(P-Q)(s) \ d(P-Q)(t)\ &\stackrel{(b)}{=} \int \int_{\mathbb{R}^d} e^{-is^T \omega} \ d(P-Q)(s) \ \int_{\mathbb{R}^d} e^{it^T \omega} \ d(P-Q)(t) \ d\Lambda(\omega)\ &= \int_{\mathbb{R}^d} |\phi_P(\omega) - \phi_Q(\omega)|^2 \ d\Lambda(\omega) \end{aligned}$$

Fourier representation of MMD on \mathbb{R}^d :

$$MMD^2(P, oldsymbol{Q}; F) = \int \left| arphi_P(\omega) - arphi_{oldsymbol{Q}}(\omega)
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(a) Using Bochner's theorem...

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(a) Using Bochner's theorem.....(b) and using Fubini's theorem. 60/89

Fourier representation of MMD on \mathbb{R}^d :

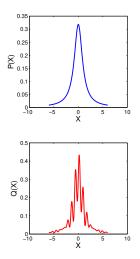
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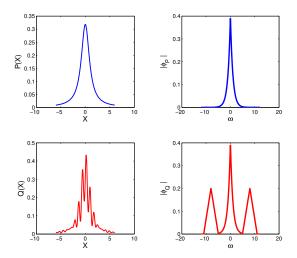
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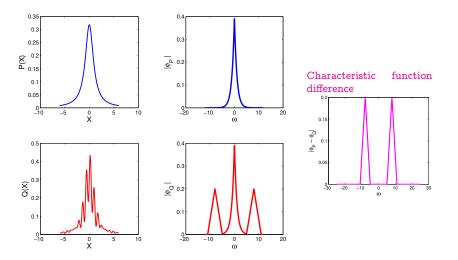
Example: P differs from Q at roughly one frequency:



Example: P differs from Q at roughly one frequency:



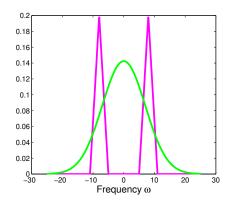
Example: P differs from Q at roughly one frequency:



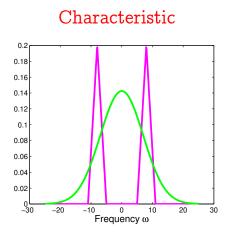
Example: P differs from Q at (roughly) one frequency:

Exponentiated quadraric kernel spectrum $\Lambda(\omega)$

Difference $|\varphi_P - \varphi_Q|$



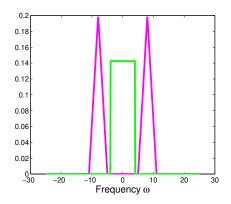
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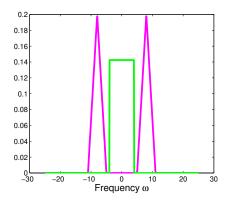
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Sinc kernel spectrum \Lambda(\omega)
```

Difference $|\varphi_P - \varphi_Q|$



Example: P differs from Q at (roughly) one frequency:

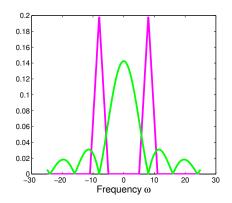
Not characteristic



Example: P differs from Q at (roughly) one frequency:

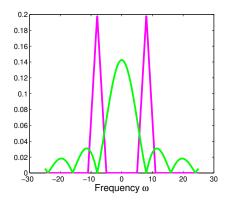
Triangle (B-spline) kernel spectrum $\Lambda(\omega)$

Difference $|\phi_P - \phi_Q|$

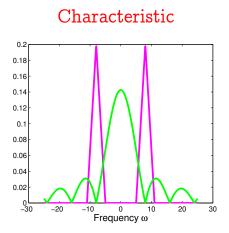


Example: P differs from Q at (roughly) one frequency:

???

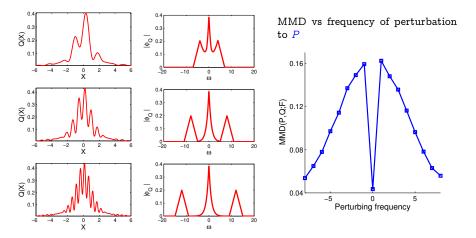


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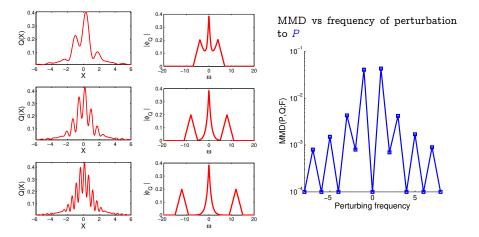
Choosing the best kernel (Fourier)

Exponentiated quadratic kernel:



Choosing the best kernel (Fourier)

B-Spline kernel:



MMD decay with increasing perturbation freq.

Recall simple MMD, Fourier series on $[-\pi, \pi]$:

$$MMD^2(P, oldsymbol{Q}; \mathcal{F}) = \sum_{\ell=-\infty}^\infty |arphi_{P,\ell} - arphi_{oldsymbol{Q},\ell}|^2 \hat{k}_\ell$$

where \hat{k}_{ℓ} decays as ℓ grows.

Fourier series representation for more general case on \mathbb{R}^d :

$$MMD^2(\textit{P},\textit{Q};\mathcal{F}) = \int_{\mathbb{R}^d} \left| \phi_P(\omega) - \phi_{\textit{Q}}(\omega)
ight|^2 \left. d\Lambda(\omega)
ight.$$

has similar behaviour.

Summary: characteristic kernels on \mathbb{R}^d

Characteristic kernel: MMD = 0 iff P = Q Fukumizu et al. [NIPS07b], Sriperumbudur et al.[COLT08]

Main theorem: A translation invariant k is characteristic for prob. measures on \mathbb{R}^d if and only if

 $\mathrm{supp}(\Lambda)=\mathbb{R}^d$

(i.e. support zero on at most a countable set) Sriperumbudur et al. [COLT08, JMLR10]

Corollary: any continuous, compactly supported k characteristic (since Fourier spectrum $\Lambda(\omega)$ cannot be zero on an interval). 1-D proof sketch from [Mallat, 99, Theorem 2.6], proof on \mathbb{R}^4 via distribution theory in Sriperumbudue et al. [MMLB10 Corollary 10 p. 1535]

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Characteristic kernels: MMD = 0 iff P = Q

Classical result: P = Q if and only if $E_P(f(x)) = E_Q(f(y))$ for all $f \in C(\mathcal{X})$, the space of bounded continuous functions on \mathcal{X} Dudley (2002)

Universal RKHS: k(x, x') continuous, \mathcal{X} compact, and \mathcal{F} dense in $C(\mathcal{X})$ with respect to L_{∞} Steinwart (2001)

If ${\mathcal F}$ universal, then $MMD(P, Q; {\mathcal F}) = 0$ iff P = Q

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Proof:

First, it is clear that P = Q implies $MMD(P, Q; \mathcal{F})$ is zero. Converse: by the universality of \mathcal{F} , for any given $\epsilon > 0$ and $f \in C(\mathcal{X})$, $\exists g \in \mathcal{F}$

$$\|f-g\|_{\infty}\leq\epsilon.$$

We next make the expansion

 $egin{aligned} &|\mathrm{E}_P f(x) - \mathrm{E}_{\mathcal{Q}} f(y)| \ &\leq |\mathrm{E}_P f(x) - \mathrm{E}_P g(x)| + |\mathrm{E}_P g(x) - \mathrm{E}_{\mathcal{Q}} g(y)| + |\mathrm{E}_{\mathcal{Q}} g(y) - \mathrm{E}_{\mathcal{Q}} f(y)| \,. \end{aligned}$

The first and third terms satisfy

 $|\mathrm{E}_P f(x) - \mathrm{E}_P g(x)| \leq \mathrm{E}_P \left|f(x) - g(x)
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$$|\mathrm{E}_P f(x) - \mathrm{E}_P g(x)| \leq \mathrm{E}_P \left| f(x) - g(x)
ight| \leq \epsilon.$$

Proof (continued):

$$\mathrm{E}_P g(x) - \mathrm{E}_{oldsymbol{Q}} g(y) = \left\langle g(\cdot), \mu_P - \mu_{oldsymbol{Q}}
ight
angle_{\mathcal{F}} = 0,$$

since $MMD(P, Q; \mathcal{F}) = 0$ implies $\mu_P = \mu_Q$. Hence

$$|\mathrm{E}_{P}f(x) - \mathrm{E}_{\mathcal{Q}}f(y)| \leq 2\epsilon$$

for all $f \in C(\mathcal{X})$ and $\epsilon > 0$, which implies P = Q.

How to choose the best kernel (2) optimising the kernel parameters

- A test's power depends on k(x, x'), P, and Q (and n)
- With characteristic kernel, MMD test has power $\rightarrow 1$ as $n \rightarrow \infty$ for any (fixed) problem
 - But, for many P and Q, will have terrible power with reasonable n!

- A test's power depends on k(x, x'), P, and Q (and n)
- With characteristic kernel, MMD test has power $\rightarrow 1$ as $n \rightarrow \infty$ for any (fixed) problem
 - But, for many P and Q, will have terrible power with reasonable n!
- You *can* choose a good kernel for a given problem
- You *can't* get one kernel that has good finite-sample power for all problems

Simple choice: exponentiated quadratic

$$k(x,y) = \exp\left(-rac{1}{2\sigma^2}\|x-y\|^2
ight)$$

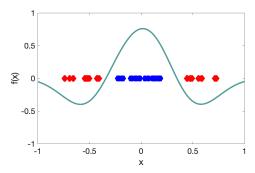
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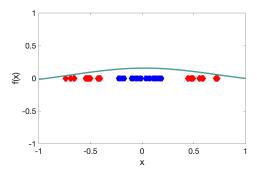
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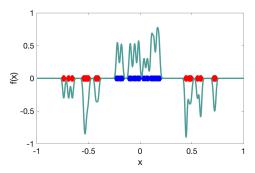
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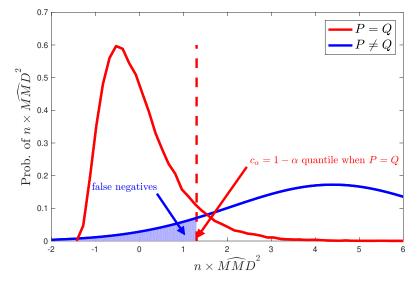
$$k(x,y) = \exp\left(-rac{1}{2\sigma^2}\|x-y\|^2
ight)$$

Characteristic: for any σ: for any P and Q, power → 1 as n → ∞
But choice of σ is very important for finite n...

• ... and some problems (e.g. images) might have no good choice for σ

Graphical illustration

• Maximising test power same as minimizing false negatives



The power of our test (Pr₁ denotes probability under $P \neq Q$):

$$\Pr_1\left(n\widehat{\mathrm{MMD}}^2 > \hat{c}_{oldsymbol{lpha}}
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ight) \end{aligned}$$

where

- Φ is the CDF of the standard normal distribution.
- \hat{c}_{α} is an estimate of c_{α} test threshold.

The power of our test (Pr₁ denotes probability under $P \neq Q$):

$$\Pr_{1}\left(n\widehat{\mathrm{MMD}}^{2} > \hat{c}_{\alpha}\right) \\ \rightarrow \Phi\left(\underbrace{\frac{\mathrm{MMD}^{2}(P,Q)}{\sqrt{V_{n}(P,Q)}}}_{O(n^{1/2})} - \underbrace{\frac{c_{\alpha}}{n\sqrt{V_{n}(P,Q)}}}_{O(n^{-1/2})}\right)$$

For large n, second term negligible!

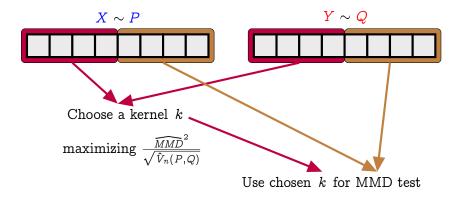
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To maximize test power, maximize

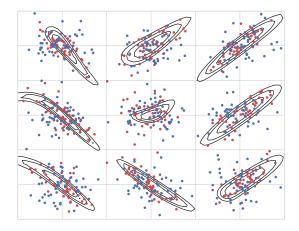
$$\frac{\mathrm{MMD}^2(P,Q)}{\sqrt{V_n(P,Q)}}$$

Data splitting



Learning a kernel helps a lot

Kernel with deep learned features: $k_{\theta}(x, y) = [(1 - \epsilon)\kappa(\Phi_{\theta}(x), \Phi_{\theta}(y)) + \epsilon] q(x, y)$ κ and q are Gaussian kernels



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■ CIFAR-10 vs CIFAR-10.1, null rejected 75% of time



CIFAR-10 test set (Krizhevsky 2009) $X \sim P$



CIFAR-10.1 (Recht+ ICML 2019)

 $Y \sim Q$

Learning a kernel helps a lot

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■ CIFAR-10 vs CIFAR-10.1, null rejected 75% of time

arXiv.org > stat > arXiv:2002.09116

Statistics > Machine Learning

[Submitted on 21 Feb 2020]

Learning Deep Kernels for Non-Parametric Two-Sample Tests

Feng Liu, Wenkai Xu, Jie Lu, Guangquan Zhang, Arthur Gretton, D. J. Sutherland

ICML 2020

How to choose the best kernel (2) minimax test without data splitting

Two-sample problem

Our aim: find a condition on $\|p - q\|_2$ to control Type II error β $\mathbb{P}_{p \times q}(\Delta(\mathbb{X}_m, \mathbb{Y}_n) = 0) \leq \beta$

Remaining settings:

samples X_m := (X₁,..., X_m), X_i $\stackrel{\text{iid}}{\sim} p$ in R^d
samples Y_n := (Y₁,..., Y_n), Y_i $\stackrel{\text{iid}}{\sim} q$ in R^d

 $\begin{array}{ll} \mathcal{H}_0 \colon p = q & \text{against} & \mathcal{H}_1 \colon p \neq q \\ \Delta(\mathbb{X}_m, \mathbb{Y}_n) = 1 & \Longleftrightarrow & \text{reject } \mathcal{H}_0 \\ \Delta(\mathbb{X}_m, \mathbb{Y}_n) = 0 & \Longleftrightarrow & \text{fail to reject } \mathcal{H}_0 \end{array}$

Type I error: controlled by α by design

 $\mathbb{P}_{p imes p}(\Delta(\mathbb{X}_m,\mathbb{Y}_n)=1) \ \leq \ lpha$

79/89

Two-sample problem

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$$\mathbb{P}_{p \times p}(\Delta(\mathbb{X}_m, \mathbb{Y}_n) = 1) \leq \alpha$$
 79/89

Kernels and bandwidths

$$\text{Kernel:} \ k_{\!\lambda}(\boldsymbol{x},\boldsymbol{y}) \coloneqq \prod_{i=1}^d K_i\!\left(\frac{x_i-y_i}{\lambda_i}\right)$$

Bandwidth: $\lambda \in (0, \infty)^d$

Assumptions: K_1, \ldots, K_d integrable and square integrable

Examples: Gaussian $(K_i(u) = e^{-u^2})$, Laplace $(K_i(u) = e^{-|u|})$, Matérn

 $ext{Gaussian kernel: } k_{\lambda}(\boldsymbol{x}, \boldsymbol{y}) \coloneqq \exp \Big(- \sum rac{(x_i - y_i)^2}{\lambda_i^2} \Big)$

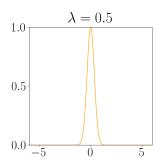
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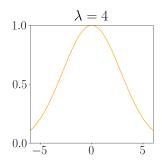
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Bonferroni multiple testing: non-asymptotic level α

$$\Delta^{\boldsymbol{\Lambda}}_{\boldsymbol{\alpha}}(\mathbb{X}_m,\mathbb{Y}_n) \coloneqq \mathbb{1}\left(\widehat{\mathrm{MMD}}_{\boldsymbol{\lambda}}^2(\mathbb{X}_m,\mathbb{Y}_n) > \widehat{q}_{1-\boldsymbol{\alpha}/|\boldsymbol{\Lambda}|}^{\boldsymbol{\lambda}} \text{ for some } \boldsymbol{\lambda} \in \boldsymbol{\Lambda}\right)$$

time complexity $\mathcal{O}(|\Lambda| B_1 (m+n)^2)$

 B_1 permutations for each $\lambda \in \Lambda$ to obtain thresholds $\widehat{q}_{1-\alpha/|\Lambda|}^{\lambda}$

Bonferroni multiple testing: non-asymptotic level α

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time complexity $\mathcal{O}(|\Lambda| B_1 (m+n)^2)$

MMDAgg (MMD Aggregation): non-asymptotic level α

$$\Delta^{\Lambda}_{\alpha}(\mathbb{X}_m,\mathbb{Y}_n) \coloneqq \mathbb{1}\left(\widehat{\mathrm{MMD}}^2_{\lambda}(\mathbb{X}_m,\mathbb{Y}_n) > \widehat{q}^{\lambda}_{1-u_{\alpha}/|\Lambda} \text{ for some } \lambda \in \Lambda\right)$$

Bonferroni multiple testing: non-asymptotic level α

$$\Delta^{\Lambda}_{\alpha}(\mathbb{X}_m,\mathbb{Y}_n) := \mathbb{1}\left(\widehat{\mathrm{MMD}}^2_{\lambda}(\mathbb{X}_m,\mathbb{Y}_n) > \widehat{q}^{\lambda}_{1-\alpha/|\Lambda|} \text{ for some } \lambda \in \Lambda\right)$$

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Correction u_{α} defined as

$$\sup \biggl\{ u > 0 : \mathbb{P}_{p \times p} \biggl(\max_{\lambda \in \Lambda} \left(\widehat{\mathrm{MMD}}_{\lambda}^2(\mathbb{X}_m, \mathbb{Y}_n) - \widehat{q}_{1-u/|\Lambda}^{\lambda} \right) > 0 \biggr) \leq \alpha \biggr\}$$

more powerful than Bonferroni correction as $u_{lpha} \geq lpha$

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$$\Delta^{\Lambda}_{\alpha}(\mathbb{X}_m,\mathbb{Y}_n) := \mathbb{1}\left(\widehat{\mathrm{MMD}}^2_{\lambda}(\mathbb{X}_m,\mathbb{Y}_n) > \widehat{q}^{\lambda}_{1-\alpha/|\Lambda|} \text{ for some } \lambda \in \Lambda\right)$$

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more powerful than Bonferroni correction as $u_{\alpha} \ge \alpha$ B_2 permutations to simulate draws from $\mathbb{P}_{p \times p}$, search over u

Bonferroni multiple testing: non-asymptotic level α

$$\Delta^{\Lambda}_{\alpha}(\mathbb{X}_m,\mathbb{Y}_n) := \mathbb{1}\left(\widehat{\mathrm{MMD}}^2_{\lambda}(\mathbb{X}_m,\mathbb{Y}_n) > \widehat{q}^{\lambda}_{1-\alpha/|\Lambda|} \text{ for some } \lambda \in \Lambda\right)$$

time complexity $\mathcal{O}(|\Lambda| B_1 (m+n)^2)$

MMDAgg (MMD Aggregation): non-asymptotic level α

$$\Delta^{\Lambda}_{\alpha}(\mathbb{X}_{m},\mathbb{Y}_{n}) \coloneqq \mathbb{1}\left(\widehat{\mathrm{MMD}}^{2}_{\lambda}(\mathbb{X}_{m},\mathbb{Y}_{n}) > \widehat{q}^{\lambda}_{1-u_{\alpha}/|\Lambda} \text{ for some } \lambda \in \Lambda\right)$$

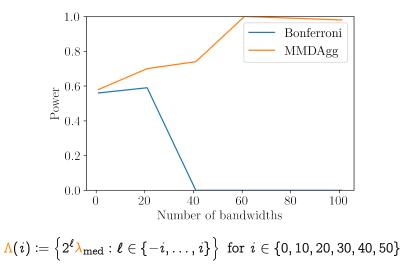
Correction u_{α} defined as

$$\sup \biggl\{ u > 0 : \mathbb{P}_{p \times p} \biggl(\max_{\lambda \in \Lambda} \left(\widehat{\mathrm{MMD}}_{\lambda}^2(\mathbb{X}_m, \mathbb{Y}_n) - \widehat{q}_{1-u/|\Lambda}^{\lambda} \right) > 0 \biggr) \leq \alpha \biggr\}$$

more powerful than Bonferroni correction as $u_{lpha} \geq lpha$ Time complexity $\mathcal{O}ig(|\Lambda| (B_1 + B_2) (m + n)^2ig)$

Multiple testing correction comparison

Simple example: 3-d Gaussians with different means



MMDAgg parameter-free user-friendly implementation

mmdagg package: github.com/antoninschrab/mmdagg

from mmdagg import mmdagg # X shape (m, d) output = mmdagg(X, Y) # 0 or 1 # Y shape (n, d)

JAX: runs on either CPU or GPU (significant speed improvements)JAX GPU runs 100 times faster than Numpy CPU



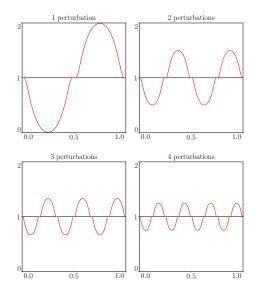
Statistics > Machine Learning

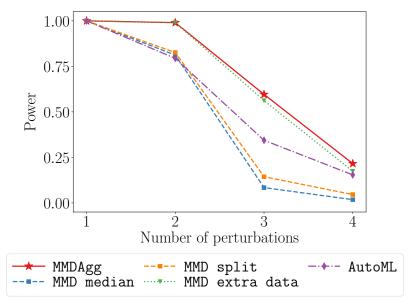
[Submitted on 28 Oct 2021 (v1), last revised 21 Aug 2023 (this version, v4)]

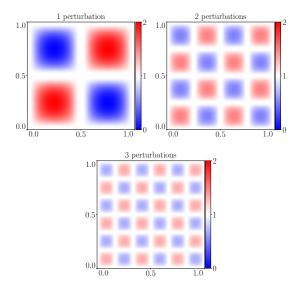
MMD Aggregated Two-Sample Test

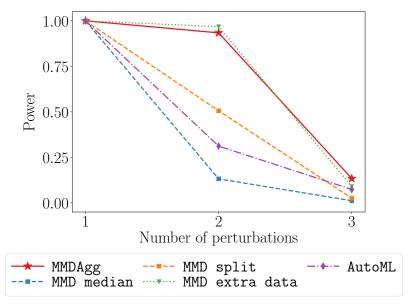
Antonin Schrab, Ilmun Kim, Mélisande Albert, Béatrice Laurent, Benjamin Guedj, Arthur Gretton

JMLR 2023









Experiment on image shifts on MNIST & CIFAR-10

Failing Loudly Benchmark: Rabanser et al., 2019











Experiment on image shifts on MNIST & CIFAR-10

