Probability Divergences and Generative Models

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Training generative models

- Have: One collection of samples X from unknown distribution P.
- Goal: generate samples Q that look like P



LSUN bedroom samples P



Generated Q, MMD GAN

Role of divergence D(P, Q)?

Testing for differences in samples

Given samples $X \sim P$ and $Y \sim Q$, are P and Q distinguishable (via D(P, Q))?

■ Application: detecting domain shift (did I train for the right task?)



CIFAR-10 test set (Krizhevsky 2009) $X \sim P$



CIFAR-10.1 (Recht+ ICML 2019)



Outline

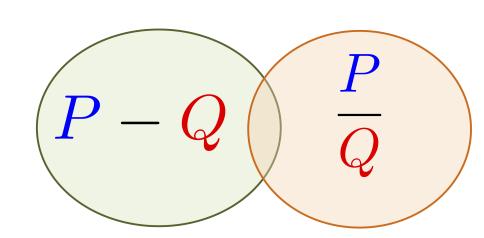
- Integral probability metrics (MMD, Wasserstein)
- lacktriangledown ϕ -divergences (f-divergences) and a variational lower bound (KL)
- Generalized energy-based models
 - "Like a GAN" but incorporate critic into sample generation
 - Perform better than using generator alone

Arbel, Zhou, G., Generalized Energy Based Models (ICLR 2021)

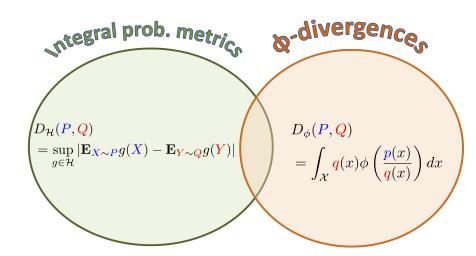
■ Comparing samples with MMD

Liu, Xu, Lu, Zhang, G. Sutherland, Learning Deep Kernels for Non-Parametric Two-Sample Tests (ICML 2020) Divergence measures (critics)

Divergences



Divergences



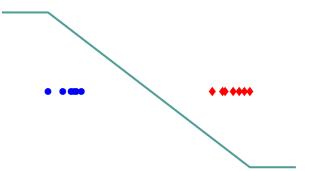
The Integral Probability Metrics

Wasserstein distance



A helpful critic witness:

$$egin{align} W_1ig(P, rac{m{Q}}{m{Q}}ig) &= \sup_{\|f\|_L \leq 1} E_P fig(Xig) - E_{m{Q}} fig(Yig). \ \|f\|_L &:= \sup_{x
eq y} |f(x) - f(y)| / \|x - y\| \ W_1 = 0.88 \ \end{split}$$



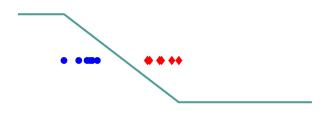
Santambrogio, Optimal Transport for Applied Mathematicians (2015, Section 5.4) G Peyré, M Cuturi, Computational Optimal Transport (2019) M. Cuturi, J. Solomon, NeurIPS tutorial (2017)

Wasserstein distance



A helpful critic witness:

$$egin{align} W_1(P, \begin{subarray}{c} Q \end{subarray} &= \sup_{\|f\|_L \leq 1} E_P f(X) - E_{oldsymbol{Q}} f(\begin{subarray}{c} Y \end{subarray}. \ &\|f\|_L := \sup_{x
eq y} |f(x) - f(y)| / \|x - y\| \ &W_1 = 0.65 \end{aligned}$$



Santambrogio, Optimal Transport for Applied Mathematicians (2015, Section 5.4) G Peyré, M Cuturi, Computational Optimal Transport (2019) M. Cuturi, J. Solomon, NeurIPS tutorial (2017)

The Maximum Mean Discrepancy

Maximum mean discrepancy: smooth function for P vs Q

$$MMD(P, Q; F) := \sup_{\|f\| \le 1} [\mathbb{E}_P f(X) - \mathbb{E}_Q f(Y)]$$

$$(F = \text{unit ball in RKHS } \mathcal{F})$$

The Maximum Mean Discrepancy

Maximum mean discrepancy: smooth function for P vs Q

$$egin{aligned} MMD(P, m{\mathcal{Q}}; F) := \sup_{\|f\| \leq 1} \left[\mathrm{E}_P f(X) - \mathrm{E}_{m{\mathcal{Q}}} f(m{Y})
ight] \ (F = \mathrm{unit} \,\, \mathrm{ball} \,\, \mathrm{in} \,\, \mathrm{RKHS} \,\, \mathcal{F}) \end{aligned}$$

Functions are linear combinations of features:

$$f(x) = \langle f, \varphi(x) \rangle_{\mathcal{F}} = \sum_{\ell=1}^{\infty} f_{\ell} \varphi_{\ell}(x) = \begin{bmatrix} f_1 \\ f_2 \\ f_3 \\ \vdots \end{bmatrix}^{\top} \begin{bmatrix} \varphi_{1}(x) \\ \varphi_{2}(x) \\ \vdots \\ \varphi_{3}(x) \end{bmatrix}$$

$$||f||_{\mathcal{T}}^{2} := \sum_{i=1}^{\infty} f_{i}^{2} < 1$$

Infinitely many features using kernels

Kernels: dot products of features

Feature map $\varphi(x) \in \mathcal{F}$,

$$oldsymbol{arphi}(x) = [\dots arphi_i(x) \dots] \in oldsymbol{\ell}_2$$

For positive definite k,

$$k(x,x')=\langle arphi(x),arphi(x')
angle_{\mathcal{F}}$$

Infinitely many features $\varphi(x)$, dot product in closed form!

Infinitely many features using kernels

Kernels: dot products of features

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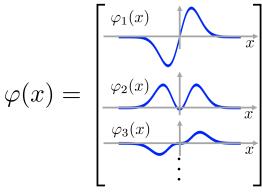
For positive definite k,

$$k(x,x') = \langle arphi(x), arphi(x')
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Infinitely many features $\varphi(x)$, dot product in closed form!

Exponentiated quadratic kernel

$$k(x, x') = \exp\left(-\gamma \left\|x - x'\right\|^2\right)$$



Features: Gaussian Processes for Machine learning, Rasmussen and Williams, Ch. 4.

The MMD: an integral probability metric

Maximum mean discrepancy: smooth function for P vs Q

$$egin{aligned} MMD(P, oldsymbol{\mathcal{Q}}; F) := \sup_{\|f\| \leq 1} \left[\operatorname{E}_P f(X) - \operatorname{E}_{oldsymbol{\mathcal{Q}}} f(oldsymbol{Y})
ight] \ (F = \operatorname{unit\ ball\ in\ RKHS\ } \mathcal{F}) \end{aligned}$$

For characteristic RKHS \mathcal{F} , MMD(P, Q; F) = 0 iff P = Q

■ Energy distance is a special case [Sejdinovic, Sriperumbudur, G. Fukumizu, 2013]

The MMD: an integral probability metric

Maximum mean discrepancy: smooth function for P vs Q

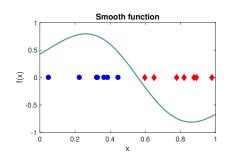
$$egin{aligned} \mathit{MMD}(P, \column{Q}{Q}; F) := & \sup_{\|f\| \leq 1} \left[\operatorname{E}_P f(X) - \operatorname{E}_{\column{Q}} f(\column{Y}{Y})
ight] \ & (F = \operatorname{unit\ ball\ in\ RKHS\ } \mathcal{F}) \end{aligned}$$

Expectations of functions are linear combinations of expected features

$$\mathrm{E}_P(f(X)) = \langle f, \mathrm{E}_P arphi(X)
angle_{\mathcal{F}} = \langle f, \mu_P
angle_{\mathcal{F}}$$

(always true if kernel is bounded)

$$egin{aligned} MMD(P, & oldsymbol{Q}; F) \ &= \sup_{\|f\| \le 1} \left[\mathbb{E}_P f(X) - \mathbb{E}_{oldsymbol{Q}} f(Y)
ight] \end{aligned}$$



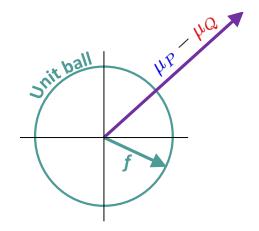
The MMD:

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ight] \ &= \sup_{\|f\| \leq 1} \left\langle f, \mu_P - \mu_{oldsymbol{Q}}
ight
angle_{\mathcal{F}} \end{aligned}$$

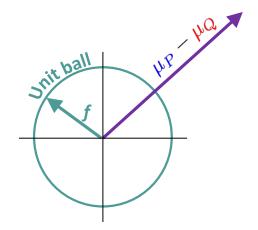
use

$$\mathbb{E}_P f(X) = \langle \mu_P, f
angle_{\mathcal{F}}$$

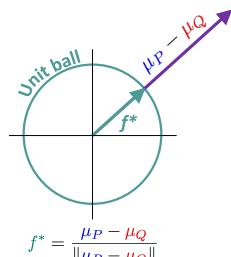
$$egin{aligned} &MMD(P, \cline{Q}; F) \ &= \sup_{\|f\| \leq 1} \left[\operatorname{E}_P f(X) - \operatorname{E}_Q f(Y)
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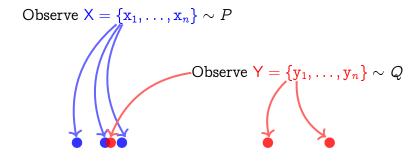


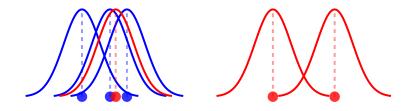
$$f^* = \frac{\mu_P - \mu_Q}{\|\mu_P - \mu_Q\|}$$

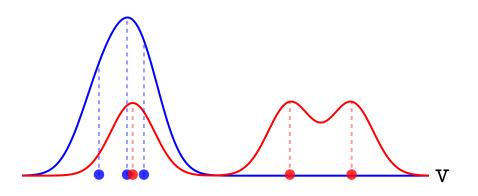
The MMD:

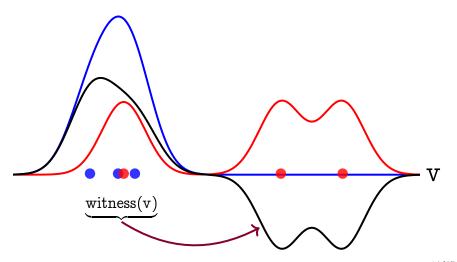
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egin{aligned} & MMD(P, \column{2}{Q}; F) \ &= \sup_{\|f\| \leq 1} \left[ \operatorname{E}_P f(X) - \operatorname{E}_Q f(\column{2}{Y}) 
ight] \ &= \sup_{\|f\| \leq 1} \left\langle f, \mu_P - \mu_Q 
ight
angle_{\mathcal{F}} \ &= \|\mu_P - \mu_Q\|_{\mathcal{F}} \end{aligned}
```

IPM view equivalent to feature mean difference (kernel case only)









Recall the witness function expression

$$f^* \propto \mu_P - \mu_Q$$

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The empirical feature mean for P

$$\widehat{\pmb{\mu}}_P := rac{1}{n} \sum_{i=1}^n arphi(x_i)$$

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$$egin{aligned} f^*(v) &= \langle f^*, arphi(v)
angle_{\mathcal{F}} \ &\propto \langle \widehat{m{\mu}}_P - \widehat{m{\mu}}_{m{\mathcal{Q}}}, arphi(v)
angle_{\mathcal{F}} \end{aligned}$$

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angle_{\mathcal{F}} \ &\propto \langle \widehat{\pmb{\mu}}_P - \widehat{\pmb{\mu}}_{m{\mathcal{Q}}}, arphi(v)
angle_{m{\mathcal{F}}} \ &= rac{1}{n} \sum_{i=1}^n k(\pmb{x}_i, v) - rac{1}{n} \sum_{i=1}^n k(\pmb{ extbf{y}}_i, v) \end{aligned}$$

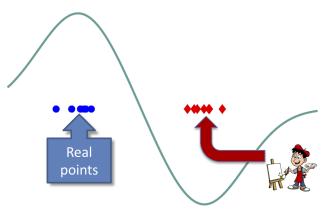
Don't need explicit feature coefficients $f^* := \begin{bmatrix} f_1^* & f_2^* & \dots \end{bmatrix}$



A helpful critic:

$$MMD(P, {\color{red} Q}) = \sup_{\|f\|_{\mathcal{F}} \leq 1} E_P f(X) - E_{{\color{red} Q}} f({\color{red} Y}).$$

MMD=1.8

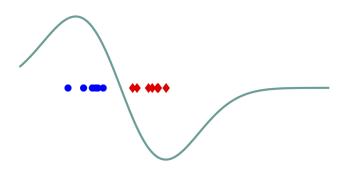




A helpful critic:

$$MMD(P, \begin{cases} Q \end{cases}) = \sup_{\|f\|_{\mathcal{F}} < 1} E_P f(X) - E_{Q} f(Y)$$

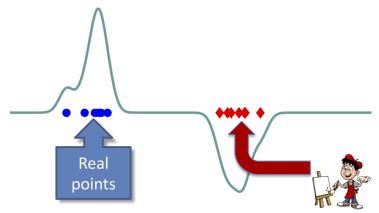
MMD=1.1





An unhelpful critic: MMD(P, Q) with a narrow kernel.

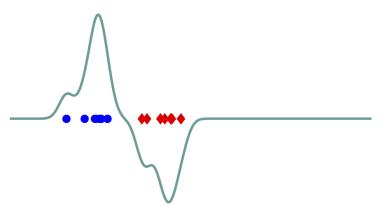
MMD=0.64



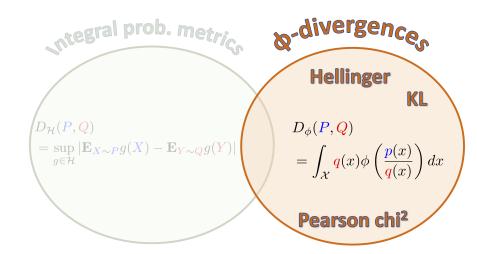


An unhelpful critic: MMD(P, Q) with a narrow kernel.

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The ϕ -divergences



The ϕ -divergences

Define the ϕ -divergence(f-divergence):

$$D_{\phi}(P, Q) = \int \phi\left(rac{p(z)}{q(z)}
ight) rac{q}{q}(z)dz$$

where ϕ is convex, lower-semicontinuous, $\phi(1) = 0$.

Example: $\phi(u) = u \log(u)$ gives KL divergence,

$$egin{aligned} D_{KL}(P,m{Q}) &= \int \log\left(rac{p(z)}{m{q}(z)}
ight) p(z) dz \ &= \int \left(rac{p(z)}{m{q}(z)}
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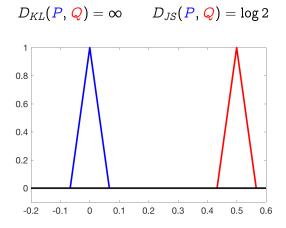
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ight) oldsymbol{q}(z) dz \end{aligned}$$

Are ϕ -divergences good critics?



Simple example: disjoint support.

Goodfellow et al. (NeurIPS 2014), Arjovsky and Bottou [ICLR 2017]

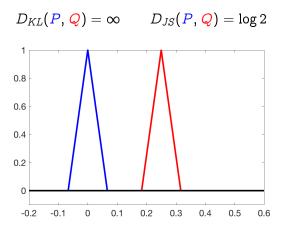


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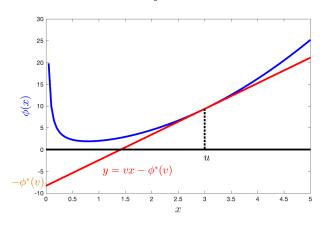
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ϕ -divergences in practice

Background: the conjugate (Fenchel) dual

$$\phi^*(v) = \sup_{u \in \mathbb{R}} \left\{ uv - \phi(u)
ight\}.$$



 $\phi^*(v)$ is negative intercept of tangent to ϕ with slope v

ϕ -divergences in practice

Background: the conjugate (Fenchel) dual

$$\phi^*(v) = \sup_{u \in \mathbb{R}} \left\{ uv - \phi(u) \right\}.$$

For a convex l.s.c. ϕ we have

$$\phi^{**}(x)=\phi(x)=\sup_{v\in\mathbb{R}}\left\{xv-\phi^{*}(v)
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ϕ -divergences in practice

Background: the conjugate (Fenchel) dual

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ight\}.$$

For a convex l.s.c. ϕ we have

$$\phi^{**}(x)=\phi(x)=\sup_{v\in\mathbb{R}}\left\{xv-\phi^*(v)
ight\}$$

■ KL divergence:

$$\phi(x) = x \log(x)$$
 $\phi^*(v) = \exp(v - 1)$

A lower-bound ϕ -divergence approximation:

$$D_{\phi}(P, Q) = \int q(z) \phi\left(rac{p(z)}{q(z)}
ight) dz$$

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$$egin{aligned} D_{\phi}(P,Q) &= \int q(z) \phi\left(rac{p(z)}{q(z)}
ight) dz \ &= \int q(z) \sup_{f_{ar{z}}} \left(rac{p(z)}{q(z)}f_{ar{z}} - \phi^*(f_{ar{z}})
ight) \ & \phi^*(v) \end{aligned}$$

 $\phi^*(v)$ is dual of $\phi(x)$.

A lower-bound ϕ -divergence approximation:

$$egin{aligned} D_{\phi}(P, oldsymbol{\mathcal{Q}}) &= \int oldsymbol{q}(z) \phi\left(rac{p(z)}{oldsymbol{q}(z)}
ight) dz \ &= \int oldsymbol{q}(z) \sup_{f_z} \left(rac{p(z)}{oldsymbol{q}(z)} f_z - \phi^*(f_z)
ight) \ &\geq \sup_{f \in \mathcal{H}} \mathrm{E}_P f(X) - \mathrm{E}_{oldsymbol{\mathcal{Q}}} \phi^*\left(f(oldsymbol{Y})
ight) \end{aligned}$$

(restrict the function class)

A lower-bound ϕ -divergence approximation:

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ight) \end{aligned}$$

(restrict the function class)

Bound tight when:

$$f^{\diamond}(z) = \partial \phi \left(rac{p(z)}{q(z)}
ight)$$

if ratio defined.

$$D_{KL}(P, rac{Q}{Q}) = \int \log \left(rac{p(z)}{rac{q}{Q}(z)}
ight) p(z) dz$$

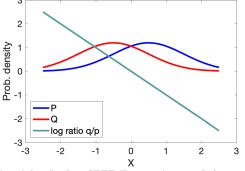
$$egin{aligned} D_{KL}(P, oldsymbol{Q}) &= \int \log\left(rac{p(z)}{oldsymbol{q}(z)}
ight) p(z) dz \ &\geq \sup_{f \in \mathcal{H}} - \mathrm{E}_P f(X) + 1 - \mathrm{E}_{oldsymbol{Q}} \underbrace{\exp\left(-f(oldsymbol{Y})
ight)}_{oldsymbol{\phi}^*(-f(oldsymbol{Y})+1)} \end{aligned}$$

$$egin{aligned} D_{KL}(P, \ oldsymbol{Q}) &= \int \log \left(rac{p(z)}{q(z)}
ight) p(z) dz \ &\geq \sup_{f \in \mathcal{H}} -\mathrm{E}_P f(X) + 1 - \mathrm{E}_{oldsymbol{Q}} \exp \left(-f(rac{oldsymbol{Y}}{Y})
ight) \end{aligned}$$

Bound tight when:

$$f^{\diamond}(z) = -\lograc{p(z)}{{m q}(z)}$$

if ratio defined.



$$egin{aligned} D_{KL}(P, oldsymbol{Q}) &= \int \log\left(rac{p(z)}{oldsymbol{q}(z)}
ight) p(z) dz \ &\geq \sup_{f \in \mathcal{H}} - \operatorname{E}_P f(X) + 1 - \operatorname{E}_{oldsymbol{Q}} \exp\left(-f(oldsymbol{Y})
ight) & x_i \overset{\mathrm{i.i.d.}}{\sim} P \ &y_i \overset{\mathrm{i.i.d.}}{\sim} oldsymbol{Q} \ &pprox \sup_{f \in \mathcal{H}} \left[-rac{1}{n} \sum_{i=1}^n f(x_i) - rac{1}{n} \sum_{i=1}^n \exp(-f(y_i))
ight] + 1 \end{aligned}$$

$$egin{aligned} D_{KL}(P, oldsymbol{Q}) &= \int \log\left(rac{p(z)}{oldsymbol{q}(z)}
ight) p(z) dz \ &\geq \sup_{f \in \mathcal{H}} -\mathrm{E}_P f(X) + 1 - \mathrm{E}_{oldsymbol{Q}} \exp\left(-f(oldsymbol{Y})
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ight] + 1 \end{aligned}$$

This is a

KL

Approximate

Lower-bound

Estimator.

$$egin{aligned} D_{KL}(P, oldsymbol{Q}) &= \int \log\left(rac{p(z)}{oldsymbol{q}(z)}
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This is a

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The KALE divergence



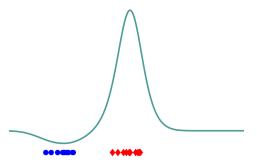
$$egin{aligned} \mathit{KALE}(P, \column{Q}; \mathcal{H}) &= \sup_{f \in \mathcal{H}} -E_P f(X) - E_{\column{Q}} \exp\left(-f(\column{Y})
ight) + 1 \ & \ f = \langle w, \phi(x)
angle_{\mathcal{H}} & \mathcal{H} \text{ an RKHS} \ & \|w\|_{\mathcal{H}}^2 & ext{penalized} : \end{aligned}$$



$$egin{aligned} \mathit{KALE}(P, \column{Q}; \mathcal{H}) &= \sup_{f \in \mathcal{H}} -E_P f(X) - E_{\column{Q}} \exp\left(-f(\column{Y})
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angle_{\mathcal{H}} & \mathcal{H} \text{ an RKHS} \ & \|w\|_{\mathcal{H}}^2 & ext{penalized} : ext{KALE smoothie} \end{aligned}$$



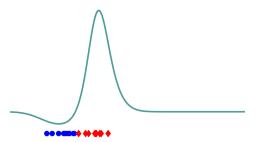
$$KALE(P, \colongledown) = \sup_{f \in \mathcal{H}} -E_P f(X) - E_{\colongledown} \exp\left(-f(\colongledown) + 1
ight)$$
 $f = \langle w, \phi(x) \rangle_{\mathcal{H}} \qquad \mathcal{H} \text{ an RKHS}$
 $\|w\|_{\mathcal{H}}^2 \quad \text{penalized} : KALE \text{ smoothie}$
 $KALE(\colongledown), P; \mathcal{H}) = 0.18$



Glaser, Arbel, G. "KALE Flow: A Relaxed KL Gradient Flow for Probabilities with Disjoint Support," (arXiv, 2021, Section 2)



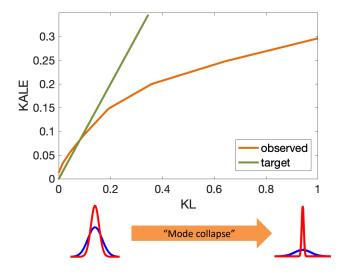
$$KALE(P, \colongledown) = \sup_{f \in \mathcal{H}} -E_P f(X) - E_{\colongledown} \exp\left(-f(\colongledown) + 1
ight)$$
 $f = \langle w, \phi(x) \rangle_{\mathcal{H}} \qquad \mathcal{H} \text{ an RKHS}$
 $\|w\|_{\mathcal{H}}^2 \quad \text{penalized} : KALE \text{ smoothie}$
 $KALE(\colongledown), P; \mathcal{H}) = 0.12$



Glaser, Arbel, G. "KALE Flow: A Relaxed KL Gradient Flow for Probabilities with Disjoint Support," (arXiv, 2021, Section 2)

The KALE smoothie and "mode collapse"

■ Two Gaussians with same means, different variance



Topological properties of KALE (1)

Key requirements on \mathcal{H} and \mathcal{X} :

- Compact domain \mathcal{X} ,
- \mathcal{H} dense in the space $C(\mathcal{X})$ of continuous functions on \mathcal{X} wrt $\|\cdot\|_{\infty}$.
- If $f \in \mathcal{H}$ then $-f \in \mathcal{H}$ and $cf \in \mathcal{H}$ for $0 \le c \le C_{\max}$.

```
Theorem: KALE(P, Q; \mathcal{H}) \geq 0 and KALE(P, Q; \mathcal{H}) = 0 iff P = Q.
```

Zhang, Liu, Zhou, Xu, and He. "On the Discrimination-Generalization Tradeoff in GANs" (ICLR 2018, Corollary 2.4; Theorem B.1)
Arbel, Liang, G. (ICLR 2021, Proposition 1)

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Theorem:
$$KALE(P, Q; \mathcal{H}) \geq 0$$
 and $KALE(P, Q; \mathcal{H}) = 0$ iff $P = Q$.

 \mathcal{H} dense in $C(\mathcal{X})$ for $\mathcal{X} \subset \mathbb{R}^d$ when:

$$\mathcal{H} = \operatorname{span}\{\sigma(w \top x + b) : [w, b] \in \Theta\}$$

$$\sigma(u) = \max\{u,0\}^{\alpha}, \ \alpha \in \mathbb{N}, \ \mathrm{and} \ \{\lambda \theta : \lambda \geq 0, \theta \in \Theta\} = \mathbb{R}^{d+1}.$$

Zhang, Liu, Zhou, Xu, and He. "On the Discrimination-Generalization Tradeoff in GANs" (ICLR 2018, Corollary 2.4; Theorem B.1)
Arbel, Liang, G. (ICLR 2021, Proposition 1)

Topological properties of KALE (2)

Additional requirement: all functions in ${\mathcal H}$ Lipschitz in their inputs with constant L

Theorem: $KALE(P, \mathbb{Q}^n; \mathcal{H}) \to 0$ iff $\mathbb{Q}^n \to P$ under the weak topology.

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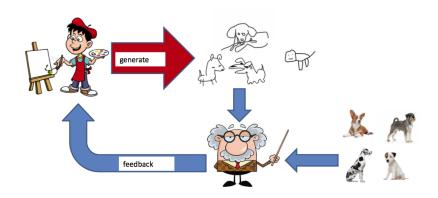
Partial proof idea:

$$egin{aligned} KALE(P, oldsymbol{Q}; \mathcal{H}) &= -\int f dP - \int \exp(-f) doldsymbol{Q} + 1 \ &= \int f(x) doldsymbol{Q}(x) - f(x') dP(x') \ &- \int \underbrace{\left(\exp(-f) + f - 1\right)}_{\geq 0} doldsymbol{Q} \ &\leq \int f(x) doldsymbol{Q}(x) - f(x') dP(x') \leq LW_1(P, oldsymbol{Q}) \end{aligned}$$

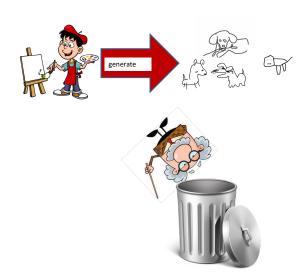
Liu, Bousquet, Chaudhuri. "Approximation and Convergence Properties of Generative Adversarial Learning" (NeurIPS 2017); Arbel, Liang, G. (ICLR 2021, Proposition 1)

How to train your GAN Generalized Energy-Based Model

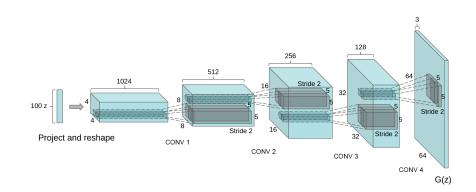
Visual notation: GAN setting



Visual notation: GAN setting



Reminder: the generator



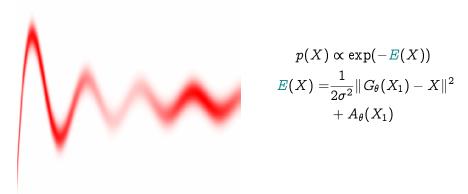
Radford, Metz, Chintala, ICLR 2016

Target distribution P

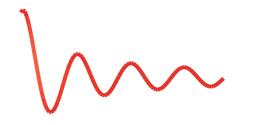


$$egin{aligned} z &\sim \mathit{Unif}\left[0,1
ight] \ & \widetilde{z} = au(z) \ & X = G_{ heta^\star}(\widetilde{z}), \quad X_1 = \widetilde{z} \end{aligned}$$

EBM approximation to target:



GAN (generator) distribution Q_{θ}



$$egin{aligned} & ext{Generator} \ z \sim unif[0,1] \ X = egin{aligned} & E_{m{ heta}}(z) \end{aligned} \end{aligned} egin{aligned} & ext{Critic} \ & ext{$MLP(X)$} \end{aligned}$$

Mass of GEBM corrected by critic



Generator

$$z \sim unif[0,1] \ X = {\color{red}B_{ heta}(z)}$$

Re-weight using importance weights defined by energy:

$$w(x) \propto \exp(-E(x))$$

Generalized energy-based models

Define a model $Q_{B_{\theta},E}$ as follows:

■ Sample from generator with parameters θ

$$X \sim Q_{\theta} \quad \iff \quad X = B_{\theta}(Z), \quad Z \sim \eta$$

■ Reweight the samples according to importance weights:

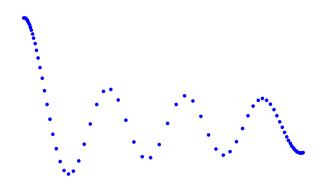
$$f_{oldsymbol{Q},E}(x) = rac{\exp(-E(x))}{Z_{oldsymbol{Q}_{oldsymbol{ heta},E}}}, \qquad Z_{oldsymbol{Q},E} = \int \exp(-E(x)) d rac{oldsymbol{Q}_{oldsymbol{ heta}}(x),}{2}$$

where $E \in \mathcal{E}$, the energy function class.

$$f_{Q,E}(x)$$
 is Radon-Nikodym derivative of $Q_{B_{\theta},E}$ wrt Q_{θ} .

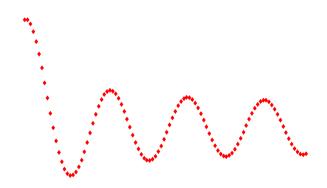
■ When Q_{θ} has density wrt Lebesgue on \mathcal{X} , this is a standard energy-based model.

Target distribution P



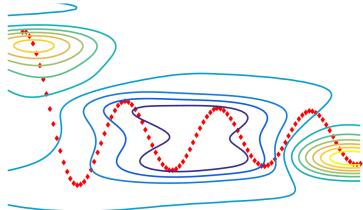
Example thanks to M. Arbel

GAN (generator) Q_{θ} , correct support but wrong mass



Example thanks to M. Arbel

Log energy function and Q_{θ}

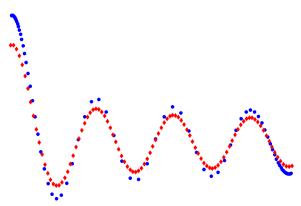


Key:

■ Orange: increase mass

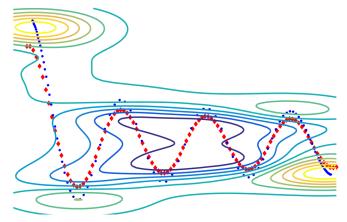
■ Blue: reduce mass

Target distribution P and GAN (generator) Q_{θ} , wrong support and wrong mass



Example thanks to M. Arbel

Log energy function, P, and Q_{θ}



Key:

- Orange: increase weight
- Blue: reduce weight

How do we learn the energy E?

How do we learn the energy E?

Fit the model using Generalized Log-Likelihood:

$$\mathcal{L}_{P,oldsymbol{Q}}(E) := \int \log(f_{oldsymbol{Q},E}) dP = - \int E \, dP - \log Z_{oldsymbol{Q},E}$$

- When $KL(P, \mathbb{Q}_{\theta})$ well defined, above is Donsker-Varadhan lower bound on KL
 - tight when $E(z) = -\log(p(z)/q(z))$.
- However, Generalized Log-Likelihood still defined when P and Q_{θ} mutually singular (as long as E smooth)!

Fit the model using Generalized Log-Likelihood:

$$\mathcal{L}_{P,\,oldsymbol{Q}}(E) := \int \log(f_{oldsymbol{Q},E}) dP = -\int E\, dP - \log\int \exp(-E) d\, oldsymbol{Q}_{oldsymbol{ heta}}$$

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One last trick...(convexity of exponential)

$$-\log\int\exp(-E)dQ_{ heta}\geq -c-e^{-c}\int\exp(-E)dQ_{ heta}+1$$

tight whenever $c = \log \int \exp(-E) dQ_{\theta}$.

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tight whenever $c = \log \int \exp(-E) dQ_{\theta}$.

Generalized Log-Likelihood has the lower bound:

$$egin{aligned} \mathcal{L}_{P,oldsymbol{Q}}(E) &\geq -\int (E+c)dP - \int \exp(-E-c)doldsymbol{Q}_{ heta} + 1 \ &:= \mathcal{F}(P,oldsymbol{Q}_{ heta};\mathcal{E}+\mathbb{R}) \end{aligned}$$

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This is the KALE! with function class $\mathcal{E} + \mathbb{R}$.

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Jointly maximizing yields the maximum likelihood energy E^* and corresponding $c^* = \log \int \exp(-E) dQ_{\theta}$.

Training the base measure (generator)

Recall the generator:

$$X = B_{\theta}(Z), \quad Z \sim \eta$$

Define: $\mathcal{K}(\theta) := \mathcal{F}(P, Q_{\theta}; \mathcal{E} + \mathbb{R})$

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Recall the generator:

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Define: $\mathcal{K}(\theta) := \mathcal{F}(P, Q_{\theta}; \mathcal{E} + \mathbb{R})$

Theorem: \mathcal{K} is lipschitz and differentiable for almost all $\theta \in \Theta$ with:

$$\nabla \mathcal{K}(\boldsymbol{\theta}) = Z_{\boldsymbol{Q},E^*}^{-1} \int \nabla_x E^*(\underline{B}_{\boldsymbol{\theta}}(z)) \nabla_{\boldsymbol{\theta}} \underline{B}_{\boldsymbol{\theta}}(z) \exp(-E^*(\underline{B}_{\boldsymbol{\theta}}(z))) \eta(z) dz.$$

where E^* achieves supremum in $\mathcal{F}(P, \mathbb{Q}; \mathcal{E} + \mathbb{R})$.

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where E^* achieves supremum in $\mathcal{F}(P, Q; \mathcal{E} + \mathbb{R})$.

Assumptions:

- Functions in \mathcal{E} parametrized by $\psi \in \Psi$, where Ψ compact,
 - jointly continous w.r.t. (ψ, x) , L-lipschitz and L-smooth w.r.t. x.
- $(\theta, z) \mapsto B_{\theta}(z)$ jointly continuous wrt (θ, z) , $z \mapsto B_{\theta}(z)$ uniformly Lipschitz w.r.t. z, lipschitz and smooth wrt θ (see paper: constants depend on z)

Sampling from the model

Consider end-to-end model $Q_{B_{\theta},E}$, where recall that

$$X = \mathcal{B}_{\theta}(Z), \quad Z \sim \eta,$$

$$f_{\mathcal{B},E}(x) := rac{\exp(-E(x))}{Z_{oldsymbol{Q},E}}$$

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$$X = \mathcal{B}_{\theta}(Z), \quad Z \sim \eta,$$

$$f_{B,E}(x) := rac{\exp(-E(x))}{Z_{oldsymbol{Q},E}}$$

For a test function g,

$$\int g(x)dQ_{B,E}(x) = \int g(B(z))f_{B,E}(B(z))\eta(z)dz$$

Posterior latent distribution therefore

$$\nu_{B,E}(z) = \eta(z) f_{B,E}(B(z))$$

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Posterior latent distribution therefore

$$u_{B,E}(z) = \eta(z) f_{B,E}(B(z))$$

Sample $z \sim \nu_{B,E}$ via Langevin diffusion-derived algorithms (MALA, ULA, HMC,...) to exploit gradient information.

Generate new samples in X via

$$X \sim Q_{B,E} \iff Z \sim \nu_{B,E}, \quad X = B_{\theta}(Z).$$

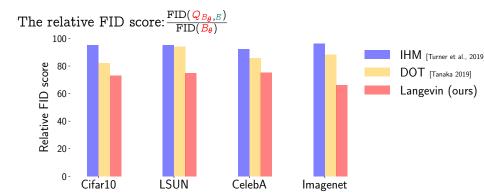
Experiments

Examples: sampling at modes

Tempered GEBM Cifar10 samples at different stages of sampling using a Kinetic Langevin Algorithm (KLA). Early samples \rightarrow late samples. Model run at *low temperature* ($\beta = 100$) for better quality samples.



Sampling at modes: results



For a given generator B_{θ} and energy E, samples always better (FID score) than generator alone.

Examples: moving between modes

Tempered GEBM Cifar10 samples at different stages of sampling using KLA. Early samples \rightarrow late samples.

Model run at *lower friction* (but still low temperature, $\beta = 100$) for mode exploration.



Summary

- Generalized energy based model:
 - End-to-end model incorporating generator and critic
 - Always better samples than generator alone.
- ICLR 2021

https://github.com/MichaelArbel/GeneralizedEBM

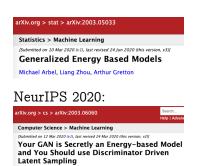


Summary

- Generalized energy based model:
 - End-to-end model incorporating generator and critic
 - Always better samples than generator alone.

■ ICLR, 2021

https://github.com/MichaelArbel/GeneralizedEBM



Tong Che, Ruixiang Zhang, Jascha Sohl-Dickstein, Hugo Larochelle,

Liam Paull, Yuan Cao, Yoshua Bengio

arXiv: org > cs > arXiv:2012.00780 Computer Science > Machine Learning Bubberniate ni Da 2009. liju strende Jan 2021 bits version, viii) Refining Deep Generative Models via Discriminator Gradient Flow Abdul Fatir Ansari, Ming Liang Ang, Harold Soh ICLR 2021: arXiv:org > cs > arXiv:2010.00654 Loss of the Computer Science > Machine Learning [Indemned on 1 0x 2009 bit, last revised 9 (vii) 2021 (bits version, viii) VAEBM: A Symbiosis between Variational Autoencoders and Energy-based Models Zhibheno Xia, Astraet Kreis, land Autur. Arash Valdat

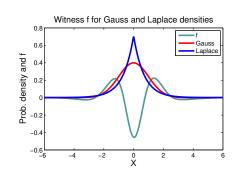
ICLR 2021:

How to find the best kernel for MMD

Integral prob. metric vs feature difference

The MMD:

$$egin{aligned} & MMD(P, \column{Q}; F) \ &= \sup_{\|f\|_{\mathcal{F}} \leq 1} \left[\operatorname{E}_P f(X) - \operatorname{E}_{\column{Q}} f(\column{Y})
ight] \ &= \left\| egin{aligned} & \mu_P - eta_{\column{Q}}
ight\|_{\mathcal{F}} \end{aligned}$$



The maximum mean discrepancy in terms of kernel means:

$$\begin{split} MMD^{2}(P,Q) &= \|\mu_{P} - \mu_{Q}\|_{\mathcal{F}}^{2} \\ &= \langle \mu_{P} - \mu_{Q}, \mu_{P} - \mu_{Q} \rangle_{\mathcal{F}} \\ &= \langle \mu_{P}, \mu_{P} \rangle_{\mathcal{F}} + \langle \mu_{Q}, \mu_{Q} \rangle_{\mathcal{F}} - 2 \langle \mu_{P}, \mu_{Q} \rangle_{\mathcal{F}} \\ &= \underbrace{\mathbb{E}_{P}k(X,X')}_{\text{(a)}} + \underbrace{\mathbb{E}_{Q}k(Y,Y')}_{\text{(a)}} - 2\underbrace{\mathbb{E}_{P,Q}k(X,Y)}_{\text{(b)}} \end{split}$$

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(a)= within distrib. similarity, (b)= cross-distrib. similarity.

The maximum mean discrepancy in terms of kernel means:

$$\begin{split} MMD^{2}(P,Q) &= \left\| \mu_{P} - \mu_{Q} \right\|_{\mathcal{F}}^{2} \\ &= \left\langle \mu_{P} - \mu_{Q}, \mu_{P} - \mu_{Q} \right\rangle_{\mathcal{F}} \\ &= \left\langle \mu_{P}, \mu_{P} \right\rangle_{\mathcal{F}} + \left\langle \mu_{Q}, \mu_{Q} \right\rangle_{\mathcal{F}} - 2 \left\langle \mu_{P}, \mu_{Q} \right\rangle_{\mathcal{F}} \\ &= \underbrace{\mathbb{E}_{P}k(X,X')}_{\text{(a)}} + \underbrace{\mathbb{E}_{Q}k(Y,Y')}_{\text{(a)}} - 2\underbrace{\mathbb{E}_{P,Q}k(X,Y)}_{\text{(b)}} \end{split}$$

Illustration of MMD

- Dogs (= P) and fish (= Q) example revisited
- Each entry is one of $k(dog_i, dog_j)$, $k(dog_i, fish_j)$, or $k(fish_i, fish_j)$

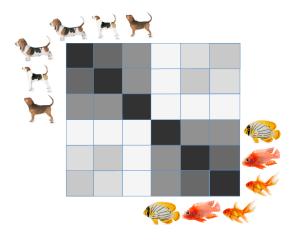


Illustration of MMD

The maximum mean discrepancy:

$$egin{aligned} \widehat{MMD}^2 = & rac{1}{n(n-1)} \sum_{i
eq j} k(\operatorname{dog}_i, \operatorname{dog}_j) + rac{1}{n(n-1)} \sum_{i
eq j} k(\operatorname{fish}_i, \operatorname{fish}_j) \\ & - rac{2}{n^2} \sum_{i,j} k(\operatorname{dog}_i, \operatorname{fish}_j) \\ & k(\operatorname{dog}_i, \operatorname{dog}_j) \quad k(\operatorname{dog}_i, \operatorname{fish}_j) \end{aligned}$$

A statistical test using MMD

The empirical MMD:

$$egin{aligned} \widehat{MMD}^2 = & rac{1}{n(n-1)} \sum_{i
eq j} k(\pmb{x_i}, \pmb{x_j}) + rac{1}{n(n-1)} \sum_{i
eq j} k(\pmb{ extbf{y}}_i, \pmb{ extbf{y}}_j) \ & - rac{2}{n^2} \sum_{i,j} k(\pmb{x_i}, \pmb{ extbf{y}}_j) \end{aligned}$$

How does this help decide whether P = Q?

A statistical test using MMD

The empirical MMD:

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Perspective from statistical hypothesis testing:

- Null hypothesis \mathcal{H}_0 when P = Q
 - should see \widehat{MMD}^2 "close to zero".
- Alternative hypothesis \mathcal{H}_1 when $P \neq Q$
 - should see \widehat{MMD}^2 "far from zero"

A statistical test using MMD

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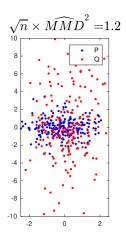
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 - should see \widehat{MMD}^2 "close to zero".
- Alternative hypothesis \mathcal{H}_1 when $P \neq Q$
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Want Threshold c_{α} for \widehat{MMD}^2 to get false positive rate α

Behaviour of \widehat{MMD}^2 when $P \neq Q$

Draw n = 200 i.i.d samples from P and Q

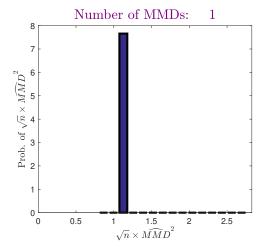
- Laplace with different y-variance.
- $\sqrt{n} \times \widehat{MMD}^2 = 1.2$

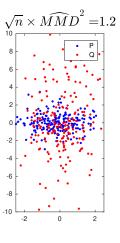


Draw n = 200 i.i.d samples from P and Q

■ Laplace with different y-variance.

$$\sqrt{n} \times \widehat{MMD}^2 = 1.2$$

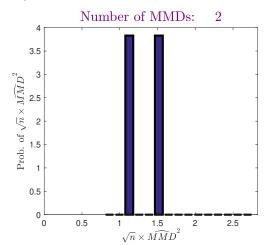


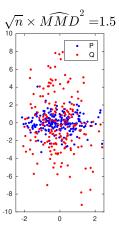


Draw n = 200 new samples from P and Q

■ Laplace with different y-variance.

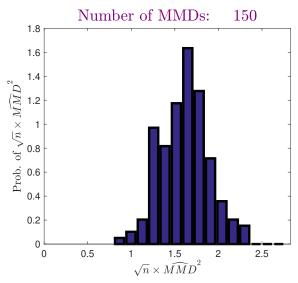
$$\sqrt{n} \times \widehat{MMD}^2 = 1.5$$



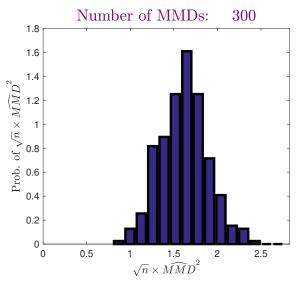


50/67

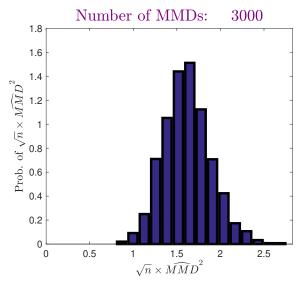
Repeat this 150 times ...



Repeat this 300 times ...



Repeat this 3000 times ...

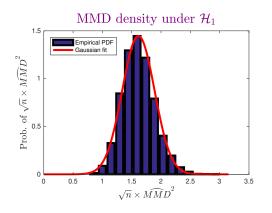


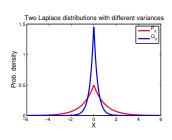
Asymptotics of \widehat{MMD}^2 when $P \neq Q$

When $P \neq Q$, statistic is asymptotically normal,

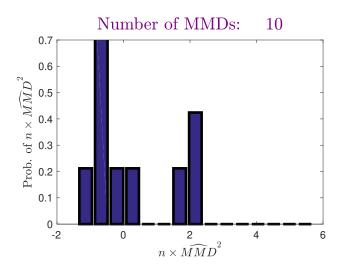
$$rac{\widehat{ ext{MMD}}^2 - ext{MMD}^2(extit{P}, extit{Q})}{\sqrt{V_n(extit{P}, extit{Q})}} \stackrel{D}{\longrightarrow} \mathcal{N}(0, 1),$$

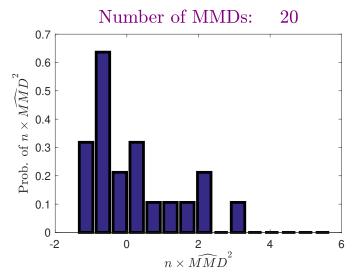
where variance $V_n(P,Q) = O(n^{-1})$.

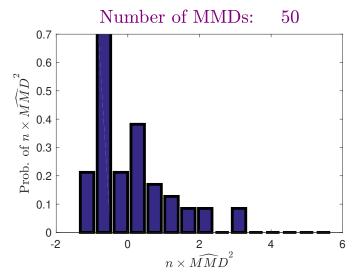


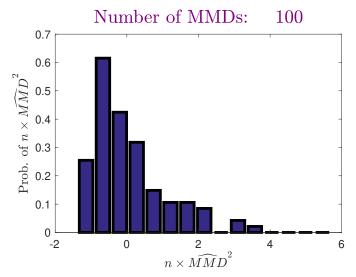


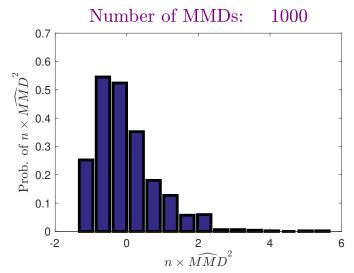
What happens when P and Q are the same?







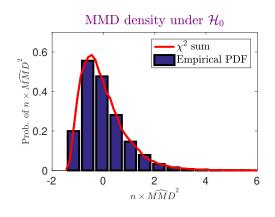




Asymptotics of \widehat{MMD}^2 when P = Q

Where P = Q, statistic has asymptotic distribution

$$n\widehat{ ext{MMD}}^2 \sim \sum_{l=1}^\infty \lambda_l \left[z_l^2 - 2
ight]$$

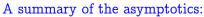


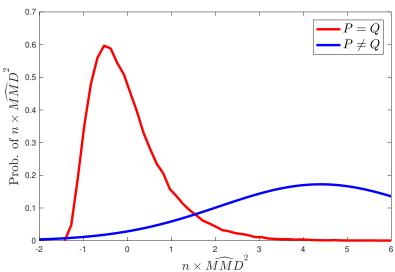
where

$$\lambda_i \psi_i(x') = \int_{\mathcal{X}} \underbrace{ ilde{k}(x,x')}_{ ext{centred}} \psi_i(x) dP(x)$$

$$z_l \sim \mathcal{N}(0, 2)$$
 i.i.d

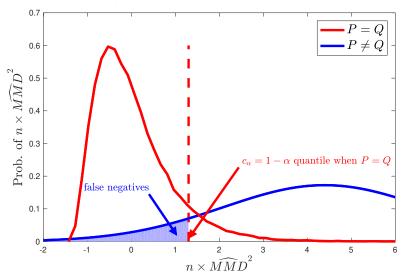
A statistical test





A statistical test

Test construction: (G., Borgwardt, Rasch, Schoelkopf, and Smola, JMLR 2012)

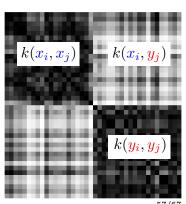


How do we get test threshold c_{α} ?

Original empirical MMD for dogs and fish:

$$X = \begin{bmatrix} & & & \\ & & &$$

$$egin{aligned} \widehat{MMD}^2 = & rac{1}{n(n-1)} \sum_{i
eq j} k(\pmb{x_i}, \pmb{x_j}) \ & + rac{1}{n(n-1)} \sum_{i
eq j} k(\pmb{y_i}, \pmb{y_j}) \ & - rac{2}{n^2} \sum_{i,j} k(\pmb{x_i}, \pmb{y_j}) \end{aligned}$$



How do we get test threshold c_{α} ?

Permuted dog and fish samples (merdogs):

$$\widetilde{X} = \begin{bmatrix} & & & \\ & & & \\ & & & \end{bmatrix}$$

$$\widetilde{Y} = [$$



How do we get test threshold c_{α} ?

Permuted dog and fish samples (merdogs):

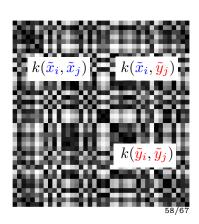
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eq i} k(ilde{m{x}}_i, ilde{m{y}}_j) \end{aligned}$$

Permutation simulates

$$P = Q$$



The best test for the job

- A test's power depends on k(x, x'), P, and Q (and n)
- With characteristic kernel, MMD test has power \rightarrow 1 as $n \rightarrow \infty$ for any (fixed) problem
 - But, for many P and Q, will have terrible power with reasonable n!

The best test for the job

- A test's power depends on k(x, x'), P, and Q (and n)
- With characteristic kernel, MMD test has power \rightarrow 1 as $n \rightarrow \infty$ for any (fixed) problem
 - But, for many P and Q, will have terrible power with reasonable n!
- You can choose a good kernel for a given problem
- You *can't* get one kernel that has good finite-sample power for all problems
 - No one test can have all that power

■ Simple choice: exponentiated quadratic

$$k(x,y) = \exp\left(-rac{1}{2\sigma^2}\|x-y\|^2
ight)$$

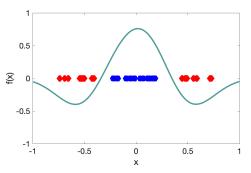
Characteristic: for any σ : for any P and Q, power $\to 1$ as $n \to \infty$

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- **The Example 2** Characteristic: for any σ : for any P and Q, power $\to 1$ as $n \to \infty$
- But choice of σ is very important for finite n...

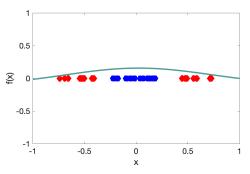
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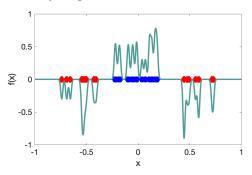
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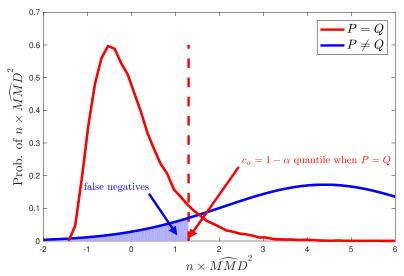


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- Characteristic: for any σ : for any P and Q, power $\to 1$ as $n \to \infty$
- But choice of σ is very important for finite n...
- lacksquare ... and some problems (e.g. images) might have no good choice for σ

Graphical illustration

Maximising test power same as minimizing false negatives



The power of our test (Pr₁ denotes probability under $P \neq Q$):

$$ext{Pr}_1\left(n\widehat{ ext{MMD}}^2>\hat{c}_lpha
ight)$$

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ight) \ & o \Phi \left(rac{ ext{MMD}^2(P,Q)}{\sqrt{V_n(P,Q)}} - rac{c_{lpha}}{n \sqrt{V_n(P,Q)}}
ight) \end{split}$$

where

- \blacksquare Φ is the CDF of the standard normal distribution.
- \hat{c}_{α} is an estimate of c_{α} test threshold.

The power of our test (Pr₁ denotes probability under $P \neq Q$):

$$ext{Pr}_1\left(n\widehat{ ext{MMD}}^2>\hat{c}_{lpha}
ight) \ o \Phi\left(\underbrace{rac{ ext{MMD}^2(P,Q)}{\sqrt{V_n(P,Q)}}}_{O(n^{1/2})} - \underbrace{rac{c_{lpha}}{n\sqrt{V_n(P,Q)}}}_{O(n^{-1/2})}
ight)$$

For large n, second term negligible!

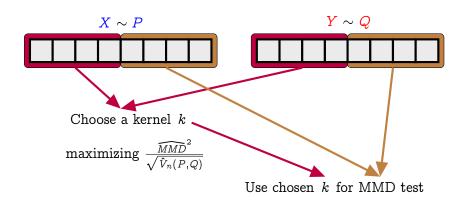
The power of our test (Pr₁ denotes probability under $P \neq Q$):

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ight) \end{split}$$

To maximize test power, maximize

$$\frac{\text{MMD}^2(P,Q)}{\sqrt{V_n(P,Q)}}$$

Data splitting

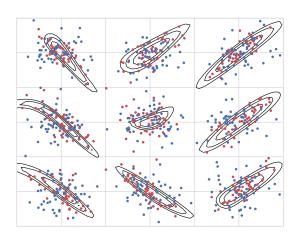


Learning a kernel helps a lot

Kernel with deep learned features:

$$k_{ heta}(x,y) = \left[(1-\epsilon) \kappa(\Phi_{ heta}(x),\Phi_{ heta}(y)) + \epsilon
ight] rac{oldsymbol{q}}{oldsymbol{q}}(x,y)$$

 κ and q are Gaussian kernels



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- κ and q are Gaussian kernels
- CIFAR-10 vs CIFAR-10.1, null rejected 75% of time



CIFAR-10 test set (Krizhevsky 2009)

$$X \sim P$$



CIFAR-10.1 (Recht+ ICML 2019)

$$Y \sim Q$$

Learning a kernel helps a lot

Kernel with deep learned features:

$$k_{\theta}(x, y) = [(1 - \epsilon)\kappa(\Phi_{\theta}(x), \Phi_{\theta}(y)) + \epsilon] q(x, y)$$

 κ and q are Gaussian kernels

■ CIFAR-10 vs CIFAR-10.1, null rejected 75% of time

arXiv.org > stat > arXiv:2002.09116

Statistics > Machine Learning

[Submitted on 21 Feb 2020]

Learning Deep Kernels for Non-Parametric Two-Sample Tests

Feng Liu, Wenkai Xu, Jie Lu, Guangquan Zhang, Arthur Gretton, D. J. Sutherland

ICML 2020

Questions?



Post-credit scene: MMD flow

From NeurIPS 2019:

Maximum Mean Discrepancy Gradient Flow

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Sanity check: reduction to EBM case

