Convergence rates of spectral methods for statistical inverse learning problems

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Rates for statistical inverse learning

DETERMINISTIC AND STATISTICAL INVERSE PROBLEMS

- ▶ Let A be a bounded operator between Hilbert spaces $\mathcal{H}_1 \to \mathcal{H}_2$ (assumed known)
- ▶ Classical (deterministic) inverse problem: observe

$$y^\sigma = Af^* + \sigma\eta, \tag{IP}$$

under the assumption $\|\eta\| \leq 1$.

- ▶ Note: the $\mathcal{H}_2$-norm measures the observation error; the $\mathcal{H}_1$-norm measures the reconstruction error.
- ▶ Classical deterministic theory: see Engl, Hanke and Neubauer (2000).

DETERMINISTIC AND STATISTICAL INVERSE PROBLEMS

- Inverse problem

$$y^\sigma = Af^* + \sigma\eta. \tag{IP}$$

- What if noise is random? Classical statistical inverse problem model: η is a Gaussian white noise process on $\mathcal{H}_2$.
- Note: in this case (IP) is not an equation between elements in $\mathcal{H}_2$, but is to be interpreted as process on $\mathcal{H}_2$.
- Under Hölder source condition of order r and polynomial ill-posedness (eigenvalue decay) of order $1/s$, sharp minimax rates are known in this setting:

$$\left\| (A^*A)^\theta (\widehat{f} - f^*) \right\|_{\mathcal{H}_1} \asymp O\left(\sigma^{\frac{2(r+\theta)}{2r+1+s}} \right) \asymp O\left(\sigma^{\frac{2(\nu+b\theta)}{2\nu+b+1}} \right),$$

for $\theta \in [0, \frac{1}{2}]$ ($\theta = 0$: inverse problem; $\theta = \frac{1}{2}$: direct problem.)
(Alternate parametrization: $b := 1/s$, $\nu := rb$ "intrinsic regularity".)

LINEAR SPECTRAL REGULARIZATION METHODS

- ▶ Inverse problem (deterministic or statistical) where A is known.
- ▶ First consider the so-called "normal equation":

$$A^* y^\sigma = (A^* A) f^* + \sigma(A^* \eta).$$

- ▶ Linear spectral methods: let $\zeta_\lambda(x) : \mathbb{R}_+ \to \mathbb{R}_+$ be a real function of 1 real variable which is an "approximation of $1/x$" and $\lambda > 0$ a tunig parameter.
- ▶ Define

$$\widehat{f}_\lambda = \zeta_\lambda(A^* A) A^* y^\sigma$$

- ▶ Examples: Tikhonov $\zeta_\lambda(x) = (x + \lambda)^{-1}$, spectral cut-off $\zeta_\lambda(x) = x^{-1} \mathbf{1}\{x \geq \lambda\}$, Landweber iteration polynomials, ν-methods . . .
- ▶ Under general conditions on ζ_λ, optimal/mimimax rates can be attained by such methods (Deterministic: Engl et al. , 2000; Stochastic noise: Bissantz et al, 2007)

STATISTICAL LEARNING

- "Learning" usually refers to the following setting:

$$(X_i, Y_i)_{i=1,\ldots,n} \quad \text{i.i.d.} \ \sim \mathbb{P}_{XY} \text{ on } \mathcal{X} \times \mathcal{Y}$$

  where $\mathcal{Y} \subset \mathbb{R}$,

- Goal: estimate some functional related to the dependency between X and Y,

- for instance (nonparametric) least squares regression: estimate

$$f^*(x) := \mathbb{E}\left[Y|X=x\right],$$

  and measure the quality of an estimator $\widehat{f}$ via

$$\left\| f^* - \widehat{f} \right\|_{L^2(\mathbb{P}_X)}^2 = \mathbb{E}_{X \sim \mathbb{P}_X}\left[\left(\widehat{f}(X) - f^*(X) \right)^2 \right]$$

SETTING: "INVERSE LEARNING" PROBLEM

- ▶ We refer to "inverse learning" for an inverse problem where we have noisy observations at random design points:

$$(X_i, Y_i)_{i=1,\dots,n} \text{ i.i.d. } : \qquad Y_i = (Af^*)(X_i) + \varepsilon_i . \qquad \text{(ILP)}$$

- ▶ the goal is to recover $f^* \in \mathcal{H}_1$.
- ▶ early works on closely related subjects: from the splines literature in the 80's (e.g. O'Sullivan '90)

MAIN ASSUMPTION FOR INVERSE LEARNING

Model: $Y_i = (Af^*)(X_i) + \varepsilon_i$, $i = 1, \ldots, n$, where $A : \mathcal{H}_1 \to \mathcal{H}_2$. (ILP)

Observe:

- $\mathcal{H}_2$ should be a space of real-values functions on $\mathcal{X}$.
- the geometrical structure of the "measurement errors" will be dictated by the statistical properties of the sampling scheme – we do not need to assume or consider any a priori Hilbert structure on $\mathcal{H}_2$
- the crucial stuctural assumption we make is the following:

Assumption

The family of evaluation functionals (S_x), $x \in \mathcal{X}$, defined by

$$S_x : \mathcal{H}_1 \longrightarrow \mathbb{R}$$
$$f \longmapsto (S_x)(f) := (Af)(x)$$

is uniformly bounded, i.e., there exists $\kappa < \infty$ such that for any $x \in \mathcal{X}$

$$|S_x(f)| \leq \kappa \|f\|_{\mathcal{H}_1} .$$

GEOMETRY OF INVERSE LEARNING

The inverse learning setting was essentially introduced by Caponnetto et al. (2006).

- ▶ Riesz's theorem implies the existence for any $x \in \mathcal{X}$ of $F_x \in \mathcal{H}_1$:

$$\forall f \in \mathcal{H}_1 : \qquad (Af)(x) = \langle f, F_x \rangle$$

- ▶ $K(x, y) := \langle F_x, F_y \rangle$ defines a positive semidefinite kernel on $\mathcal{X}$ with associated reproducing kernel Hilbert space (RKHS) denoted $\mathcal{H}_K$.
- ▶ as a pure function space, $\mathcal{H}_K$ coincides with $Im(A)$.
- ▶ assuming A injective, A is in fact an isometric isomorphism between $\mathcal{H}_1$ and $\mathcal{H}_K$.

GEOMETRY OF INVERSE LEARNING

- ▶ Main assumption implies that as a function space, *Im(A)* is endowed with a natural RKHS structure with a kernel *K* bounded by κ.
- ▶ Furthermore this RKHS $\mathcal{H}_K$ is isometric to $\mathcal{H}_1$ (through A^{-1}).
- ▶ Therefore, the inverse learning problem is formally equivalent to the kernel learning problem

$$Y_i = h^*(X_i) + \varepsilon_i, \qquad i = 1, \ldots, n$$

where $h^* \in \mathcal{H}_K$, and we measure the quality of an estimator $\widehat{h} \in \mathcal{H}_K$ via the RKHS norm $\left\|\widehat{h} - h^*\right\|_{\mathcal{H}_K}$

- ▶ Indeed, if we put $\widehat{f} := A^{-1}\widehat{h}$, then

$$\left\|\widehat{f} - f^*\right\|_{\mathcal{H}_1} = \left\|A(\widehat{f} - f^*)\right\|_{\mathcal{H}_K} = \left\|\widehat{h} - h^*\right\|_{\mathcal{H}_K}$$

SETTING, REFORMULATED

- ► We are actually back to the familiar regression setting on a random design,

$$Y_i = h^*(X_i) + \varepsilon_i \,,$$

where $(X_i, Y_i)_{1 \leq i \leq n}$ is an i.i.d. sample from $\mathbb{P}_{XY}$ on the space $\mathcal{X} \times \mathbb{R}$,

- ► with $\mathbb{E}[\varepsilon_i | X_i] = 0$.
- ► Noise assumptions:

(BernsteinNoise) $\quad \mathbb{E}\left[\varepsilon_i^p | X_i\right] \leq \frac{1}{2} p! M^p, \quad p \geq 2$

- ► h^* is assumed to lie in a (known) RKHS $\mathcal{H}_K$ with bounded kernel K.
- ► The criterion for measuring the quality of an estimator $\widehat{h}$ is the RKHS norm

$$\left\| \widehat{h} - h^* \right\|_{\mathcal{H}_K} \,.$$

EMPIRICAL AND POPULATION OPERATORS

▶ Define the (random) empirical evaluation operator

$$T_n : h \in \mathcal{H} \mapsto (h(X_1), \ldots, h(X_n)) \in \mathbb{R}^n$$

and its population counterpart the inclusion operator

$$T : h \in \mathcal{H} \mapsto h \in L_2(\mathcal{X}, \mathbb{P}_X);$$

▶ the (random) empirical kernel integral operator

$$T_n^* : (v_1, \ldots, v_n) \in \mathbb{R}^n \mapsto \frac{1}{n} \sum_{i=1}^{n} K(X_i, .) v_i \in \mathcal{H}$$

and its population counterpart, the kernel integral operator

$$T^* : f \in L_2(\mathcal{X}, \mathbb{P}_X) \mapsto T^*(f) = \int f(x) k(x, .) d\mathbb{P}_X(x) \in \mathcal{H}.$$

▶ finally, define the empirical covariance operator $S_n = T_n^* T_n$ and its population counterpart $S = T^* T$.
▶ observe that S_n, S are both opertors $\mathcal{H}_K \to \mathcal{H}_K$; the intuition is that S_n is a (random) approximation of S.

- Recall the model with $h^* \in \mathcal{H}_K$:

$$Y_i = h^*(X_i) + \varepsilon_i \qquad \text{i.e.} \qquad \mathbf{Y} = T_n h^* + \varepsilon,$$

  where $\mathbf{Y} := (Y_1, \ldots, Y_n)$.

- Associated "normal equation":

$$Z = T_n^* \mathbf{Y} = T_n^* T_n h^* + T_n^* \varepsilon = S_n h^* + T_n^* \varepsilon$$

- Idea (Rosasco, Caponnetto, De Vito, Odone): use methods from inverse problems literature
- Observe that there is also an error on the operator
- Use concentration principles to bound $\|T_n^* \varepsilon\|$ and $\|S_n - S\|$.

LINEAR SPECTRAL REGULARIZATION METHODS

▶ Linear spectral methods:

$$\widehat{h}_\zeta = \zeta(S_n)Z$$

for somme well-chosen function $\zeta : \mathbb{R} \to \mathbb{R}$ acting on the spectrum and "approximating" the function $x \mapsto x^{-1}$.

▶ Examples: Tikhonov $\zeta_\lambda(t) = (t + \lambda)^{-1}$, spectral cut-off $\zeta_\lambda(t) = t^{-1}\mathbf{1}\{t \geq \lambda\}$, Landweber iteration polynomials, ν-methods ...

SPECTRAL REGULARIZATION IN KERNEL SPACE

- Linear spectral regularization in kernel space is written

$$\widehat{h}_\zeta = \zeta(S_n) T_n^* \mathbf{Y}$$

- notice

$$\zeta(S_n) T_n^* = \zeta(T_n^* T_n) T_n^* = T_n^* \zeta(T_n T_n^*) = T_n^* \zeta(K_n),$$

where $K_n = T_n T_n^* : \mathbb{R}^n \to \mathbb{R}^n$ is the kernel Gram matrix,

$$K_n(i,j) = \frac{1}{n} K(X_i, X_j).$$

- equivalently:

$$\widehat{h}_\zeta = \sum_{i=1}^n \alpha_{\zeta,i} K(X_i, .)$$

with

$$\alpha_\zeta = \frac{1}{n} \zeta\left(\frac{1}{n} K_n\right) \mathbf{Y}.$$

STRUCTURAL ASSUMPTIONS

- Two parameters determine attainable convergence rates:
- (Hölder) Source condition for the signal: for $r > 0$, define

$$\mathbf{SC}(r, R): \quad h^* = S^r h_0 \text{ with } \|h_o\| \leq R$$

(can be generalized to "extended source conditions", see e.g. Mathé and Pereverzev 2003)

- Ill-posedness: if $(\lambda_i)_{i \geq 1}$ is the sequence of positive eigenvalues of S in nonincreasing order, then define

$$\mathbf{IP}^+(s, \beta): \quad \lambda_i \leq \beta i^{-\frac{1}{s}}$$

and

$$\mathbf{IP}^-(s, \beta'): \quad \lambda_i \geq \beta' i^{-\frac{1}{s}}$$

ERROR/RISK MEASURE

▶ We are measuring the error (risk) of an estimator $\widehat{h}$ in the family of norms

$$\left\| S^\theta(\widehat{h} - h^*) \right\|_{\mathcal{H}_K} \qquad (\theta \in [0, \frac{1}{2}])$$

▶ Note $\theta = 0$: inverse problem; $\theta = 1/2$: direct problem, since

$$\left\| S^{\frac{1}{2}}(\widehat{h} - h^*) \right\|_{\mathcal{H}_K} = \left\| \widehat{h} - h^* \right\|_{L^2(\mathbb{P}_X)}.$$

[1]: Smale and Zhou (2007)
[2]: Bauer, Pereverzev, Rosasco (2007)
[3]: Caponnetto, De Vito (2007)
[4]: Caponnetto (2006)

Error	[1]	[2]	[3]	[4]
$\left\| \widehat{h} - h^* \right\|_{L^2(\mathbb{P}_X)}$	$\left(\frac{1}{\sqrt{n}}\right)^{\frac{2r+1}{2r+2}}$	$\left(\frac{1}{\sqrt{n}}\right)^{\frac{2r+1}{2r+2}}$	$\left(\frac{1}{\sqrt{n}}\right)^{\frac{(2r+1)}{2r+1+s}}$	$\left(\frac{1}{\sqrt{n}}\right)^{\frac{(2r+1)}{2r+1+s}}$
$\left\| \widehat{h} - h^* \right\|_{\mathcal{H}_K}$	$\left(\frac{1}{\sqrt{n}}\right)^{\frac{r}{r+1}}$	$\left(\frac{1}{\sqrt{n}}\right)^{\frac{r}{r+1}}$		
Assumptions (q: qualification)	$r \leq \frac{1}{2}$	$r \leq q - \frac{1}{2}$	$r \leq \frac{1}{2}$	$0 \leq r \leq q - \frac{1}{2}$ +unlabeled data if $2r + s < 1$
Method	Tikhonov	General	Tikhonov	General

Matching lower bound: only for $\left\| \widehat{h} - h^* \right\|_{L^2(\mathbb{P}_X)}$ [2].

Compare to results known for regularization methods under Gaussian White Noise model: Mair and Ruymgaart (1996), Nussbaum and Pereverzev (1999), Bissantz, Hohage, Munk and Ruymgaart (2007).

ASSUMPTIONS ON REGULARIZATION FUNCTION

From now on we assume $\kappa = 1$ for simplicity. Standard assmptions on the regularization family $\zeta_\lambda : [0, 1] \to \mathbb{R}$ are:

(i) There exists a constant $D < \infty$ such that

$$\sup_{0 < \lambda \leq 1} \sup_{0 < t \leq 1} |t\zeta_\lambda(t)| \leq D,$$

(ii) There exists a constant $M' < \infty$ such that

$$\sup_{0 < \lambda \leq 1} \sup_{0 < t \leq 1} \lambda |\zeta_\lambda(t)| \leq \frac{M'}{,}$$

(iii) *Qualification:*

$$\forall \lambda \leq 1 : \qquad \sup_{0 < t \leq 1} |1 - t\zeta_\lambda(t)| \, t^\nu \leq \gamma_\nu \lambda^\nu.$$

holds for $\nu = 0$ and $\nu = q > 0$.

UPPER BOUND ON RATES

Theorem (Mücke, Blanchard)

Assume r, R, b, β are fixed positive constants and let $\mathcal{P}(r, R, s, \beta)$ denote the set of distributions on $\mathcal{X} \times \mathcal{Y}$ satisfying **(IP$^+$)**(s, β), **(SC)**(r, R) *and* **(BernsteinNoise)**. *Define*

$$\widehat{h}_{\lambda_n}^{(n)} = \zeta_{\lambda_n}(S_n)Z^{(n)}$$

using a regularization family (ζ_λ) satisfying the standard assumptions with qualification $q \geq r + \theta$, and the parameter choice rule

$$\lambda_n = \left(\frac{R^2\sigma^2}{n} \right)^{-\frac{1}{2r+1+s}}.$$

it holds for any $\theta \in [0, \frac{1}{2}], \eta \in (0, 1)$:

$$\sup_{P \in \mathcal{P}(r,R,s,\beta)} P^{\otimes n} \left(\left\| S^\theta(h^* - \widehat{h}_{\lambda_n}^{(n)}) \right\|_{\mathcal{H}_K} > C(\log \eta^{-1})R \left(\frac{\sigma^2}{R^2 n} \right)^{-\frac{(r+\theta)}{2r+1+s}} \right) \leq \eta.$$

COMMENTS

- it follows that the convergence rate obtained is of order

$$C.R \left(\frac{\sigma^2}{R^2 n} \right)^{-\frac{(r+\theta)}{2r+1+s}}$$

- the "constant" C depends on the various parameters entering in the assumptions, but **not** on n, R, σ!
- the result applies to all linear spectral regularization methods but assuming a precise tuning of the regularization constant λ as a function of the assumed regularization parameters of the target – not adaptive.

"WEAK" LOWER BOUND ON RATES

Theorem (Mücke, Blanchard)

Assume r, R, s, β *are fixed positive constants and let* $\mathcal{P}'(r, R, s, \beta)$ *denote the set of distributions on* $\mathcal{X} \times \mathcal{Y}$ *satisfying* **(IP$^-$)**(s, β), **(SC)**(r, R) *and* **(BernsteinNoise)**. *(We assume this set to be non empty!) Then*

$$\limsup_{n \to \infty} \inf_{\widehat{h}} \sup_{P \in \mathcal{P}'(r, R, s, \beta)} P^{\otimes n} \left(\left\| S^\theta (h^* - \widehat{h}) \right\|_{\mathcal{H}_K} > CR \left(\frac{\sigma^2}{R^2 n} \right)^{-\frac{(r+\theta)}{2r+1+s}} \right) > 0$$

Proof: Fano's lemma technique

"STRONG" LOWER BOUND ON RATES

Assume additionally "no big jumps in eigenvalues":

$$\inf_{k \geq 1} \frac{\lambda_{2k}}{\lambda_k} > 0$$

Theorem (Mücke, Blanchard)

Assume r, R, s, β are fixed positive constants and let $\mathcal{P}'(r, R, s, \beta)$ denote the set of distributions on $\mathcal{X} \times \mathcal{Y}$ satisfying **(IP$^-$)**(s, β), **(SC)**(r, R) *and* **(BernsteinNoise)**. *(We assume this set to be non empty!) Then*

$$\liminf_{n \to \infty} \inf_{\widehat{h}} \sup_{P \in \mathcal{P}'(r,R,s,\beta)} P^{\otimes n} \left(\left\| S^{\theta}(h^* - \widehat{h}) \right\|_{\mathcal{H}_K} > CR \left(\frac{\sigma^2}{R^2 n} \right)^{-\frac{(r+\theta)}{2r+1+s}} \right) > 0$$

Proof: Fano's lemma technique

COMMENTS

- obtained rates are minimax (but not adaptive) in the parameters $R, n, \sigma \ldots$
- ... provided **(IP$^-$)**$(s, \beta) \cap$ **(IP$^+$)**(s, α) is not empty.

STATISTICAL ERROR CONTROL

Error controls were introduced and used by Caponnetto and De Vito (2007), Caponnetto (2007), using Bernstein's inequality for Hilbert space-valued variables (see Pinelis and Sakhanenko; Yurinski).

Theorem (Caponetto, De Vito)

Define

$$\mathcal{N}(\lambda) = \mathrm{Tr}(\,(S + \lambda)^{-1}S\,) \, ,$$

then under assumption **(BernsteinNoise)** *we have the following:*

$$\mathbb{P}\left[\left\|(S+\lambda)^{-\frac{1}{2}}(T_n^* \mathbf{Y} - S_n h^*)\right\| \le 2M\left(\sqrt{\frac{\mathcal{N}(\lambda)}{n}} + \frac{2}{\sqrt{\lambda}n}\right)\log\frac{6}{\delta}\right] \ge 1 - \delta \, .$$

Also, the following holds:

$$\mathbb{P}\left[\left\|(S+\lambda)^{-\frac{1}{2}}(S_n - S)\right\|_{HS} \le 2\left(\sqrt{\frac{\mathcal{N}(\lambda)}{n}} + \frac{2}{\sqrt{\lambda}n}\right)\log\frac{6}{\delta}\right] \ge 1 - \delta \, .$$

PARTIAL LEAST SQUARES REGULARIZATION

Consider first the classical linear regression setting

$$\mathbf{Y} = \mathbf{X}\omega + \varepsilon \,,$$

where $\mathbf{Y} := (Y_1, \ldots, Y_n)$; $\mathbf{X} := (X_1, \ldots, X_n)^t$; $\varepsilon = (\varepsilon_1, \ldots, \varepsilon_n)$.

- ▶ Algorithmic description of Partial Least Squares:
- ▶ find direction v_1 s.t.

$$v_1 = \underset{v \in \mathbb{R}^d}{\operatorname{Arg\,Max}} \frac{\widehat{\operatorname{Cov}}(\langle X, v \rangle, Y)}{\|v\|} = \underset{v \in \mathbb{R}^d}{\operatorname{Arg\,Max}} \frac{\mathbf{Y}^t \mathbf{X} v}{\|v\|} \propto \mathbf{X}^t \mathbf{Y}$$

- ▶ project $\mathbf{Y}$ orthogonally on $\mathbf{X}v$ yielding $\mathbf{Y}_1$
- ▶ iterate the procedure on the residual $\mathbf{Y} - \mathbf{Y}_1$
- ▶ The fit at step m is $\sum_{i=1}^m \mathbf{Y}_i$.
- ▶ Regularization is obtained by early stopping.

PLS AND CONJUGATE GRADIENT

▶ An equivalent definition of PLS:

$$\omega_m = \underset{\omega \in \mathcal{K}_m(\mathbf{X}\mathbf{X}^t, \mathbf{X}^t\mathbf{Y})}{\text{Arg Min}} \|\mathbf{Y} - \mathbf{X}\omega\|^2$$

where

$$\mathcal{K}_m(A, z) = \text{span}\left\{z, Az, \ldots, A^{m-1}z\right\}$$

is a Krylov space of order m.

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- This definition is equivalent to m steps of the conjugate gradient algorithm applied to iteratively solve the linear equation

$$A\omega = \mathbf{X}^t\mathbf{X}\omega = \mathbf{X}^t\mathbf{Y} = z$$

- For any fixed m, the fit $\mathbf{Y}_m = \mathbf{X}\omega_m$ is a nonlinear function of $\mathbf{Y}$.

PROPERTIES OF CONJUGATE GRADIENT

▶ by definition ω_m has the form

$$\omega_m = p_m(A)z = \mathbf{X}^t p_m(\mathbf{X}\mathbf{X}^t)\mathbf{Y},$$

where p_m is a polynomial of degree $\leq m-1$.

▶ of particular interest are the residual polynomials

$$r_m(t) = 1 - tp_m(t); \qquad \|\mathbf{Y} - \mathbf{Y}_m\| = \left\| r_m(\mathbf{X}\mathbf{X}^t)\mathbf{Y} \right\|$$

▶ the polynomials r_m form a family of **orthogonal polynomials** for the inner product

$$\langle p, q \rangle = \left\langle p(\mathbf{X}\mathbf{X}^t)\mathbf{Y}, \mathbf{X}\mathbf{X}^t q(\mathbf{X}\mathbf{X}^t)\mathbf{Y} \right\rangle$$

and with the normalization $r_m(0) = 1$.

▶ the polynomials r_m follow an order 2 recurrence relation of the type

$$r_{m+1}(t) = a_m t r_m(t) + b_m r_m(t) + c_m r_{m-1}(t)$$

($\rightarrow$ simple implementation)

ALGORITHM FOR CG/PLS

Initialize: $\omega_0 = 0$; $r_0 = \mathbf{X}^t \mathbf{Y}$; $g_0 = r_0$
for $m = 0, \ldots, (m_{\max} - 1)$ **do**
   $\alpha_m = \|\mathbf{X}r_m\|^2 / \|\mathbf{X}^t\mathbf{X}g_m\|^2$
   $\omega_{m+1} = \omega_m + \alpha_m g_m$ (update)
   $r_{m+1} = r_m - \alpha_m \mathbf{X}^t\mathbf{X}g_m$ (residuals)
   $\beta_m = \|\mathbf{X}r_{m+1}\|^2 / \|\mathbf{X}r_m\|^2$
   $g_{m+1} = r_{m+1} + \beta_m g_m$ (next direction)
end for
Return: approximate solution $\omega_{m_{\max}}$

KERNEL-CG REGULARIZATION
(≈ KERNEL PARTIAL LEAST SQUARES)

▶ Define the *m*-th iterate of CG as

$$\widehat{h}_{CG(m)} = \operatorname*{Arg\,Min}_{h \in \mathcal{K}_m(S_n, T_n^* \mathbf{Y})} \|T_n^* \mathbf{Y} - h\|_{\mathcal{H}},$$

where $\mathcal{K}_m$ denotes Krylov space:

$$\mathcal{K}_m(A, z) = \operatorname{span} \left\{ z, Az, \ldots, A^{m-1} z \right\}$$

▶ equivalently:

$$\alpha_{CG(m)} = \operatorname*{Arg\,Min}_{\alpha \in \mathcal{K}_m(K_n, \mathbf{Y})} \left\| K_n^{\frac{1}{2}} \left(\mathbf{Y} - K_n \alpha \right) \right\|^2$$

and

$$\widehat{h}_{CG(m)} = \sum_{i=1}^{n} \alpha_{CG(m),i} K(X_i, .) .$$

RATES FOR CG

Consider the following stopping rule for some fixed τ

$$\widehat{m} := \min \left\{ m \geq 0 : \left\| T_n^*(T_n \widehat{h}_{CG(m)} - \mathbf{Y}) \right\| \leq \tau \left(\frac{1}{n} \log^2 \frac{6}{\delta} \right)^{\frac{r+1}{2r+1+s}} \right\} . \quad (1)$$

Theorem (Blanchard, Krämer)

Assume **(BernsteinNoise)**, **SC(r, R)**, **IP(s, β)** *hold; let* $\theta \in [0, \frac{1}{2})$. *Then for* τ *large enough, with probability larger than* $1 - \delta$ *:*

$$\left\| S^\theta (\widehat{h}_{CG(\widehat{m})} - h^*) \right\|_{\mathcal{H}_k} \leq c(r, R, s, \beta, \tau) \left(\frac{1}{n} \log^2 \frac{6}{\delta} \right)^{\frac{r+\theta}{2r+1+s}} .$$

Technical tools: again, concentration of the error in appropriate norm, and suitable reworking of the arguments of Nemirovskii (1980) for deterministic CG.

OUTER RATES

▶ It it natural (for the prediction problem) to assume extension of source condition for $h^* \notin \mathcal{H}$ (now assuming $h^* \in L^2(\mathbb{P}_X)$)

$$\mathbf{SC_{outer}}(r, R): \quad \left\| B^{-(r+\frac{1}{2})} h^* \right\|_{L^2} \leq R \qquad (\text{for } B := TT^*)$$

to include the possible range $r \in (-\frac{1}{2}, 0]$.

▶ For such "outer" source conditions, even for Kernel ridge regression and for the direct (=prediction) problem, there are no known results without additional assumptions to reach the optimal rate $\mathcal{O}\left(n^{-\frac{r+\frac{1}{2}}{2r+1+s}}\right)$.

▶ Mendelson and Neeman (2009) make assumptions on the sup norm of the eigenfunctions of the integral operator

▶ Caponnetto (2006) assumes additional unlabeled examples $X_{n+1}, \ldots, X_{\widetilde{n}}$ are available, with

$$\frac{\widetilde{n}}{n} \sim \mathcal{O}\left(n^{\frac{(1-2r-s)_+}{2r+1+s}}\right)$$

CONSTRUCTION WITH UNLABELED DATA

- assume $\widehat{n}$ i.i.d. X-examples are available, out of which n are labeled.
- extend the n vector $\mathbf{Y}$ to a $\widetilde{n}$-vector

$$\widetilde{\mathbf{Y}} = \frac{\widetilde{n}}{n} \, (Y_1, \ldots, Y_n, 0, \ldots, 0)$$

- perform the same algorithm as before on $\mathbf{X}, \widetilde{\mathbf{Y}}$.
- notice in particular that

$$T_{\widetilde{n}}^* \widetilde{\mathbf{Y}} = T_n^* \mathbf{Y}.$$

CONSTRUCTION WITH UNLABELED DATA

- assume $\widehat{n}$ i.i.d. X-examples are available, out of which n are labeled.
- extend the n vector **Y** to a $\widetilde{n}$-vector

$$\widetilde{\mathbf{Y}} = \frac{\widetilde{n}}{n}\left(Y_1, \ldots, Y_n, 0, \ldots, 0\right)$$

- perform the same algorithm as before on **X**, $\widetilde{\mathbf{Y}}$.
- notice in particular that

$$T_{\widetilde{n}}^* \widetilde{\mathbf{Y}} = T_n^* \mathbf{Y}.$$

- Recall:

$$\widehat{h}_{CG1(m)} = \underset{h \in \mathcal{K}_m(\widetilde{S}_n, T_n^* \mathbf{Y})}{\operatorname{Arg\,Min}} \|T_n \mathbf{Y} - h\|_{\mathcal{H}}$$

- equivalently:

$$\alpha = \underset{\omega \in \mathcal{K}_m(\widetilde{K}_n, \widetilde{\mathbf{Y}})}{\operatorname{Arg\,Min}} \left\| \widetilde{K}_n^{\frac{1}{2}} \left(\widetilde{\mathbf{Y}} - \widetilde{K}_n \alpha\right) \right\|^2$$

OUTER RATES FOR CG REGULARIZATION

Consider the following stopping rule for some fixed $\tau > \frac{3}{2}$,

$$\widehat{m} := \min \left\{ m \geq 0 : \left\| T_n^*(T_n \widehat{h}_{CG(m)} - \mathbf{Y}) \right\| \leq \tau M \left(\frac{4\beta}{n} \log^2 \frac{6}{\delta} \right)^{\frac{r+1}{2r+1+s}} \right\} . \quad (2)$$

Furthermore assume

$$\textbf{(BoundedY)} : \quad |Y| \leq M \qquad \text{a.s.}$$

Theorem

Assume **(BoundedY)**, **SC$_{\text{outer}}$**(r, R), **IP$^+$**(s, β), and $r \in (-\min(s, \frac{1}{2}), 0)$.

Assume unlabeled data is available with $\frac{\widetilde{n}}{n} \geq \left(\frac{16\beta^2}{n} \log^2 \frac{6}{\delta} \right)^{-\frac{(-2r)_+}{2r+1+s}}$. *Then for* $\theta \in [0, r + \frac{1}{2})$, *with probability larger than* $1 - \delta$:

$$\left\| B^{-\theta}(T h_{\widehat{m}} - h^*) \right\|_{L^2} \leq c(r, \tau)(M + R) \left(\frac{16\beta^2}{n} \log^2 \frac{6}{\delta} \right)^{\frac{r+\frac{1}{2}-\theta}{2r+1+s}} .$$

OVERVIEW:

- inverse problem setting under random i.i.d. design scheme ("learning setting"),
- for source condition: Hölder of order r;
- for ill-posedness: polynomial decay of eigenvalues of order s;
- rates of the form (for $\theta \in [0, \frac{1}{2}]$):

$$\left\| S^\theta(h^* - \widehat{h}) \right\|_{\mathcal{H}_K} \leq O\left(n^{-\frac{(r+\theta)}{2r+1+s}} \right).$$

- rates established for general linear spectral methods, as well as CG.
- matching lower bound.
- matches "classical" rates in the white noise model (=sequence model) with $\sigma^{-2} \leftrightarrow n$.
- extension to "outer rates" ($r \in (-\frac{1}{2}, 0)$) if additional **unlabeled** data available.

CONCLUSION/PERSPECTIVES

- We filled gaps in the existing picture for inverse learning methods...
- Adaptivity?
- Ideally attain optimal rates without a priori knowledge of *r* **nor** of *s*!
    - Lepski's method/balancing principle: in progress. Need a good estimator for $\mathcal{N}(\lambda)$! (Prior work on this: Caponnetto; need some sharper bound)
    - Hold-out principle: only valid for direct problem? But optimal parameter does not depend on risk norm: hope for validity in inverse case.