

Kernels Based Tests with Non-asymptotic Bootstrap Approaches for Two-sample Problems

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Task: two-sample problem

- Given: 2 independent samples

$$Z^{(1)} = \{Z_1^{(1)}, \dots, Z_{(N_1)}^{(1)}\} \sim s_1,$$

$$Z^{(2)} = \{Z_1^{(2)}, \dots, Z_{(N_2)}^{(2)}\} \sim s_2.$$

- We want to test
 - $H_0 : s_1 = s_2$, against
 - $H_1 : s_1 \neq s_2$.

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 - $H_0 : s_1 = s_2$, against
 - $H_1 : s_1 \neq s_2$.
- [\mathcal{Z} : measurable space, e.g. 'nice' $\mathcal{Z} \subseteq \mathbb{R}^d$.]

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- Error of
 - **first** kind: H_0 is true, but we reject it.
 - **second** kind: H_0 is false, but we accept it.

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 - **first** kind: H_0 is true, but we reject it.
 - **second** kind: H_0 is false, but we accept it.
- Test is called of **level $\alpha \in (0, 1)$** if
$$\mathbb{P}(\text{first kind of error}) \leq \alpha.$$
- Difficulty: hard to find the *non-asymptotic* quantile for α -levelness! \rightarrow small-sample domain.

Two-sample problem: density model – classical setup

- Examples = different $Z^{(i)}$ generating mechanisms.
- Here: $N_1 = n_1$, $N_2 = n_2$: fixed,
 - $Z^{(1)} = \{Z_1^{(1)}, \dots, Z_{(n_1)}^{(1)}\} \sim s_1$
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 - $Z^{(2)} = \{Z_1^{(2)}, \dots, Z_{(n_2)}^{(2)}\} \sim s_2 [\ll \nu]$,
 - $[\nu:$
 - non-atomic: $z \in \mathcal{Z}$ s.t. $\nu(\{z\}) > 0$ can not happen.
 - σ -finite measure: $\mathcal{Z} = \dot{\cup}_{i \in I} P_i$, I : countable, $\nu(P_i) < \infty$.]
 - $[s_1, s_2 \in L^2(\mathcal{Z}, \nu)]$

Two-sample problem: heteroscedastic regression model

- $N_1 = n_1, N_2 = n_2$: fixed,
- $Z^{(1)}, Z^{(2)}$ i.i.d.

$$Z_i^{(1)} = (X_i^{(1)}, Y_i^{(1)}) \quad Y_i^{(1)} = \textcolor{red}{s}_1(X_i^{(1)}) + \sigma(X_i^{(1)})\xi_i^{(1)},$$

$$Z_i^{(2)} = (X_i^{(2)}, Y_i^{(2)}) \quad Y_i^{(2)} = \textcolor{red}{s}_2(X_i^{(2)}) + \sigma(X_i^{(2)})\xi_i^{(2)},$$

$$(X_i^{(1)}, \xi_i^{(1)}) \stackrel{\text{distr}}{=} (X_i^{(2)}, \xi_i^{(2)}), \quad \mathbb{E} [\xi_i^{(1)} | X_i^{(1)}] = 0, \quad \mathbb{E} \left[(\xi_i^{(1)})^2 | X_i^{(1)} \right] = 1,$$

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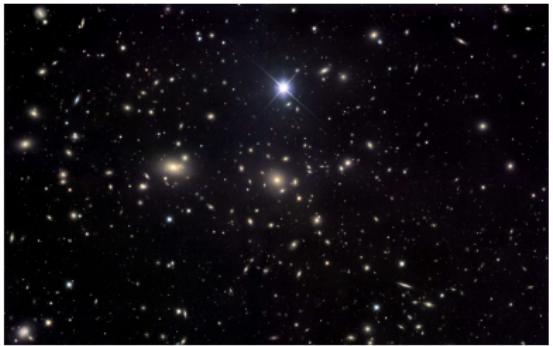
- $[s_1, s_2, \sigma \in L^2(X, P_X), \mathcal{Z} = \mathcal{X} \times \mathcal{Y}, \mathcal{Y} \subseteq \mathbb{R}]$

Two-sample problem: Poisson process model

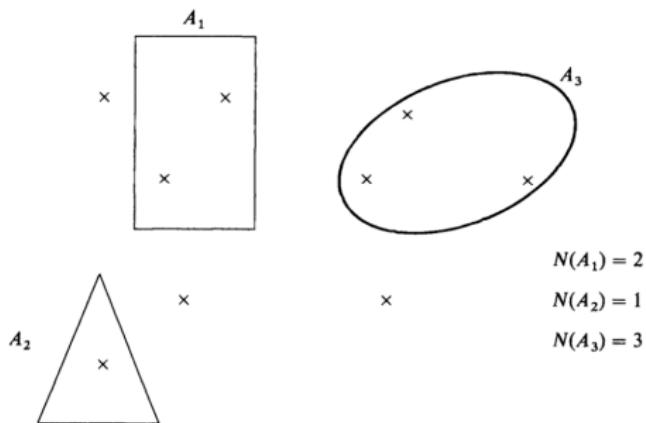
- $X \sim \text{Poisson}[\lambda] : \mathbb{P}(X = n) = \frac{\lambda^k}{k!} e^{-\lambda}, n = 0, 1, \dots ; \lambda \geq 0.$

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- # of raisins in a cake / stars on the sky:



Poisson process model



An N process producing random points in $(\mathcal{Z}, \mathcal{A}, m)$ is called PP if

- ① $N(A) \sim \text{Poisson}[m(A)]$, for all $A \in \mathcal{A}$
- ② $\{A_i\}_{i \in I}$ disjunct, I : countable $\Rightarrow \{N(A_i)\}_{i \in I}$ are independent.

Poisson process: it exists, construction

- If \mathcal{Z} is finite ($m(\mathcal{Z}) < \infty$):
 - Z_1, Z_2, \dots i.i.d., $\mathbb{P}(Z_i \in A) = \frac{m(A)}{m(\mathcal{Z})}, \forall A \in \mathcal{A}$,
 - $p \sim \text{Poisson}[m(\mathcal{Z})]$, independent of Z_i -s.
 - $N(A) = \sum_{i=1}^p \chi\{Z_i \in A\}$ is good!

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 - $N(A) = \sum_{i=1}^p \chi\{Z_i \in A\}$ is good!
- If \mathcal{Z} is σ -finite:

$$\mathcal{Z} = \dot{\cup}_{i \in I} P_i, \quad P_i \leftrightarrow N_i, \quad N = \sum_{i \in I} N_i.$$

Two-sample problem: Poisson process model

- $Z^{(1)} = \{Z_1^{(1)}, \dots, Z_{(N_1)}^{(1)}\}$, $m := s_1$ intensity
- $Z^{(2)} = \{Z_1^{(2)}, \dots, Z_{(N_2)}^{(2)}\}$, $m := s_2$
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- $Z^{(2)} = \{Z_1^{(2)}, \dots, Z_{(N_2)}^{(2)}\}$, $m := s_2$ [$\ll \mu$],
- N_1, N_2 : Poisson r.v.-s.
- $[\mu: \text{non-atomic, } \sigma\text{-finite.}]$
- $[\mu = n\nu; \nu: \text{non-atomic, } \sigma\text{-finite; } ???]$

After the examples, back to two-sampling (density model)

- Pooled samples: $Z = Z^{(1)} \cup Z^{(2)}$, $n = n_1 + n_2$. Test statistics:

$$\begin{aligned}T_K &= \sum_{i \neq j \in \{1, \dots, n\}} K(Z_i, Z_j) (\epsilon_i^0 \epsilon_j^0 + c_{n_1, n_2}), \\c_{n_1, n_2} &= \frac{1}{n_1 n_2 (n_1 - 1 + n_2 - 1)}, \\a_{n_1, n_2} &= \sqrt{\frac{1}{n_1(n_1 - 1)} - c_{n_1, n_2}}, \quad b_{n_1, n_2} = -a_{n_2, n_1}, \\\epsilon_i^0 &= \begin{cases} a_{n_1, n_2} & Z_i \in Z^{(1)} \\ b_{n_1, n_2} & Z_i \in Z^{(2)}. \end{cases}\end{aligned}$$

- Example: K = reproducing kernel.

T_K makes sense

Let $(K \diamond p)(z) = \langle K(\cdot, z), p \rangle$, $\langle f, g \rangle = \int_{\mathcal{Z}} f(z)g(z)d\nu(z)$. Unbiasedness:

$$\mathbb{E}_{s_1, s_2}[T_K] \stackrel{?}{=} \langle K \diamond (s_1 - s_2), s_1 - s_2 \rangle$$

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where $c_{n_\ell} = \frac{1}{n_\ell(n_{\ell-1})}$ ($\ell = 1, 2$).

$\langle K \diamond (s_1 - s_2), s_1 - s_2 \rangle$ is finite

By CBS ($\|\cdot\| = \|\cdot\|_{L^2(\mathcal{Z}, \nu)}$)

$$|\langle K \diamond (s_1 - s_2), s_1 - s_2 \rangle| \leq \underbrace{\|K \diamond (s_1 - s_2)\|}_{=: (*)_1} \underbrace{\|s_1 - s_2\|}_{=: (*)_2},$$

where

- $(*)_1$: assumed to be finite,
- $(*)_2 < \infty$ since $s_1, s_2 \in L^2(\mathcal{Z}, \nu)$.

If K is a reproducing kernel:

$$\begin{aligned}\mathbb{E}_{s_1, s_2}[T_K] &= \int_{\mathcal{Z}^2} K(z, z')(s_1 - s_2)(z)(s_1 - s_2)(z') d\nu(z) d\nu(z') \\ &= \left\| \int_{\mathcal{Z}} K(\cdot, z)(s_1 - s_2)(z) d\nu(z) \right\|_{H(K)}^2 \\ &= \|\mu_{s_1} - \mu_{s_2}\|_{H(K)}^2, \quad \mu_s = \mu_{s, K}.\end{aligned}$$

$\Rightarrow \mathbb{E}_{s_1, s_2}[T_K] = 0 \Leftrightarrow s_1 = s_2$, if K is characteristic.

K -examples: K is a projection kernel

Let $\{\phi_\lambda\}_{\lambda \in \Lambda}$ be an ONS.

$$K(z, z') = \sum_{\lambda \in \Lambda} \phi_\lambda(z) \phi_\lambda(z'),$$

$$\int_{\mathcal{Z}} K(z, z') f(z') d\nu(z') = \text{proj}_S(f)(z),$$

$$S = \text{span}(\{\phi_\lambda\}_{\lambda \in \Lambda}),$$

$$\langle K \diamond (s_1 - s_2), s_1 - s_2 \rangle = \|\text{proj}_S(s_1 - s_2)\|_{L^2(\mathcal{Z}, \nu)}.$$

Example: $\{\phi_\lambda\}_{\lambda \in \Lambda}$ Fourier/Haar basis.

K -examples: K is a convolution kernel

$\mathcal{Z} = \mathbb{R}^d$, $\nu = \text{Lebesgue measure}$, $k \in L^2(\mathbb{R}^d)$, $k(-x) = k(x)$, $h_i > 0$ ($\forall i$).

$$K(z, z') = \frac{1}{\prod_{i=1}^d h_i} k \left(\frac{z_1 - z'_1}{h_1}, \dots, \frac{z_d - z'_d}{h_d} \right),$$

$$\langle K \diamond (s_1 - s_2), s_1 - s_2 \rangle = \langle k_h * (s_1 - s_2), s_1 - s_2 \rangle_{L^2(\mathbb{R}^d)},$$

$$(f * g)(x) = \int_{\mathbb{R}^d} f(y)g(x - y)dy,$$

$$k_h(u_1, \dots, u_d) = \frac{1}{\prod_{i=1}^d h_i} k \left(\frac{u_1}{h_1}, \dots, \frac{u_d}{h_d} \right).$$

Distribution of $T_K|Z$

- Recall: $T_K = \sum_{i \neq j \in \{1, \dots, n\}} K(Z_i, Z_j) (\epsilon_i^0 \epsilon_j^0 + c_{n_1, n_2})$,

$$\epsilon_i^0 = \begin{cases} a_{n_1, n_2} & i \in \{1, \dots, n_1\} \\ b_{n_1, n_2} & i \in \{n_1 + 1, \dots, n_1 + n_2\}. \end{cases}$$

- $R = (R_1, \dots, R_n)$: rnd permutation, independent of Z .

$$\epsilon_i = \begin{cases} a_{n_1, n_2} & i \in \{R_1, \dots, R_{n_1}\} \\ b_{n_1, n_2} & i \in \{R_{n_1+1}, \dots, R_{n_1+n_2}\}. \end{cases},$$

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- $T_K|Z \stackrel{\text{distr}}{=} T_K^\epsilon|Z$.

Final test, level

- $q_{K,1-\alpha}^{(Z)}$: $(1 - \alpha)$ -quantile of $T_K^\epsilon | Z \stackrel{distr}{=} T_K | Z$.
- Reject H_0 if $T_K > q_{K,1-\alpha}^{(Z)}$.

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- Taking expectation over Z , $\sup [\mathbb{P}_{(H_0)}(A) = \sup_{(s_1, s_2), s_1 = s_2} \mathbb{P}_{s_1, s_2}(A)]$:

$$\mathbb{P}_{(H_0)}(\Phi_{K,\alpha} = 1) \leq \alpha,$$

$$\Phi_{K,\alpha} = \chi \left\{ T_K > q_{K,1-\alpha}^{(Z)} \right\}.$$

Aggregated test

- Given:
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$$u_\alpha^{(Z)} := \sup \left\{ u > 0 : \mathbb{P} \left(\underbrace{\sup_m (T_{K_m}^\epsilon - q_{m,1-ue^{-w_m}}^{(Z)}) > 0}_{\geq 1 \text{ test rejects}} \mid Z \right) \leq \alpha \right\}.$$

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- The test: $\exists m : T_{K_m}^\epsilon - q_{m,1-u_\alpha^{(Z)}e^{-w_m}}^{(Z)} > 0$; it is of level α .

Summary

- Focus: two-sample problem.
- Unbiasedness and exact quantile for the test statistics.
- Single kernel and aggregated tests of level α .

