Probabilistic & Unsupervised Learning Approximate Inference

Exponential families: convexity, duality and free energies

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Term 1, Autumn 2021

Exponential families: the log partition function

Consider an exponential family distribution with sufficient statistic s(X) and natural parameter θ (and no base factor in X alone). We can write its probability or density function as

$$p(X|\theta) = \exp\left(\theta^{\mathsf{T}}s(X) - \Phi(\theta)\right)$$

where $\Phi(\theta)$ is the log partition function

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 $\Phi(\theta)$ plays an important role in the theory of the exponential family. For example, it maps natural parameters to the moments of the sufficient statistics:

$$\begin{split} \frac{\partial}{\partial \theta} \Phi(\theta) &= e^{-\Phi(\theta)} \sum_{x} s(x) e^{\theta^{\mathsf{T}} s(x)} = \mathbb{E}_{\theta} \left[s(X) \right] = \mu(\theta) = \mu \\ \frac{\partial^{2}}{\partial \theta^{2}} \Phi(\theta) &= e^{-\Phi(\theta)} \sum_{x} s(x)^{2} e^{\theta^{\mathsf{T}} s(x)} - e^{-2\Phi(\theta)} \left[\sum_{x} s(x) e^{\theta^{\mathsf{T}} s(x)} \right]^{2} = \mathbb{V}_{\theta} \left[s(X) \right] \end{split}$$

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The second derivative is thus positive semi-definite, and so $\Phi(\theta)$ is convex in θ .

Exponential families: mean parameters and negative entropy

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Consider the negative entropy of the distribution as a function of the mean parameter:

$$\Psi(\mu) = \mathbb{E}_{\theta} \left[\log p(X|\theta(\mu)) \right] = \theta^{\mathsf{T}} \mu - \Phi(\theta)$$

so

$$\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{\mu} = \Phi(\boldsymbol{\theta}) + \Psi(\boldsymbol{\mu})$$

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The negative entropy is dual to the log-partition function. For example,

$$\begin{split} \frac{d}{d\mu} \Psi(\mu) &= \frac{\partial}{\partial \mu} \left(\theta^{\mathsf{T}} \mu - \Phi(\theta) \right) + \frac{d\theta}{d\mu} \frac{\partial}{\partial \theta} \left(\theta^{\mathsf{T}} \mu - \Phi(\theta) \right) \\ &= \theta + \frac{d\theta}{d\mu} (\mu - \mu) = \theta \end{split}$$

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$$\begin{aligned} \mathsf{KL}\big[\boldsymbol{\theta}\big\|\boldsymbol{\theta}'\big] &= \mathsf{KL}\big[p(\boldsymbol{X}|\boldsymbol{\theta})\big\|p(\boldsymbol{X}|\boldsymbol{\theta}')\big] = \mathbb{E}_{\boldsymbol{\theta}}\left[-\log p(\boldsymbol{X}|\boldsymbol{\theta}') + \log p(\boldsymbol{X}|\boldsymbol{\theta})\right] \\ &= -\boldsymbol{\theta}'^{\mathsf{T}}\boldsymbol{\mu} + \Phi(\boldsymbol{\theta}') + \Psi(\boldsymbol{\mu}) \geq 0 \\ &\Rightarrow \Psi(\boldsymbol{\mu}) \geq \boldsymbol{\theta}'^{\mathsf{T}}\boldsymbol{\mu} - \Phi(\boldsymbol{\theta}') \end{aligned}$$

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where μ are the mean parameters corresponding to θ .

Now, the minimum KL divergence of zero is reached iff $oldsymbol{ heta} = oldsymbol{ heta}'$, so

$$\Psi(\mu) = \sup_{\theta'} \left[\theta'^{\mathsf{T}} \mu - \Phi(\theta') \right] \qquad \text{ and, if finite } \quad \theta(\mu) = \operatorname*{argmax}_{\theta'} \left[\theta'^{\mathsf{T}} \mu - \Phi(\theta') \right]$$

The left-hand equation is the definition of the conjugate dual of a convex function.

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Continuous functions are reciprocally dual, so we also have:

$$\Phi(\theta) = \sup_{\mu'} \left[\theta^\mathsf{T} \mu' - \Psi(\mu') \right] \qquad \text{ and, if finite } \quad \mu(\theta) = \operatorname*{argmax}_{\mu'} \left[\theta^\mathsf{T} \mu' - \Psi(\mu') \right]$$

Thus, duality gives us another relation between θ and μ .

Consider a joint exponential family distribution on observed ${\bf x}$ and latent ${\bf z}$.

$$p(\mathbf{x}, \mathbf{z}) = \exp\left[\boldsymbol{\theta}^{\mathsf{T}} s(\mathbf{x}, \mathbf{z}) - \Phi_{X\!Z}(\boldsymbol{\theta})\right]$$

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The posterior on \mathbf{z} is also in the exponential family, with the clamped sufficient statistic $s_Z(\mathbf{z}; \mathbf{x}) = s_{XZ}(\mathbf{x}^{\text{obs}}, \mathbf{z})$; the same (now possibly redundant) natural parameter θ ; and partition function $\Phi_Z(\theta) = \log \sum_{\mathbf{z}} \exp \theta^T s_Z(\mathbf{z})$.

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The likelihood is

$$\mathcal{L}(\theta) = p(\mathbf{x}|\theta) = \sum_{\mathbf{z}} e^{\theta^{\mathsf{T}} s(\mathbf{x}, \mathbf{z}) - \Phi_{XZ}(\theta)} = \sum_{\mathbf{z}} e^{\theta^{\mathsf{T}} s_{Z}(\mathbf{z}; \mathbf{x})} e^{-\Phi_{XZ}(\theta)} = \exp[\Phi_{Z}(\theta) - \Phi_{XZ}(\theta)]$$

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So we can write the log-likelihood as

$$\ell(\boldsymbol{\theta}) = \sup_{\boldsymbol{\mu}_{\boldsymbol{Z}}} [\underbrace{\boldsymbol{\theta}^{\mathsf{T}} \boldsymbol{\mu}_{\boldsymbol{Z}} - \boldsymbol{\Phi}_{\boldsymbol{X}\boldsymbol{Z}}(\boldsymbol{\theta})}_{\langle \log \rho(\mathbf{x}, \mathbf{z}) \rangle_q} - \underbrace{\boldsymbol{\Psi}(\boldsymbol{\mu}_{\boldsymbol{Z}})}_{-\mathbf{H}[q]}] = \sup_{\boldsymbol{\mu}_{\boldsymbol{Z}}} \mathcal{F}(\boldsymbol{\theta}, \boldsymbol{\mu}_{\boldsymbol{Z}})$$

This is the familiar free energy with $q(\mathbf{z})$ represented by its mean parameters $\mu_{\mathcal{Z}}$!

We have described inference in terms of the distribution q, approximating as needed, then computing expected suff stats. Can we describe it instead as an optimisation over μ directly?

$$\boldsymbol{\mu}_{\boldsymbol{\mathcal{Z}}}^* = \operatorname*{argmax}[\boldsymbol{\theta}^{\mathsf{T}}\boldsymbol{\mu}_{\boldsymbol{\mathcal{Z}}} - \boldsymbol{\Psi}(\boldsymbol{\mu}_{\boldsymbol{\mathcal{Z}}})]$$

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Concave maximisation(!), but two complications:

The optimum must be found over feasible means. Interdependance of the sufficient statistics may prevent arbitrary sets of mean sufficient statistics being achieved

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- The optimum must be found over feasible means. Interdependance of the sufficient statistics may prevent arbitrary sets of mean sufficient statistics being achieved
 - Feasible means are convex combinations of all the single-configuration sufficient statistics.

$$\mu = \sum_{\mathbf{x}} \nu(\mathbf{x}) s(\mathbf{x}) \qquad \sum_{\mathbf{x}} \nu(\mathbf{x}) = 1$$

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- ▶ Take a Boltzmann machine on two variables, x_1 , x_2 .
- ► The sufficient stats are $s(\mathbf{x}) = [x_1, x_2, x_1x_2]$.
- Clearly only the stats $S = \{[0, 0, 0], [0, 1, 0], [1, 0, 0], [1, 1, 1]\}$ are possible.
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- ▶ Thus μ ∈ convex hull(\mathcal{S}).
- For a discrete distribution, this space of possible means is bounded by exponentially many hyperplanes connecting the discrete configuration stats: called the marginal polytope.
- Even when restricted to the marginal polytope, evaluating $\Psi(\mu)$ can be challenging.

Convexity and undirected trees

We can parametrise a discrete pairwise MRF as follows:

$$\rho(\mathbf{X}) = \frac{1}{Z} \prod_{i} f_i(X) \prod_{(ij)} f_{ij}(X_i, X_j)$$

$$= \exp\left(\sum_{i} \sum_{k} \theta_i(k) \delta(X_i = k) + \sum_{(ij)} \sum_{k,l} \theta_{ij}(k, l) \delta(X_i = k) \delta(X_j = l) - \Phi(\boldsymbol{\theta})\right)$$

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So discrete MRFs are always exponential family, with natural and mean parameters:

$$\theta = [\theta_i(k), \theta_{ij}(k, l) \quad \forall i, j, k, l]$$

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If the MRF has tree structure T, the negative entropy can be written in terms of the single-site entropies and mutual informations on edges:

$$\Psi(\mu_{\tau}) = \mathbb{E}_{\theta_{\tau}} \left[\log \prod_{i} p(X_{i}) \prod_{(ij) \in \tau} \frac{p(X_{i}, X_{j})}{p(X_{i})p(X_{j})} \right]$$
$$= -\sum_{i} H(X_{i}) + \sum_{(ij) \in \tau} I(X_{i}, X_{j})$$

We can see the Bethe free energy problem as a relaxation of the true free-energy optimisation:

$$\boldsymbol{\mu}_{\boldsymbol{\mathcal{Z}}}^* = \operatorname*{argmax}_{\boldsymbol{\mu}_{\boldsymbol{\mathcal{Z}}} \in \mathcal{M}} [\boldsymbol{\theta}^\mathsf{T} \boldsymbol{\mu}_{\boldsymbol{\mathcal{Z}}} - \boldsymbol{\Psi}(\boldsymbol{\mu}_{\boldsymbol{\mathcal{Z}}})]$$

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- 2. Approximate $\Psi(\mu_Z)$ by the tree-structured form

$$\Psi_{\mathsf{Bethe}}(\mu_{\mathcal{Z}}) = -\sum_{i} \mathit{H}(\mathit{X}_{i}) + \sum_{(ij) \in \mathbf{G}} \mathit{I}(\mathit{X}_{i}, \mathit{X}_{j})$$

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 $\mathcal L$ is still a convex set (polytope for discrete problems). However Ψ_{Bethe} is not convex.

Convexifying BP

Consider instead an upper bound on $\Phi(\theta)$:

Imagine a set of spanning trees T for the MRF, each with its own parameters θ_T, μ_T . By padding entries corresponding to off-tree edges with zero, we can assume that θ_T has the same dimensionality as θ .

Suppose also that we have a distribution β over the spanning trees so that $\mathbb{E}_{\beta}\left[\boldsymbol{\theta}_{7}\right]=\boldsymbol{\theta}.$

Then by the convexity of $\Phi(\theta)$,

$$\Phi(oldsymbol{ heta}) = \Phi(\mathbb{E}_eta\left[oldsymbol{ heta}_ au
ight]) \leq \mathbb{E}_eta\left[\Phi(oldsymbol{ heta}_ au)
ight]$$

If we were to tighten the upper bound we might obtain a good approximation to Φ :

$$\Phi(\boldsymbol{\theta}) \leq \inf_{\boldsymbol{\beta}, \boldsymbol{\theta}_{\mathcal{T}} : \mathbb{E}_{\boldsymbol{\beta}}[\boldsymbol{\theta}_{\mathcal{T}}] = \boldsymbol{\theta}} \mathbb{E}_{\boldsymbol{\beta}} \left[\Phi(\boldsymbol{\theta}_{\mathcal{T}}) \right]$$

Convex Upper Bounds on the Log Partition Function

$$\Phi(oldsymbol{ heta}) \leq \inf_{oldsymbol{ heta}_{ au}: \mathbb{E}_{eta}[oldsymbol{ heta}_{ au}] = oldsymbol{ heta}} \mathbb{E}_{eta}\left[\Phi(oldsymbol{ heta}_{ au})
ight] \stackrel{ ext{def}}{=} \Phi_{eta}(oldsymbol{ heta})$$

Solve the constrained optimisation problem using Lagrange multipliers:

$$\mathcal{L} = \mathbb{E}_{eta} \left[\Phi(oldsymbol{ heta}_{ au})
ight] - oldsymbol{\lambda}^{\mathsf{T}} (\mathbb{E}_{eta} \left[oldsymbol{ heta}_{ au}
ight] - oldsymbol{ heta})$$

Setting the derivatives wrt θ_T to zero, we get:

$$\frac{\partial}{\partial \boldsymbol{\theta}_{T}} \sum_{\tau} \beta(T) \Phi(\boldsymbol{\theta}_{T}) - \boldsymbol{\lambda}^{\mathsf{T}} \frac{\partial}{\partial \boldsymbol{\theta}_{T}} \sum_{\tau} \beta(T) \boldsymbol{\theta}_{T} = 0$$
$$\beta(T) \boldsymbol{\mu}_{T} - \beta(T) \Pi_{T}(\boldsymbol{\lambda}) = 0$$
$$\boldsymbol{\mu}_{T} = \Pi_{T}(\boldsymbol{\lambda})$$

where $\Pi_T(\lambda)$ selects the Lagrange multipliers corresponding to elements of θ that are non-zero in the tree T.

Although each tree has its own parameters θ_T , at the optimum they are all constrained: their mean parameters are all consistent with each other (c.f. the tree-reparametrisation view of BP) and with the Lagrange multipliers λ .

Convex Upper Bounds on the Log Partition Function

$$\begin{split} \Phi_{\beta}(\theta) &= \sup_{\lambda} \inf_{\theta_{\mathcal{T}}} \mathbb{E}_{\beta} \left[\Phi(\theta_{\mathcal{T}}) \right] - \lambda^{\mathsf{T}} (\mathbb{E}_{\beta} \left[\theta_{\mathcal{T}} \right] - \theta) \\ &= \sup_{\lambda} \lambda^{\mathsf{T}} \theta + \mathbb{E}_{\beta} \left[\inf_{\theta_{\mathcal{T}}} \Phi(\theta_{\mathcal{T}}) - \theta_{\mathcal{T}}^{\mathsf{T}} \Pi_{\mathcal{T}}(\lambda) \right] \\ &= \sup_{\lambda} \lambda^{\mathsf{T}} \theta + \mathbb{E}_{\beta} \left[-\Psi(\Pi_{\mathcal{T}}(\lambda)) \right] \\ &= \sup_{\lambda} \lambda^{\mathsf{T}} \theta + \mathbb{E}_{\beta} \left[\sum_{i} H_{\lambda}(X_{i}) - \sum_{(ij) \in \mathcal{T}} I_{\lambda}(X_{i}, X_{j}) \right] \\ &= \sup_{\lambda} \lambda^{\mathsf{T}} \theta + \sum_{i} H_{\lambda}(X_{i}) - \sum_{(ij)} \beta_{ij} I_{\lambda}(X_{i}, X_{j}) \end{split}$$

- This is a convexified version of the Bethe free energy.
- Optimisation wrt λis approximate inference applied to the tighest bound on Φ(θ) for fixed β.
- ▶ The bound holds for any β and can be tightened by further minimisation.

EP free energy

A Bethe-like approach also casts EP as a variational energy fixed point method.

Consider finding marginals of a (posterior) distribution defined by clique potentials:

$$P(\mathcal{Z}) \propto f_0(\mathcal{Z}) \prod_i f_i(\mathcal{Z}_i)$$

where all factor have exponential form, f_0 is in a tractable exponential family (possibly uniform) but he f_i are jointly intractable – i.e. product cannot be marginalised, although individual terms may be (numerically) tractable.

Augment by including tractable ExpFam terms with zero natural parameters

$$\textit{P}(\mathcal{Z}) \propto e^{\theta_0^T s_0(\mathcal{Z})} \prod_{\textit{i}} e^{\theta_{\textit{i}}^T s_{\textit{i}}(\mathcal{Z}_{\textit{i}})} e^{0^T \tilde{s}_{\textit{i}}(\mathcal{Z}_{\textit{i}})} = e^{\theta_0^T s_0(\mathcal{Z}) + \sum_{\textit{i}} \left(\theta_{\textit{i}}^T s_{\textit{i}}(\mathcal{Z}_{\textit{i}}) + \tilde{\theta}^T \tilde{s}(\mathcal{Z}_{\textit{i}})\right)}$$

Now, the variational dual principle tells us that the expected sufficient statistics:

$$\mu_0^* = \langle \mathbf{s}_0 \rangle_P; \quad \mu_i^* = \langle \mathbf{s}_i(\mathcal{Z}_i) \rangle_P; \quad \tilde{\mu}_i^* = \langle \tilde{\mathbf{s}}_i \rangle_P$$

are given by

$$\{\boldsymbol{\mu}_0^*, \boldsymbol{\mu}_i^*, \tilde{\boldsymbol{\mu}}_i^*\} = \operatorname*{argmax}_{\{\boldsymbol{\mu}_0, \boldsymbol{\mu}_i, \tilde{\boldsymbol{\mu}}_i\} \in \mathcal{M}} \left[\boldsymbol{\theta}_0^\mathsf{T} \boldsymbol{\mu}_0 + \sum_i \left(\boldsymbol{\theta}_i^\mathsf{T} \boldsymbol{\mu}_i + \boldsymbol{0}^\mathsf{T} \tilde{\boldsymbol{\mu}}_i\right) - \Psi(\boldsymbol{\mu}_0, \boldsymbol{\mu}_i, \tilde{\boldsymbol{\mu}}_i)\right]$$

EP relaxation

The EP algorithm relaxes this optimisation:

- ▶ Relax \mathcal{M} to locally consistent marginals, retaining consistency across each edge connecting $\{\mu_0, \tilde{\mu}_i\}$ (as in BP on a junction graph); and between pairs $(\mu_i, \tilde{\mu}_i)$.
- ▶ Replace negative entropy by $\Psi_{\text{Bethe}}(\{\mu_0, \tilde{\mu}_i\}) \sum_i (\mathbf{H}[\mu_i, \tilde{\mu}_i] \mathbf{H}[\tilde{\mu}_i]).$
- In effect, drop links between different μ_i and run reparameterisation on a junction graph.

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The free-energy-based approximate marginals include μ_i which are refined during updates.

- ▶ Direct learning on the EP free-energy uses these marginals rather than the approximate ones and a local normaliser formed by integrating over $f_i(Z_i)q_{\neg i}(Z_i)$.
- A different derivation to the result from that in EP lecture.

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